PASS Sample Size Software NCSS.com

Chapter 247

Superiority by a Margin Tests for the Ratio of Two Poisson Rates in a Cluster-Randomized Design

Introduction

This procedure calculates power and sample size for superiority by a margin tests of the ratio of two rates in a cluster-randomized design in which the outcome variable is a count. It uses the work of Wang, Zhang, and Ahn (2018) which give the power for the case of varying cluster sizes. The analysis uses a simple z-test comparing the two rates.

Cluster-randomized designs are those in which whole clusters of subjects (classes, hospitals, communities, etc.) are put into the treatment group or the control group. Generally speaking, the larger the cluster sizes and the higher the correlation among subjects within the same cluster, the larger will be the overall sample size necessary to detect an effect with the same power.

Hypotheses

When higher rates are better, the superiority by a margin test hypotheses are

$$H_0: \frac{\lambda_1}{\lambda_2} \le R_0$$
 vs. $H_1: \frac{\lambda_1}{\lambda_2} > R_0$

where $R_0 > 1$.

When higher rates are worse, the superiority by a margin test hypotheses are

$$H_0: \frac{\lambda_1}{\lambda_2} \ge R_0$$
 vs. $H_1: \frac{\lambda_1}{\lambda_2} < R_0$

where $R_0 < 1$.

Technical Details

The following discussion summarizes the results in Wang, Zhang, and Ahn (2018).

Suppose you are interested in comparing the incidence rates of two groups (treatment and control) with a superiority by a margin test. Further suppose that the response is known to be related to other covariates (such as age, race, or gender) and so their impact needs to be adjusted for. This may be accomplished by stratifying on the covariates and forming hypotheses about a common mean difference across all clusters and strata. Often, the stratification is based on cluster size, but this is not required.

Let Y_{jki} be the count outcome of the ith $(i = 1, ..., M_{jk})$ subject in the kth $(k = 1, ..., K_j)$ cluster of the jth (j = 1, 2) group. Assuming that Y_{jki} follows a Poisson model with

$$\lambda_j = \mathrm{E}(Y_{jki}) = \mathrm{Var}(Y_{jki})$$

and a common intracluster correlation coefficient (ICC) $\rho = \operatorname{corr}(Y_{jki}, Y_{jki'})$ for $i \neq i'$.

Test Statistic

An unbiased estimator of λ_i is

$$\hat{\lambda}_j = \frac{\sum_{k=1}^{K_j} \sum_{i=1}^{M_{jk}} Y_{jki}}{\sum_{k=1}^{K_j} M_{jk}}$$

with estimated variance

$$s_j^2 = \frac{\hat{\lambda}_j \sum_{k=1}^{K_j} M_{jk} [1 + (M_{jk} - 1)\hat{\rho}]}{\left(\sum_{k=1}^{K_j} M_{jk}\right)^2}$$

Using a one-sided hypothesis $H0: \lambda_1 = \lambda_2$ versus $H1: \lambda_1 < \lambda_2$, H0 is rejected if

$$\frac{\hat{\lambda}_1 - \hat{\lambda}_2}{\sqrt{s_1^2 + s_2^2}} > z_{1-\alpha}$$

Sample Size and Power

Wang, Zhang, and Ahn (2018) provide the following formula for estimating K_2 (the number of clusters in the control group) when there are an unequal number of clusters in each group.

$$K_2 = \frac{\left(z_{1-\frac{\alpha}{2}} + z_{1-\beta}\right)^2 \left(\frac{\lambda_1}{R} + \lambda_2\right)}{(\lambda_1 - \lambda_2)^2} \left[(1-\rho)\frac{1}{\theta} + \rho + \rho\gamma^2 \right]$$

where $R = K_1/K_2$, $M = \mathrm{E}(M_{jk})$, $\tau^2 = \mathrm{Var}(M_{jk})$, and $\gamma = \tau/M$. Here M is the average cluster size of all clusters in the study and γ is the coefficient of variation of the cluster sizes. Note that $z_x = \Phi(x)$ is the standard normal distribution function.

This equation can easily be rearranged to provide a formula for power $(1 - \beta)$.

Example 1 – Finding the Number of Clusters

A superiority study is being planned to investigate whether a new intervention will decrease the incidence rate of a certain disease over the rate achieved by the current intervention. The response is a count. For a number of reasons, the researchers decide to administer the intervention to whole clusters (clinics) rather than randomize the treatment to individuals within a cluster. The number of clinics receiving each treatment will be balanced.

The average number of subjects per clinic is 21. The coefficient of variation of the cluster sizes is 0.42.

Prior studies have obtained an incidence rate of 0.35 for the current treatment and an ICC of 0.07. The superiority ratio is set to 0.86. The researchers want to compare the necessary sample size when the actual ratio is 0.43, 0.57, and 0.71.

The one-sided significance level is set to 0.025 and the power is set to 0.8.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	K1 (Number of Clusters in Group 1)
Higher Rates Are	Worse (H1: λ1 / λ2 < R0, where R0 < 1)
Power	0.80
Alpha	0.025
K2 (Clusters in Group 2)	K1
M (Average of Cluster Sizes)	21
CV (Coef of Variation of Cluster Sizes)	0.42
Incidence Rate Input Type	Ratios (R0 and R1)
R0 (Superiority Ratio)	0.86
R1 (Actual Ratio)	0.43 0.57 0.71
λ2 (Incidence Rate of Control Group)	0.35
ρ (Intracluster Correlation, ICC)	0.07

Output

Click the Calculate button to perform the calculations and generate the following output.

Numeric Reports

Numeric Results

Solve For: K1 (Number of Clusters in Group 1)

Groups: 1 = Treatment, 2 = Control

Higher Rates Are: Worse

Hypotheses: $H0: \lambda 1 / \lambda 2 \ge R0$ vs. $H1: \lambda 1 / \lambda 2 < R0$

								In					
	Number of Clusters			Cluster Size		Total Sample	Treatment			Ratio		Intracluster	
Power	Treatment K1	Control K2	Total K	Average M	cv		Superiority λ1.0	Actual λ1.1	Control λ2	Superiority R0	Actual R1	Correlation ρ	Alpha
0.80020	187	187	374	21	0.42	7854	0.301	0.1505	0.35	0.86	0.43	0.07	0.025
0.80153	208	208	416	21	0.42	8736	0.301	0.1995	0.35	0.86	0.57	0.07	0.025
0.80090	228	228	456	21	0.42	9576	0.301	0.2485	0.35	0.86	0.71	0.07	0.025
CV N λ1.0	Т	average he total s	cluste ample	r size. size, i.e.,	the to	otal numb	per of subjection the treatm	cts from	all cluste		ei sizes	divided by th	ie
λ1.1 λ2				ence rate to			nt group at w	hich the	e power is	s computed.			
R0	Т		nce ra			•	null hypothe	esis. Thi	s is often	called the s	uperiori	ty boundary.	R0 =
R1	Т			te ratio as = λ1.1 / λ		d by the	alternative h	nypothe	sis. This i	s the ratio a	t which	the power is	
ρ		he intracl cluster.	uster	correlation	coef	ficient. TI	nis is the co	rrelation	between	any two sul	bjects w	ithin a partic	ular
Alpha	Т	he proba	bility o	f rejecting	a tru	e null hvi	oothesis.						

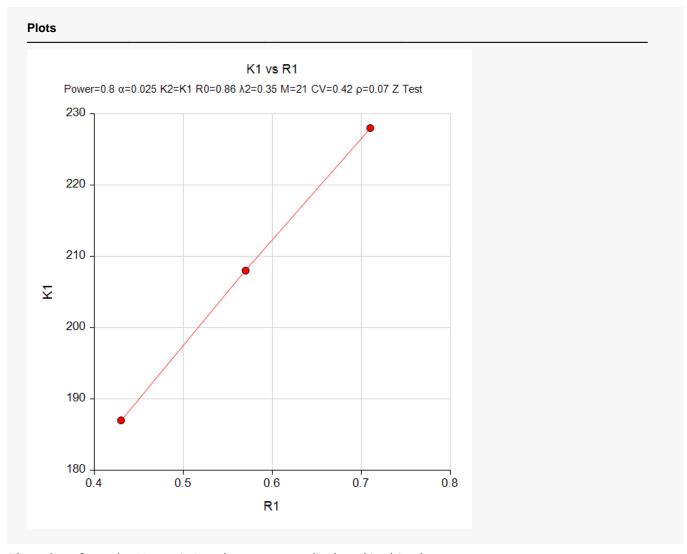
Summary Statements

A parallel two-group cluster-randomized design will be used to test whether the Group 1 (treatment) incidence rate (λ 1) is superior to the Group 2 (control) incidence rate (λ 2) by a margin, with an incidence rate superiority ratio of 0.86 (H0: λ 1 / λ 2 \geq 0.86 versus H1: λ 1 / λ 2 < 0.86). The comparison will be made using a one-sided incidence rate difference Z-test with a Type I error rate (α) of 0.025. The intracluster correlation coefficient is assumed to be 0.07. The control group incidence rate is assumed to be 0.35. The individual cluster sizes (the number of subjects per cluster) are assumed to vary according to a discrete distribution with mean 21 and coefficient of variation 0.42. To detect an incidence rate ratio (λ 1 / λ 2) of 0.43 (or treatment group incidence rate of 0.1505) with 80% power, with an average of 21 subjects per cluster, the number of clusters needed will be 187 in Group 1 (treatment) and 187 in Group 2 (control) (for an overall total of 7854 subjects).

References

Wang, J., Zhang, S., and Ahn, C. 2018. 'Sample size calculation for count outcomes in cluster randomization trials with varying cluster sizes.' Communications in Statistics--Theory and Methods, DOI: 10.1080/03610926.2018.1532004.

Plots Section



The values from the Numeric Results report are displayed in this plot.

Example 2 – Validation using a Previously Validated Procedure

We could not find a validation example in the literature so we will use a previously validated procedure in **PASS** to validate this procedure.

Suppose in a superiority design, higher rates are better, power = 0.9, alpha = 0.025, K2 = K1, M = 50, CV = 0.2, λ 1.0 = 0.6, λ 1.1 = 0.5, λ 2 = 0.5, and ρ = 0.002. Solve for K1.

This scenario can be solved using the procedure "Tests for the Difference Between Two Poisson Rates in a Cluster-Randomized Design" with the following settings.

Set the alternative hypothesis to "One-Sided (H1: δ > 0)", power = 0.9, alpha = 0.025, K2 = K1, M = 50, CV = 0.2, λ 1 = 0.6, λ 2 = 0.5, and ρ = 0.002. The solution is K1 = K2 = 26.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	K1 (Number of Clusters in Group 1)
Higher Rates Are	Better (H1: λ1 / λ2 > R0, where R0 > 1)
Power	0.9
Alpha	0.025
K2 (Clusters in Group 2)	K1
M (Average of Cluster Sizes)	50
CV (Coef of Variation of Cluster Sizes)	0.2
Incidence Rate Input Type	Incidence Rates (λ1.0 and λ1.1)
λ1.0 (Superiority Incidence Rate)	0.6
λ1.1 (Actual Incidence Rate)	0.5
λ2 (Incidence Rate of Control Group)	0.5
ρ (Intracluster Correlation, ICC)	0.002

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Superiority by a Margin Tests for the Ratio of Two Poisson Rates in a Cluster-Randomized Design

Output

Click the Calculate button to perform the calculations and generate the following output.

Solve For Groups: Higher For Hypothe	Rates Are:	K1 (Number 1 = Treatm Better H0: λ1 / λ2	ent, 2 =	Control	•	•		Inc	cidence R	ate			
		per of Clusters Cluster Size			Total	Treatment							
	Numb	er of Clust	ers	Cluster	Size		rreaum	ent		Ratio	,		
Power	Treatmen K	t Control		Average		Sample Size N			Control λ2		Actual R1	Intracluster Correlation ρ	Alpha

PASS has also obtained K1 = K2 = 26. Thus, the procedure is validated.