

## Chapter 725

# Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

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## Introduction

This procedure calculates the assurance of an equivalence test that uses a regression coefficient test from Cox proportional hazards regression. The formulation for testing the significance of a Cox regression coefficient is identical to that of the standard logrank test, so the power and sample size formulas for one analysis also works for the other. This procedure is documented in Chow, Shao, and Wang (2008) which summarizes the work of Schoenfeld (1981, 1983). Refer to the *Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model* procedure for more details. The assurance calculation used in this procedure is based on O'Hagan, Stevens, and Campbell (2005).

The Cox Regression model has the added benefit over the exponential model that it does not assume that the hazard rates are constant across time, but only that they are proportional. That is, that the hazard ratio remains constant throughout the experiment, even if the hazard rates vary from one time point to another.

The assurance calculation in this procedure is based on a user-specified prior distribution of the applicable parameters. This procedure may also be used to determine the needed sample size to obtain a specified assurance.

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## Assurance

The assurance of a design is the expected value of the power with respect to one or more prior distributions of the design parameters. Assurance is also referred to as *Bayesian assurance*, *expected power*, *average power*, *statistical assurance*, *hybrid classical-Bayesian procedure*, or *probability of success*.

The power of a design is the probability of rejecting the null hypothesis, conditional on a given set of design attributes, such as the test statistic, the significance level, the sample size, and the effect size to be detected. As many of the parameters are typically unknown quantities, the stated power may be very different from the true power if the specified parameter values are inaccurate.

While power is conditional on individual design parameter values, and is highly sensitive to those values, assurance is the average power across a presumed prior distribution of the parameters. Thus, assurance adds a Bayesian element to the frequentist framework, resulting in a hybrid approach to the probability of trial or study success. It should be noted that when it comes time to perform the statistical test on the resulting data, these methods for calculating assurance assume that the traditional (frequentist) tests will be used.

The next section describes some of the ways in which the prior distributions for effect size parameters may be determined.

## Elicitation

In order to calculate assurance, a suitable prior distribution for the applicable parameters must be determined. This process is called the *elicitation* of the prior distribution.

The elicitation may be as simple as choosing a distribution that seems plausible for the parameter(s) of interest, or as complex as combining the informed advice of several experts based on experience in the field, available pilot data, or previous studies. The accuracy of the assurance value depends on the accuracy of the elicited prior distribution. The assumption (or hope) is that an informed prior distribution will produce a more accurate estimate of the probability of trial success than a single value estimate. Because clinical trials and other studies are often costly, many institutions now routinely require an elicitation step.

Two reference texts that focus on elicitation are O'Hagan, Buck, Daneshkhah, Eiser, Garthwaite, Jenkinson, Oakley, and Rakow (2006) and Dias, Morton, and Quigley (2018).

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## Technical Details

We assume that a study is to be made comparing the survival (or healing) of a control group with an experimental group. The control group (group 1) consists of patients that will receive the existing treatment. In cases where no existing treatment exists, group 1 consists of patients that will receive a placebo. Group 2 will receive the new treatment.

We assume that the critical event of interest is death and that the two groups have survival distributions with instantaneous death (hazard) rates,  $\lambda_1$  and  $\lambda_2$ . These hazard rates are a subject's probability of death in a short period of time.

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## Cox's Proportional Hazards Regression

Cox's proportional hazards regression is widely used for survival data. The regression model is

$$h(t|z) = h(t|0) \exp(bz)$$

where

$b$  is the regression coefficient which is equal to  $\log[h(t|1)/h(t|0)] = \log(HR)$

$z$  is a binary indicator variable of the treatment group.

$t$  is elapsed time.

$h(t|z)$  is the hazard rate at time  $t$ , given covariate  $z$ .

$HR$  is the hazard ratio  $h(t|1)/h(t|0)$ .

## Equivalence Hypothesis

The equivalence of two hazard rates is established by concluding that their ratio is within a clinically insignificant margin from one. If we define  $EQHR$  as the equivalence hazard ratio, then the statistical hypotheses that yields this conclusion when the null hypothesis is rejected is

$$H_0: |\log(HR)| \geq \log(EQHR) \text{ versus } H_1: |\log(HR)| < \log(EQHR), \text{ assuming } EQHR > 1$$

## Test Statistic

It can be shown that the test of  $b$  based on the partial likelihood method of Cox (1972) coincides with the common logrank test statistic shown next.

## Logrank Test

In this case, the test statistic uses two one-sided logrank tests. Define

$$L_U = \frac{\sum_{k=1}^K \left( I_k - \frac{Y_{1i}EQHR}{Y_{1i}EQHR + Y_{2i}} \right)}{\left[ \sum_{k=1}^K \left( \frac{Y_{1i}Y_{2i}EQHR}{(Y_{1i}EQHR + Y_{2i})^2} \right) \right]^{-\frac{1}{2}}}$$

$$L_L = \frac{\sum_{k=1}^K \left( I_k - \frac{Y_{1i}(1/EQHR)}{Y_{1i}(1/EQHR) + Y_{2i}} \right)}{\left[ \sum_{k=1}^K \left( \frac{Y_{1i}Y_{2i}(1/EQHR)}{(Y_{1i}(1/EQHR) + Y_{2i})^2} \right) \right]^{-\frac{1}{2}}}$$

where  $K$  is the number of deaths,  $Y_{ij}$  is the number of subjects at risk just prior to the  $j^{\text{th}}$  observed event in the  $i^{\text{th}}$  group, and  $I_k$  is a binary variable indicating whether the  $k^{\text{th}}$  event is from group 1 or not.

The null hypothesis is rejected if  $L_U < z_\alpha$  and  $L_L > z_{1-\alpha}$ .

The distribution of  $L_U$  is approximately normal with mean  $(\log(HR) - \log(EQHR))\sqrt{P_1P_2dN}$  and unit variance and the distribution of  $L_L$  is approximately normal with mean  $(\log(HR) - \log(1/EQHR))\sqrt{P_1P_2dN}$  and unit variance, where

$P_1$  is the proportion of  $N$  that is in the control group.

$P_2$  is the proportion of  $N$  that is in the treatment group.

$N$  is the total sample size.

$N_1$  is the sample size from the control group,  $N_1 = N(P_1)$ .

$N_2$  is the sample size from the treatment group,  $N_2 = N(P_2)$ .

$Pev_1$  is probability that the event of interest is observed in the control group during the study.

$Pev_2$  is probability of the event of interest is observed in the treatment group during the study.

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$d$  is the overall probability of an event,  $d = Pev_1P_1 + Pev_2P_2$ .

$b$  is the Cox regression coefficient,  $b = \log(HR)$ .

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## Power Calculations

The power of this test is given by

$$\Phi\left(\left(\log(EQHR) - \log(HR)\right)\sqrt{P_1P_2dN} - z_{1-\alpha}\right) + \Phi\left(\left(\log(EQHR) + \log(HR)\right)\sqrt{P_1P_2dN} - z_{1-\alpha}\right) - 1$$

where  $HR$  is the actual value of the hazard ratio under the alternative hypothesis.

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## Assurance Calculation

This assurance computation described here is based on O'Hagan, Stevens, and Campbell (2005).

Let  $P_1(R|Pev_1, Pev_2, HR)$  be the power function described above where  $R$  is the event that the null hypothesis is rejected conditional on the parameter values. The specification of  $Pev_1$ ,  $Pev_2$ , and  $HR$  is critical to the power calculation, but the actual values are seldom known. Assurance is defined as the expected power where the expectation is with respect to a joint prior distribution for  $Pev_1$ ,  $Pev_2$ , and  $HR$ . Hence, assurance is

$$\begin{aligned} Assurance &= E_{Pev_1, Pev_2, HR}(P_1(R|Pev_1, Pev_2, HR)) \\ &= \iiint P_1(R|Pev_1, Pev_2, HR)f(Pev_1, Pev_2, HR)dPev_1dPev_2dHR \end{aligned}$$

where  $f(Pev_1, Pev_2, HR)$  is the joint prior distribution of  $Pev_1$ ,  $Pev_2$ , and  $HR$ .

In **PASS**, the joint prior distribution can be specified as either a discrete approximation to the joint prior distribution, or as individual prior distributions, one for each parameter.

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## Specifying a Joint Prior Distribution

If the joint prior distribution is to be specified directly, the distribution is specified in **PASS** using a discrete approximation to the function  $f(Pev_1, Pev_2, HR)$ . This provides flexibility in specifying the joint prior distribution. In the three-parameter case, four columns are entered on the spreadsheet: three for the parameters and a fourth for the probability. Each row gives a value for each parameter and the corresponding parameter-combination probability. The accuracy of the distribution approximation is controlled by the number of points (spreadsheet rows) that are used.

An example of entering a joint prior distribution is included at the end of the chapter.

## Specifying Individual Prior Distributions

Ciarleglio, Arendt, and Peduzzi (2016) suggest that more flexibility is available if the joint prior distribution is separated into three independent univariate distributions as follows:

$$f(Pev_1, Pev_2, HR) = f_1(Pev_1)f_2(Pev_2)f_3(HR)$$

where  $f_1(Pev_1)$  is the prior distribution of  $Pev_1$ ,  $f_2(Pev_2)$  is the prior distribution of  $Pev_2$ , and  $f_3(HR)$  is the prior distribution of  $w$ . This method is also available in **PASS**. The definition of assurance becomes

$$\begin{aligned} Assurance &= E_{Pev_1, Pev_2, HR}(P_1(R|Pev_1, Pev_2, HR)) \\ &= \iiint P_1(R|Pev_1, Pev_2, HR)f_1(Pev_1)f_2(Pev_2)f_3(HR)dPev_1dPev_2dHR \end{aligned}$$

Using this definition, the assurance can be calculated using numerical integration. There are a variety of pre-programmed, univariate prior distributions available in **PASS**.

### Fixed Values (No Prior) and Custom Values

For any given parameter, **PASS** also provides the option of entering a single fixed value for the prior distribution, or a series of values and corresponding probabilities (using the spreadsheet), rather than one of the pre-programmed distributions.

### Numerical Integration in PASS (and Notes on Computation Speed)

When the prior distribution is specified as independent univariate distributions, **PASS** uses a numerical integration algorithm to compute the assurance value as follows:

The domain of each prior distribution is divided into  $M$  intervals. Since many of the available prior distributions are unbounded on one (e.g., Gamma) or both (e.g., Normal) ends, an approximation is made to make the domain finite. This is accomplished by truncating the distribution to a domain between the two quantiles:  $q_{0.001}$  and  $q_{0.999}$ .

The value of  $M$  controls the accuracy and speed of the algorithm. If only one parameter is to be given a prior distribution, then a value of  $M$  between 50 and 100 usually gives an accurate result in a timely manner. However, if two parameters are given priors, the number of iterations needed increases from  $M$  to  $M^2$ . For example, if  $M$  is 100, 10000 iterations are needed. Reducing  $M$  from 100 to 50 reduces the number of iterations from 10000 to 2500.

The algorithm runtime increases when searching for sample size rather than solving for assurance, as a search algorithm is employed in this case. When solving for sample size, we recommend reducing  $M$  to 20 or less while exploring various scenarios, and then increasing  $M$  to 50 or more for a final, more accurate, result.

## List of Available Univariate Prior Distributions

This section lists the univariate prior distributions that may be used for any of the applicable parameters when the Prior Entry Method is set to Individual.

### No Prior

If 'No Prior' is chosen for a parameter, the parameter is assumed to take on a single, fixed value with probability one.

### Beta (Shape 1, Shape 2, a, c)

A random variable  $X$  that follows the beta distribution is defined on a finite interval  $[a, c]$ . Two shape parameters ( $\alpha$  and  $\beta$ ) control the shape of this distribution. Two location parameters  $a$  and  $c$  give the minimum and maximum of  $X$ .

The probability density function of the beta distribution is

$$f(x|\alpha, \beta, a, c) = \frac{\left(\frac{x-a}{c-a}\right)^{\alpha-1} \left(\frac{c-x}{c-a}\right)^{\beta-1}}{(c-a)B(\alpha, \beta)}$$

where  $B(\alpha, \beta) = \Gamma(\alpha) \Gamma(\beta) / \Gamma(\alpha + \beta)$  and  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_x = \frac{\alpha c + \beta a}{\alpha + \beta}$$

Various distribution shapes are controlled by the values of  $\alpha$  and  $\beta$ . These include

#### Symmetric and Unimodal

$$\alpha = \beta > 1$$

#### U Shaped

$$\alpha = \beta < 1$$

#### Bimodal

$$\alpha, \beta < 1$$

#### Uniform

$$\alpha = \beta = 1$$

**Parabolic**

$$\alpha = \beta = 2$$

**Bell-Shaped**

$$\alpha = \beta > 2$$

**Gamma (Shape, Scale)**

A random variable  $X$  that follows the gamma distribution is defined on the interval  $(0, \infty)$ . A shape parameter,  $\kappa$ , and a scale parameter,  $\theta$ , control the distribution.

The probability density function of the gamma distribution is

$$f(x|\kappa, \theta) = \frac{x^{\kappa-1} e^{-\frac{x}{\theta}}}{\theta^{\kappa} \Gamma(\kappa)}$$

where  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_x = \frac{\kappa}{\theta}$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

**Inverse-Gamma (Shape, Scale)**

A random variable  $X$  that follows the inverse-gamma distribution is defined on the interval  $(0, \infty)$ . If  $Y \sim \text{gamma}$ , then  $X = 1 / Y \sim \text{inverse-gamma}$ . A shape parameter,  $\alpha$ , and a scale parameter,  $\beta$ , control the distribution.

The probability density function of the inverse-gamma distribution is

$$f(x|\alpha, \beta) = \frac{\beta^{\alpha} x^{\alpha-1} e^{-\frac{\beta}{x}}}{\Gamma(\alpha)}$$

where  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_x = \frac{\beta}{\alpha - 1} \text{ for } \alpha > 1$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Logistic (Location, Scale)

A random variable  $X$  that follows the logistic distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , and a scale parameter,  $s$ , control the distribution.

The probability density function of the logistic distribution is

$$f(x|\mu, s) = \frac{e^{-\frac{x-\mu}{s}}}{s \left(1 + e^{-\frac{x-\mu}{s}}\right)^2}$$

The mean of  $X$  is

$$\mu_X = \mu$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Lognormal (Mean, SD)

A random variable  $X$  that follows the lognormal distribution is defined on the interval  $(0, \infty)$ . A location parameter,  $\mu_{\log(X)}$ , and a scale parameter,  $\sigma_{\log(X)}$ , control the distribution. If  $Z \sim$  standard normal, then  $X = e^{\mu + \sigma Z} \sim$  lognormal. Note that  $\mu_{\log(X)} = E(\log(X))$  and  $\sigma_{\log(X)} = \text{Standard Deviation}(\log(X))$ .

The probability density function of the lognormal distribution is

$$f(x|\mu, \sigma) = \frac{e^{-\frac{1}{2}\left(\frac{\log x - \mu}{\sigma}\right)^2}}{x\sigma\sqrt{2\pi}}$$

The mean of  $X$  is

$$\mu_X = e^{\mu + \frac{\sigma^2}{2}}$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.



## LogT (Mean, SD)

A random variable  $X$  that follows the logT distribution is defined on the interval  $(0, \infty)$ . A location parameter,  $\mu_{\log(X)}$ , a scale parameter,  $\sigma_{\log(X)}$ , and a shape parameter,  $\nu$ , control the distribution. Note that  $\nu$  is referred to as the *degrees of freedom*.

If  $t \sim$  Student's  $t$ , then  $X = e^{\mu + \sigma t} \sim \log T$ .

The probability density function of the logT distribution is

$$f(x|\mu, \sigma, \nu) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{x\Gamma\left(\frac{\nu}{2}\right)\sigma\sqrt{\nu\pi}} \left(1 + \frac{1}{\nu} \left(\frac{\log x - \mu}{\sigma}\right)^2\right)^{\left(\frac{-\nu-1}{2}\right)}$$

The mean of  $X$  is not defined.

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Normal (Mean, SD)

A random variable  $X$  that follows the normal distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , and a scale parameter,  $\sigma$ , control the distribution.

The probability density function of the normal distribution is

$$f(x|\mu, \sigma) = \frac{e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}}{\sigma\sqrt{2\pi}}$$

The mean of  $X$  is

$$\mu_X = \mu$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## T (Mean, SD, DF)

A random variable  $X$  that follows Student's  $t$  distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , a scale parameter,  $\sigma$ , and a shape parameter,  $\nu$ , control the distribution. Note that  $\nu$  is referred to as the *degrees of freedom* or *DF*.

The probability density function of the Student's  $t$  distribution is

$$f(x|\mu, \sigma, \nu) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)\sigma\sqrt{\nu\pi}} \left(1 + \frac{1}{\nu} \left(\frac{x-\mu}{\sigma}\right)^2\right)^{\left(\frac{-\nu-1}{2}\right)}$$

The mean of  $X$  is  $\mu$  if  $\nu > 1$ .

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

### Triangle (Mode, Min, Max)

Let  $a$  = minimum,  $b$  = maximum, and  $c$  = mode. A random variable  $X$  that follows a triangle distribution is defined on the interval  $(a, b)$ .

The probability density function of the triangle distribution is

$$f(x|a, b, c) = \begin{cases} \frac{2(x-a)}{(b-a)(c-a)} & \text{for } a \leq x < c \\ \frac{2}{b-a} & \text{for } x = c \\ \frac{2(b-x)}{(b-a)(b-c)} & \text{for } c < x \leq b \end{cases}$$

The mean of  $X$  is

$$\frac{a + b + c}{3}$$

### Uniform (Min, Max)

Let  $a$  = minimum and  $b$  = maximum. A random variable  $X$  that follows a uniform distribution is defined on the interval  $[a, b]$ .

The probability density function of the uniform distribution is

$$f(x|a, b) = \begin{cases} \frac{1}{b-a} & \text{for } a \leq x \leq b \end{cases}$$

The mean of  $X$  is

$$\frac{a + b}{2}$$

### Weibull (Shape, Scale)

A random variable  $X$  that follows the Weibull distribution is defined on the interval  $(0, \infty)$ . A shape parameter,  $\kappa$ , and a scale parameter,  $\lambda$ , control the distribution.

The probability density function of the Weibull distribution is

$$f(x|\kappa, \lambda) = \frac{\kappa}{\lambda} \left(\frac{x}{\lambda}\right)^{\kappa-1} e^{-\left(\frac{x}{\lambda}\right)^\kappa}$$

The mean of  $X$  is

$$\mu_X = \kappa \Gamma\left(1 + \frac{1}{\kappa}\right)$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(\text{Min} \leq X \leq \text{Max})$  where  $\text{Min}$  and  $\text{Max}$  are two truncation bounds.

## Custom (Values and Probabilities in Spreadsheet)

This custom prior distribution is represented by a set of user-specified points and associated probabilities, entered in two columns of the spreadsheet. The points make up the entire set of values that are used for this parameter in the calculation of assurance. The associated probabilities should sum to one. Note that custom values and probabilities can be used to approximate any continuous distribution.

For example, a prior distribution of X might be

$X_i$	$P_i$
10	0.2
20	0.2
30	0.3
40	0.2
50	0.1

In this example, the mean of X is

$$\mu_X = \sum_{i=1}^5 X_i P_i$$

## Example 1 – Assurance Over a Range of Sample Sizes

A clinical trial is being planned to test the equivalence of a new treatment compared to the current treatment using a significance level of 0.05. Previous studies with similar treatments have shown a hazard ratio of 0.5. The upper equivalence hazard ratio boundary is 1.25. The researchers conjecture that about 55% of the subjects in either group will have experienced the event by the end of the study.

The researchers want to run an assurance analysis for a range of group sample sizes from 200 to 1000. An elicitation exercise determined that the prior distribution of Pev1 is Normal (0.55, 0.05). The elicitation also concluded that prior distribution of Pev2 is Normal (0.55, 0.07). Finally, the prior distribution of HR should be Normal (1, 0.1).

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For .....	<b>Assurance</b>
Prior Entry Method .....	<b>Individual (Enter a prior distribution for each applicable parameter)</b>
Alpha.....	<b>0.05</b>
Group Allocation .....	<b>Equal (N1 = N2)</b>
Sample Size Per Group .....	<b>200 400 600 800 1000</b>
Prior Distribution of Pev1 .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>0.55</b>
SD.....	<b>0.05</b>
Truncation Boundaries.....	<b>None</b>
Prior Distribution of Pev2 .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>0.55</b>
SD.....	<b>0.07</b>
Truncation Boundaries.....	<b>None</b>
Prior Distribution of HR .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>1</b>
SD.....	<b>0.1</b>
Truncation Boundaries.....	<b>None</b>
EQHR (Equivalence Hazard Ratio).....	<b>1.25</b>

#### Options Tab

Number of Computation Points for each .....	<b>50</b>
Prior Distribution	
Maximum N1 in Sample Size Search .....	<b>5000</b>

## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Reports

#### Numeric Results

Solve For: Assurance  
 Groups: 1 = Control, 2 = Treatment  
 Hazard Ratio (HR): Hazard Rate 2 / Hazard Rate 1  
 Hypotheses: H0: HR ≤ 1 / EQHR or HR ≥ EQHR vs. H1: 1 / EQHR < HR < EQHR  
 Test Statistic: Treatment Regression Coefficient from Cox Regression  
 Prior Type: Independent Univariate Distributions

#### Prior Distributions

Pev1: Normal (Mean = 0.55, SD = 0.05).  
 Pev2: Normal (Mean = 0.55, SD = 0.07).  
 HR: Normal (Mean = 1, SD = 0.1).

Assurance*	Power‡	Required Number of Events			Sample Size			Expected Prob Event Grp 1	Expected Prob Event Grp 2	Expected Hazard Ratio E(HR)	Equiv Hazard Ratio EQHR	Alpha
		E1	E2	E	N1	N2	N	E(Pev1)	E(Pev2)			
0.01915	0.00800	110	110	220	200	200	400	0.55	0.55	1	1.25	0.05
0.36684	0.51326	220	220	440	400	400	800	0.55	0.55	1	1.25	0.05
0.54528	0.77809	330	330	660	600	600	1200	0.55	0.55	1	1.25	0.05
0.64661	0.90407	440	440	880	800	800	1600	0.55	0.55	1	1.25	0.05
0.71038	0.96018	550	550	1100	1000	1000	2000	0.55	0.55	1	1.25	0.05

\* The number of points used for computation of the prior(s) was 50.  
 ‡ Power was calculated using Pev1 = E(Pev1) = 0.55, Pev2 = E(Pev2) = 0.55, and HR = E(HR) = 1.

- Assurance The expected power where the expectation is with respect to the prior distribution(s).
- Power The power calculated using the means of the prior distributions as the values of the corresponding parameters.
- E1 The required number of events in group 1, the control group.
- E2 The required number of events in group 2, the treatment group.
- E The total number of events required. E = E1 + E2.
- N1 The number of subjects in group 1, the control group.
- N2 The number of subjects in group 2, the treatment group.
- N The total sample size. N = N1 + N2.
- E(Pev1) The expected value over its prior distribution of the probability of observing an event during the study for a subject in group 1.
- E(Pev2) The expected value over its prior distribution of the probability of observing an event during the study for a subject in group 2.
- E(HR) The expected value over its prior distribution of the hazard ratio.
- EQHR The upper equivalence hazard ratio boundary. The lower equivalence boundary is 1 / EQHR.
- Alpha The probability of rejecting a true null hypothesis.

## Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

**Summary Statements**

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A parallel two-group design will be used to test whether the Group 2 (treatment) hazard rate is equivalent to the Group 1 (control) hazard rate, with hazard ratio equivalence bounds of  $1 / 1.25$  and  $1.25$  ( $H_0: HR \leq 1 / 1.25$  or  $HR \geq 1.25$  versus  $H_1: 1 / 1.25 < HR < 1.25$ ,  $HR = \text{Hazard Rate 2} / \text{Hazard Rate 1}$ ). The comparison will be made using two one-sided, two-sample logrank tests (or appropriate Cox proportional hazards regression coefficient tests), with an overall Type I error rate ( $\alpha$ ) of 0.05. It is assumed that the hazard ratio remains constant throughout the study. The prior distribution used for the probability of observing an event in Group 1 is Normal (Mean = 0.55, SD = 0.05). The prior distribution used for the probability of observing an event in Group 2 is Normal (Mean = 0.55, SD = 0.07). The prior distribution used for the hazard ratio (Hazard Rate 2 / Hazard Rate 1) is Normal (Mean = 1, SD = 0.1). With sample sizes of 200 for Group 1 (control) and 200 for Group 2 (treatment), the assurance (average power) is 0.01915.

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**References**

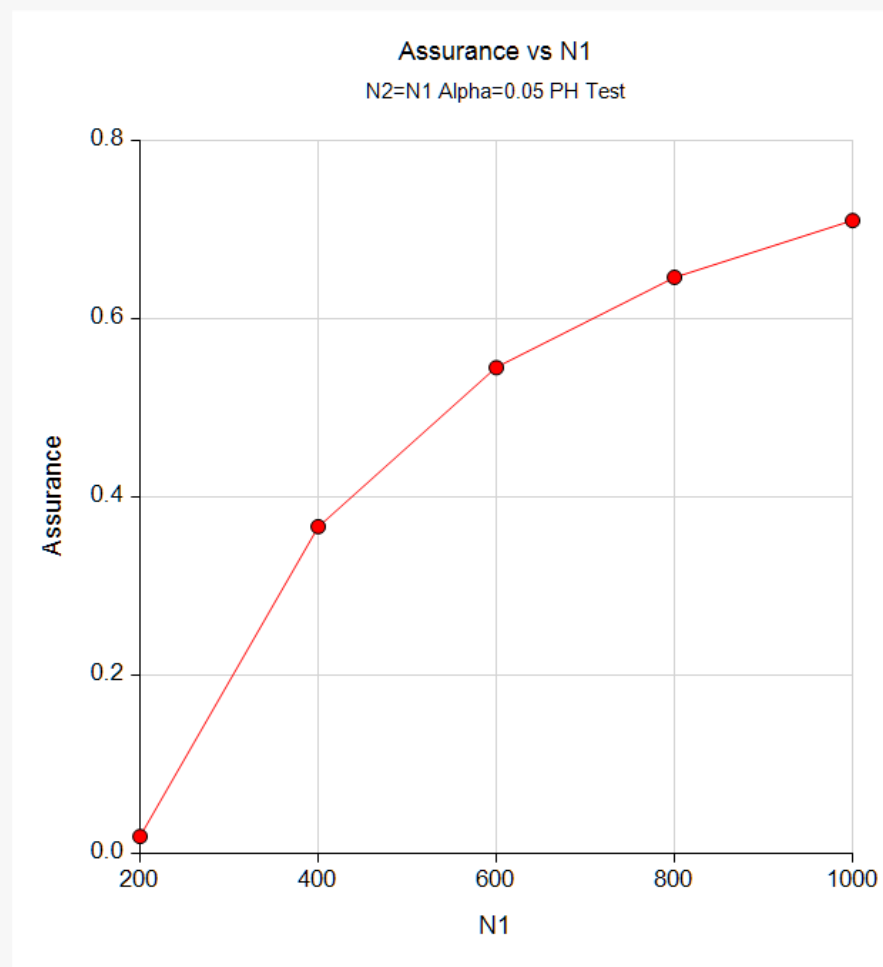
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- 

These reports show the assurance values obtained by each sample size.

## Plots Section

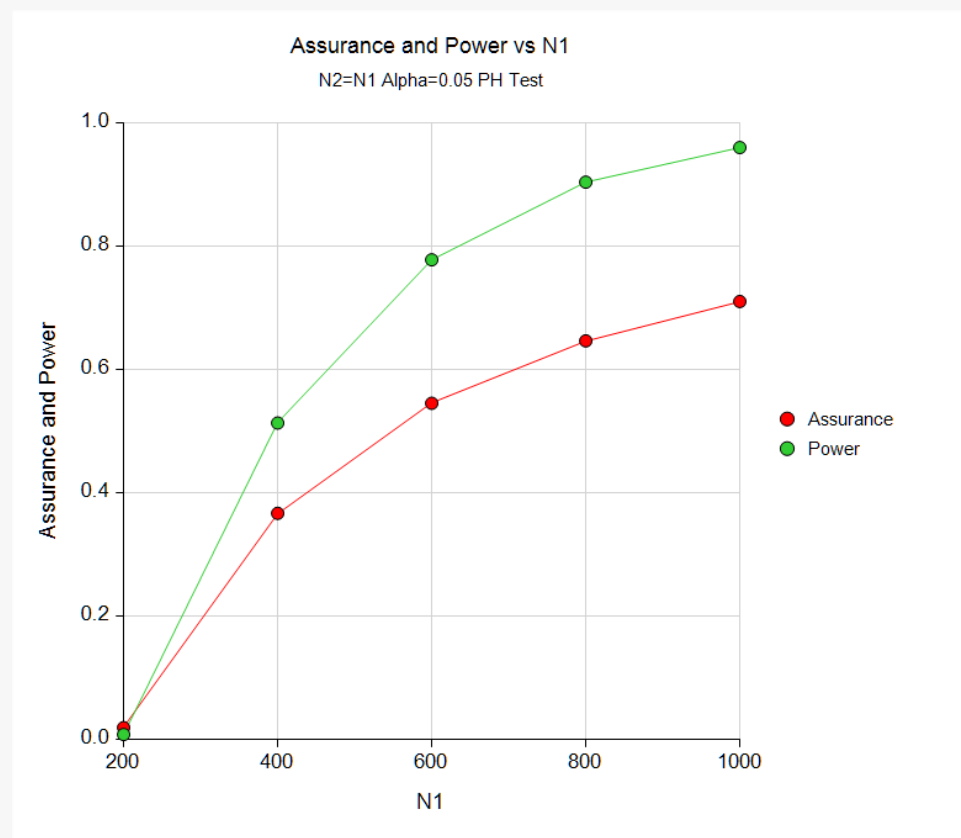
### Plots



This plot shows the relationship between the assurance and sample size. Note the diminishing impact on assurance of each increase in the number of subjects.

## Comparison Plots Section

### Comparison Plots



This plot compares the assurance and power across values of sample size.



## Example 2 – Validation using Hand Computation

We could not find a validation example in the literature for this procedure, so we have developed a validation example of our own.

Suppose an equivalence test based on a Cox regression will be used in which  $N_1 = N_2 = 800$  and the significance level is 0.05.

The prior distribution of  $Pev_1$  will be approximated by the following table.

<u>Pev1</u>	<u>Prob</u>
0.46	0.2
0.5	0.6
0.54	0.2

The prior distribution of the  $Pev_2$  will be approximated by the following table.

<u>Pev2</u>	<u>Prob</u>
0.44	0.2
0.5	0.6
0.56	0.2

The prior distribution of HR will be approximated by the following table.

<u>HR</u>	<u>Prob</u>
0.9	0.3
1.0	0.4
1.1	0.3

The *Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model* procedure is used to compute the power for each of the 27 combinations of  $Pev_1$ ,  $Pev_2$ , and  $HR$ . The results of these calculations are shown next.

### Numeric Results

Solve For: **Power**

Hypotheses:  $H_0: HR \leq 1/HR_0$  or  $HR \geq HR_0$  vs.  $H_a: 1/HR_0 < HR < HR_0$

<b>Power</b>	<b>Total Sample Size N</b>	<b>Control Sample Size N1</b>	<b>Trtmnt Sample Size N2</b>	<b>Percent Control %N1</b>	<b>Actual Hazard Ratio HR1</b>	<b>Equip Hazard Ratio HR0</b>	<b>Control Prob Event Pev1</b>	<b>Trtmnt Prob Event Pev2</b>	<b>Control Events E1</b>	<b>Trtmnt Events E2</b>	<b>Alpha</b>
0.47137	1600	800	800	50	0.9	1.25	0.46	0.44	368	352	0.05
0.82264	1600	800	800	50	1.0	1.25	0.46	0.44	368	352	0.05
0.52369	1600	800	800	50	1.1	1.25	0.46	0.44	368	352	0.05
0.49307	1600	800	800	50	0.9	1.25	0.46	0.50	368	400	0.05
0.85214	1600	800	800	50	1.0	1.25	0.46	0.50	368	400	0.05
0.54749	1600	800	800	50	1.1	1.25	0.46	0.50	368	400	0.05
0.51377	1600	800	800	50	0.9	1.25	0.46	0.56	368	448	0.05
0.87699	1600	800	800	50	1.0	1.25	0.46	0.56	368	448	0.05
0.56996	1600	800	800	50	1.1	1.25	0.46	0.56	368	448	0.05
0.48595	1600	800	800	50	0.9	1.25	0.50	0.44	400	352	0.05
0.84285	1600	800	800	50	1.0	1.25	0.50	0.44	400	352	0.05
0.53972	1600	800	800	50	1.1	1.25	0.50	0.44	400	352	0.05
0.50697	1600	800	800	50	0.9	1.25	0.50	0.50	400	400	0.05
0.86918	1600	800	800	50	1.0	1.25	0.50	0.50	400	400	0.05
0.56261	1600	800	800	50	1.1	1.25	0.50	0.50	400	400	0.05
0.52709	1600	800	800	50	0.9	1.25	0.50	0.56	400	448	0.05

Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

0.89132	1600	800	800	50	1.0	1.25	0.50	0.56	400	448	0.05
0.58429	1600	800	800	50	1.1	1.25	0.50	0.56	400	448	0.05
0.50007	1600	800	800	50	0.9	1.25	0.54	0.44	432	352	0.05
0.86090	1600	800	800	50	1.0	1.25	0.54	0.44	432	352	0.05
0.55512	1600	800	800	50	1.1	1.25	0.54	0.44	432	352	0.05
0.52048	1600	800	800	50	0.9	1.25	0.54	0.50	432	400	0.05
0.88437	1600	800	800	50	1.0	1.25	0.54	0.50	432	400	0.05
0.57719	1600	800	800	50	1.1	1.25	0.54	0.50	432	400	0.05
0.54006	1600	800	800	50	0.9	1.25	0.54	0.56	432	448	0.05
0.90407	1600	800	800	50	1.0	1.25	0.54	0.56	432	448	0.05
0.59814	1600	800	800	50	1.1	1.25	0.54	0.56	432	448	0.05

The assurance calculation is made by summing the quantities  $[(power_{i,j,k})(p(Pev1_i))(p(Pev2_j))(p(HR_k))]$  as follows

$$Assurance = (0.47137 \times 0.2 \times 0.2 \times 0.3) + (0.82264 \times 0.2 \times 0.6 \times 0.3) + \dots + (0.59814 \times 0.2 \times 0.2 \times 0.3) = 0.66788.$$

To run this example, the spreadsheet will need to be loaded with the following six columns in which the first two are for Pev1, the second two are for Pev2, and the third two are for HR.

<b>C1</b>	<b>C2</b>	<b>C3</b>	<b>C4</b>	<b>C5</b>	<b>C6</b>
<b>0.46</b>	<b>0.2</b>	<b>0.44</b>	<b>0.2</b>	<b>0.9</b>	<b>0.3</b>
<b>0.5</b>	<b>0.6</b>	<b>0.5</b>	<b>0.6</b>	<b>1.0</b>	<b>0.4</b>
<b>0.54</b>	<b>0.2</b>	<b>0.56</b>	<b>0.2</b>	<b>1.1</b>	<b>0.3</b>

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Design Tab	
Solve For .....	<b>Assurance</b>
Prior Entry Method.....	<b>Individual (Enter a prior distribution for each applicable parameter)</b>
Alpha.....	<b>0.05</b>
Group Allocation .....	<b>Equal (N1 = N2)</b>
Sample Size Per Group .....	<b>800</b>
Prior Distribution of Pev1 .....	<b>Custom (Values and Probabilities in Spreadsheet)</b>
Column of Values .....	<b>C1</b>
Column of Pr(Values).....	<b>C2</b>
Prior Distribution of Pev2 .....	<b>Custom (Values and Probabilities in Spreadsheet)</b>
Column of Values .....	<b>C3</b>
Column of Pr(Values).....	<b>C4</b>
Prior Distribution of HR .....	<b>Custom (Values and Probabilities in Spreadsheet)</b>
Column of Values .....	<b>C5</b>
Column of Pr(Values).....	<b>C6</b>
EQHR (Equivalence Hazard Ratio).....	<b>1.25</b>

Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

Options Tab

Number of Computation Points for each ..... **50**  
 Prior Distribution  
 Maximum N1 in Sample Size Search ..... **5000**

Input Spreadsheet Data

Row	C1	C2	C3	C4	C5	C6
1	0.46	0.2	0.44	0.2	0.9	0.3
2	0.50	0.6	0.50	0.6	1.0	0.4
3	0.54	0.2	0.56	0.2	1.1	0.3

## Output

Click the Calculate button to perform the calculations and generate the following output.

**Numeric Results**

Solve For: **Assurance**  
 Groups: 1 = Control, 2 = Treatment  
 Hazard Ratio (HR): Hazard Rate 2 / Hazard Rate 1  
 Hypotheses:  $H_0: HR \leq 1 / EQHR$  or  $HR \geq EQHR$  vs.  $H_1: 1 / EQHR < HR < EQHR$   
 Test Statistic: Treatment Regression Coefficient from Cox Regression  
 Prior Type: Independent Univariate Distributions

**Prior Distributions**

Pev1: Point List (Values = C1, Probs = C2).  
 C1: 0.46 0.5 0.54  
 C2: 0.2 0.6 0.2  
 Pev2: Point List (Values = C3, Probs = C4).  
 C3: 0.44 0.5 0.56  
 C4: 0.2 0.6 0.2  
 HR: Point List (Values = C5, Probs = C6).  
 C5: 0.9 1 1.1  
 C6: 0.3 0.4 0.3

Assurance	Power‡	Required Number of Events			Sample Size			Expected Prob Event Grp 1	Expected Prob Event Grp 2	Expected Hazard Ratio E(HR)	Equiv Hazard Ratio EQHR	Alpha
		E1	E2	E	N1	N2	N	E(Pev1)	E(Pev2)			
0.66788	0.86918	400	400	800	800	800	1600	0.5	0.5	1	1.25	0.05

‡ Power was calculated using  $Pev1 = E(Pev1) = 0.5$ ,  $Pev2 = E(Pev2) = 0.5$ , and  $HR = E(HR) = 1$ .

**PASS** has also calculated the assurance as 0.66788 which validates the procedure.

## Example 3 – Finding the Sample Size Needed to Achieve a Specified Assurance

Continuing with Example 1, the researchers want to investigate the sample sizes necessary to achieve assurances of 0.4, 0.6, and 0.8.

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 3** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For .....	<b>Sample Size</b>
Prior Entry Method.....	<b>Individual (Enter a prior distribution for each applicable parameter)</b>
Assurance.....	<b>0.4 0.6 0.8</b>
Alpha.....	<b>0.05</b>
Group Allocation .....	<b>Equal (N1 = N2)</b>
Prior Distribution of Pev1 .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>0.55</b>
SD.....	<b>0.05</b>
Truncation Boundaries.....	<b>None</b>
Prior Distribution of Pev2 .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>0.55</b>
SD.....	<b>0.07</b>
Truncation Boundaries.....	<b>None</b>
Prior Distribution of HR .....	<b>Normal (Mean, SD)</b>
Mean.....	<b>1</b>
SD.....	<b>0.1</b>
Truncation Boundaries.....	<b>None</b>
EQHR (Equivalence Hazard Ratio).....	<b>1.25</b>

#### Options Tab

Number of Computation Points for each.....	<b>50</b>
Prior Distribution	
Maximum N1 in Sample Size Search .....	<b>5000</b>

## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Reports

#### Numeric Results

Solve For: [Sample Size](#)  
 Groups: 1 = Control, 2 = Treatment  
 Hazard Ratio (HR): Hazard Rate 2 / Hazard Rate 1  
 Hypotheses:  $H_0: HR \leq 1 / EQHR$  or  $HR \geq EQHR$  vs.  $H_1: 1 / EQHR < HR < EQHR$   
 Test Statistic: Treatment Regression Coefficient from Cox Regression  
 Prior Type: Independent Univariate Distributions

#### Prior Distributions

Pev1: Normal (Mean = 0.55, SD = 0.05).  
 Pev2: Normal (Mean = 0.55, SD = 0.07).  
 HR: Normal (Mean = 1, SD = 0.1).

Assurance*			Required Number of Events			Sample Size			Expected Prob Event Grp 1	Expected Prob Event Grp 2	Expected Hazard Ratio E(HR)	Equiv Hazard Ratio EQHR	Alpha
Actual	Target	Power‡	E1	E2	E	N1	N2	N	E(Pev1)	E(Pev2)			
0.40030	0.4	0.56393	236	236	472	429	429	858	0.55	0.55	1	1.25	0.05
0.60030	0.6	0.85078	383	383	766	696	696	1392	0.55	0.55	1	1.25	0.05
0.80004	0.8	0.99659	840	840	1680	1527	1527	3054	0.55	0.55	1	1.25	0.05

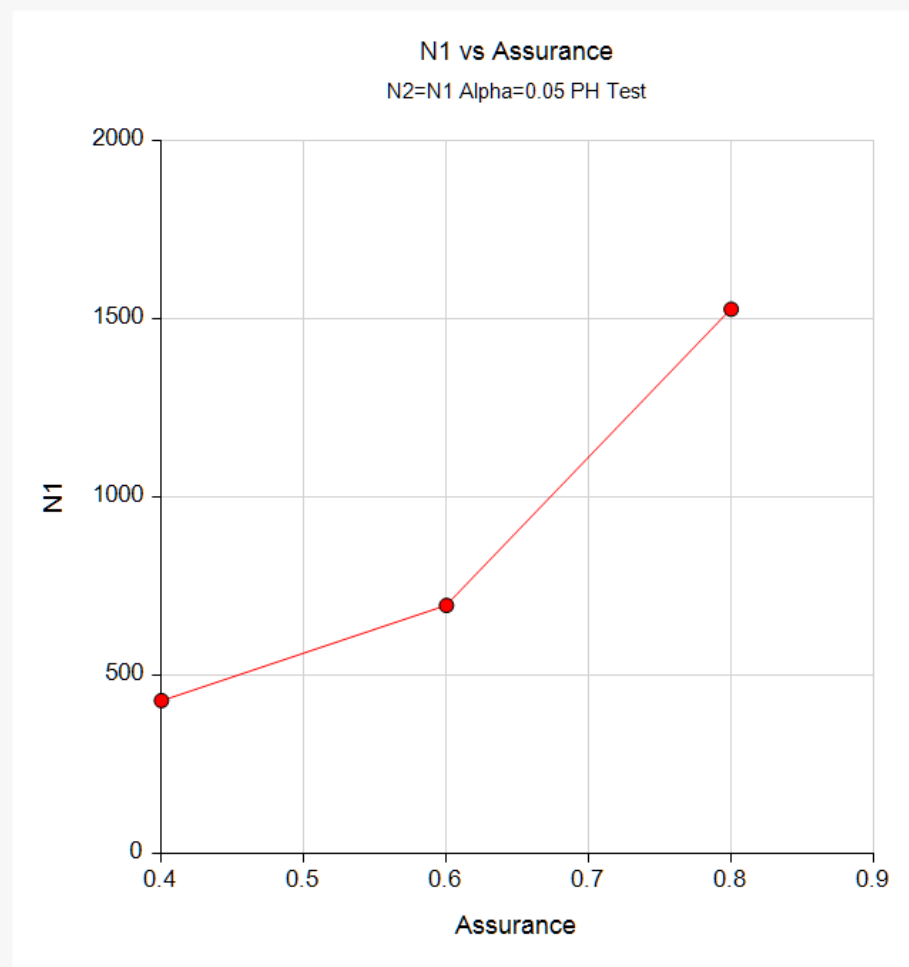
\* The number of points used for computation of the prior(s) was 50.

‡ Power was calculated using  $Pev1 = E(Pev1) = 0.55$ ,  $Pev2 = E(Pev2) = 0.55$ , and  $HR = E(HR) = 1$ .

This report shows the assurance values obtained by each sample size.

## Plots Section

### Plots



This plot shows the relationship between the assurance and sample size.

## Example 4 – Joint Prior Distribution

The following example shows the complexity required to specify a joint distribution for three or more parameters.

Suppose an equivalence test will be used in which  $N_1 = N_2 = (200\ 400\ 600\ 800)$  and the significance level is 0.05. The upper equivalence hazard ratio boundary is set at 1.25. Further suppose that the joint prior distribution of the  $Pev_1$ ,  $Pev_2$ , and HR is approximated by the following table. In a real study, the values in this table would be provided by an elicitation study.

Note that the program will rescale the probabilities so they sum to one.

<u>Pev1</u>	<u>Pev2</u>	<u>HR</u>	<u>Prob</u>
0.60	0.60	0.9	0.2
0.65	0.65	0.9	0.4
0.70	0.70	0.9	0.2
0.63	0.63	0.9	0.4
0.68	0.68	0.9	0.6
0.73	0.73	0.9	0.4
0.66	0.66	0.9	0.2
0.71	0.71	0.9	0.4
0.76	0.76	0.9	0.2
0.60	0.60	1.1	0.1
0.65	0.65	1.1	0.2
0.70	0.70	1.1	0.1
0.63	0.63	1.1	0.2
0.68	0.68	1.1	0.3
0.73	0.73	1.1	0.2
0.66	0.66	1.1	0.1
0.71	0.71	1.1	0.3
0.76	0.76	1.1	0.1

To run this example, the spreadsheet will need to be loaded with the following four columns.

<u>C1</u>	<u>C2</u>	<u>C3</u>	<u>C4</u>
0.60	0.60	0.9	0.2
0.65	0.65	0.9	0.4
0.70	0.70	0.9	0.2
0.63	0.63	0.9	0.4
0.68	0.68	0.9	0.6
0.73	0.73	0.9	0.4
0.66	0.66	0.9	0.2
0.71	0.71	0.9	0.4
0.76	0.76	0.9	0.2
0.60	0.60	1.1	0.1
0.65	0.65	1.1	0.2
0.70	0.70	1.1	0.1
0.63	0.63	1.1	0.2
0.68	0.68	1.1	0.3

## Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

<b>0.73</b>	<b>0.73</b>	<b>1.1</b>	<b>0.2</b>
<b>0.66</b>	<b>0.66</b>	<b>1.1</b>	<b>0.1</b>
<b>0.71</b>	<b>0.71</b>	<b>1.1</b>	<b>0.3</b>
<b>0.76</b>	<b>0.76</b>	<b>1.1</b>	<b>0.1</b>

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 4** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

### Design Tab

Solve For ..... **Assurance**  
 Prior Entry Method ..... **Combined (Enter parameter values and probabilities on spreadsheet)**  
 Alpha ..... **0.05**  
 Group Allocation ..... **Equal (N1 = N2)**  
 Sample Size Per Group ..... **200 400 600 800 1000**  
 Column of Pev1 Values ..... **C1**  
 Column of Pev2 Values ..... **C2**  
 Column of HR Values ..... **C3**  
 Column of Pr(Values)..... **C4**  
 EQHR (Equivalence Hazard Ratio)..... **1.25**

### Options Tab

Number of Computation Points for each ..... **50**  
 Prior Distribution  
 Maximum N1 in Sample Size Search ..... **5000**

### Input Spreadsheet Data

Row	C1	C2	C3	C4
1	0.60	0.60	0.9	0.2
2	0.65	0.65	0.9	0.4
3	0.70	0.70	0.9	0.2
4	0.63	0.63	0.9	0.4
5	0.68	0.68	0.9	0.6
6	0.73	0.73	0.9	0.4
7	0.66	0.66	0.9	0.2
8	0.71	0.71	0.9	0.4
9	0.76	0.76	0.9	0.2
10	0.60	0.60	1.1	0.1
11	0.65	0.65	1.1	0.2
12	0.70	0.70	1.1	0.1
13	0.63	0.63	1.1	0.2
14	0.68	0.68	1.1	0.3
15	0.73	0.73	1.1	0.2
16	0.66	0.66	1.1	0.1
17	0.71	0.71	1.1	0.3
18	0.76	0.76	1.1	0.1



## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Results

Solve For: [Assurance](#)  
 Groups: 1 = Control, 2 = Treatment  
 Hazard Ratio (HR): Hazard Rate 2 / Hazard Rate 1  
 Hypotheses: H0:  $HR \leq 1 / EQHR$  or  $HR \geq EQHR$  vs. H1:  $1 / EQHR < HR < EQHR$   
 Test Statistic: Treatment Regression Coefficient from Cox Regression  
 Prior Type: Joint Multivariate Distribution

### Prior Distribution

Point Lists

Pev1: C1: 0.6 0.65 0.7 0.63 0.68 0.73 0.66 0.71 0.76 0.6 0.65 0.7 0.63 0.68 0.73 0.66 0.71 0.76

Pev2: C2: 0.6 0.65 0.7 0.63 0.68 0.73 0.66 0.71 0.76 0.6 0.65 0.7 0.63 0.68 0.73 0.66 0.71 0.76

HR: C3: 0.9 0.9 0.9 0.9 0.9 0.9 0.9 0.9 0.9 1.1 1.1 1.1 1.1 1.1 1.1 1.1 1.1 1.1

Prob: C4: 0.2 0.4 0.2 0.4 0.6 0.4 0.2 0.4 0.2 0.1 0.2 0.1 0.2 0.3 0.2 0.1 0.3 0.1

Assurance	Power‡	Required Number of Events			Sample Size			Prob Event Grp 1	Expected Prob Event Grp 2	Expected Hazard Ratio E(HR)	Equiv Hazard Ratio EQHR	Alpha
		E1	E2	E	N1	N2	N	E(Pev1)	E(Pev2)			
0.10851	0.15057	137	136	273	200	200	400	0.68065	0.68065	0.96957	1.25	0.05
0.39239	0.63156	273	272	545	400	400	800	0.68065	0.68065	0.96957	1.25	0.05
0.53313	0.84125	409	408	817	600	600	1200	0.68065	0.68065	0.96957	1.25	0.05
0.63801	0.93120	545	545	1090	800	800	1600	0.68065	0.68065	0.96957	1.25	0.05
0.72159	0.97021	681	681	1362	1000	1000	2000	0.68065	0.68065	0.96957	1.25	0.05

‡ Power was calculated using Pev1 = E(Pev1) = 0.68065, Pev2 = E(Pev2) = 0.68065, and HR = E(HR) = 0.96957.

This report shows the assurance values obtained by each sample size.

## Example 5 – Joint Prior Validation

The problem given in Example 2 will be used to validate the joint prior distribution method. This will be done by running the independent-prior scenario used in that example through the joint-prior method and checking that the assurance values match. In Example 2, the prior distributions of the parameters are

<u>Pev1</u>	<u>Prob</u>
0.46	0.2
0.5	0.6
0.54	0.2

<u>Pev2</u>	<u>Prob</u>
0.44	0.2
0.5	0.6
0.56	0.2

<u>HR</u>	<u>Prob</u>
0.9	0.3
1.0	0.4
1.1	0.3

The joint prior distribution can be found by multiplying the three independent probabilities in each row. This results in the following discrete probability distribution.

<u>Pev1</u>	<u>Pev2</u>	<u>HR</u>	<u>Prob</u>
0.46	0.44	0.9	0.012
0.46	0.44	1.0	0.016
0.46	0.44	1.1	0.012
0.46	0.5	0.9	0.036
0.46	0.5	1.0	0.048
0.46	0.5	1.1	0.036
0.46	0.56	0.9	0.012
0.46	0.56	1.0	0.016
0.46	0.56	1.1	0.012
0.5	0.44	0.9	0.036
0.5	0.44	1.0	0.048
0.5	0.44	1.1	0.036
0.5	0.5	0.9	0.108
0.5	0.5	1.0	0.144
0.5	0.5	1.1	0.108
0.5	0.56	0.9	0.036
0.5	0.56	1.0	0.048
0.5	0.56	1.1	0.036
0.54	0.44	0.9	0.012
0.54	0.44	1.0	0.016
0.54	0.44	1.1	0.012
0.54	0.5	0.9	0.036
0.54	0.5	1.0	0.048

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<b>0.54</b>	<b>0.5</b>	<b>1.1</b>	<b>0.036</b>
<b>0.54</b>	<b>0.56</b>	<b>0.9</b>	<b>0.012</b>
<b>0.54</b>	<b>0.56</b>	<b>1.0</b>	<b>0.016</b>
<b>0.54</b>	<b>0.56</b>	<b>1.1</b>	<b>0.012</b>

To run this example, the spreadsheet is loaded with the following four columns.

<b><u>C1</u></b>	<b><u>C2</u></b>	<b><u>C3</u></b>	<b><u>C4</u></b>
<b>0.46</b>	<b>0.44</b>	<b>0.9</b>	<b>0.012</b>
<b>0.46</b>	<b>0.44</b>	<b>1.0</b>	<b>0.016</b>
<b>0.46</b>	<b>0.44</b>	<b>1.1</b>	<b>0.012</b>
<b>0.46</b>	<b>0.5</b>	<b>0.9</b>	<b>0.036</b>
<b>0.46</b>	<b>0.5</b>	<b>1.0</b>	<b>0.048</b>
<b>0.46</b>	<b>0.5</b>	<b>1.1</b>	<b>0.036</b>
<b>0.46</b>	<b>0.56</b>	<b>0.9</b>	<b>0.012</b>
<b>0.46</b>	<b>0.56</b>	<b>1.0</b>	<b>0.016</b>
<b>0.46</b>	<b>0.56</b>	<b>1.1</b>	<b>0.012</b>
<b>0.5</b>	<b>0.44</b>	<b>0.9</b>	<b>0.036</b>
<b>0.5</b>	<b>0.44</b>	<b>1.0</b>	<b>0.048</b>
<b>0.5</b>	<b>0.44</b>	<b>1.1</b>	<b>0.036</b>
<b>0.5</b>	<b>0.5</b>	<b>0.9</b>	<b>0.108</b>
<b>0.5</b>	<b>0.5</b>	<b>1.0</b>	<b>0.144</b>
<b>0.5</b>	<b>0.5</b>	<b>1.1</b>	<b>0.108</b>
<b>0.5</b>	<b>0.56</b>	<b>0.9</b>	<b>0.036</b>
<b>0.5</b>	<b>0.56</b>	<b>1.0</b>	<b>0.048</b>
<b>0.5</b>	<b>0.56</b>	<b>1.1</b>	<b>0.036</b>
<b>0.54</b>	<b>0.44</b>	<b>0.9</b>	<b>0.012</b>
<b>0.54</b>	<b>0.44</b>	<b>1.0</b>	<b>0.016</b>
<b>0.54</b>	<b>0.44</b>	<b>1.1</b>	<b>0.012</b>
<b>0.54</b>	<b>0.5</b>	<b>0.9</b>	<b>0.036</b>
<b>0.54</b>	<b>0.5</b>	<b>1.0</b>	<b>0.048</b>
<b>0.54</b>	<b>0.5</b>	<b>1.1</b>	<b>0.036</b>
<b>0.54</b>	<b>0.56</b>	<b>0.9</b>	<b>0.012</b>
<b>0.54</b>	<b>0.56</b>	<b>1.0</b>	<b>0.016</b>
<b>0.54</b>	<b>0.56</b>	<b>1.1</b>	<b>0.012</b>

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 5** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

### Design Tab

---

Solve For ..... **Assurance**  
 Prior Entry Method ..... **Combined (Enter parameter values and probabilities on spreadsheet)**  
 Alpha ..... **0.05**  
 Group Allocation ..... **Equal (N1 = N2)**  
 Sample Size Per Group ..... **800**  
 Column of Pev1 Values ..... **C1**  
 Column of Pev2 Values ..... **C2**  
 Column of HR Values ..... **C3**  
 Column of Pr(Values)..... **C4**  
 EQHR (Equivalence Hazard Ratio)..... **1.25**

### Options Tab

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Number of Computation Points for each ..... **50**  
 Prior Distribution  
 Maximum N1 in Sample Size Search ..... **5000**

### Input Spreadsheet Data

---

Row	C1	C2	C3	C4
1	0.46	0.44	0.9	0.012
2	0.46	0.44	1.0	0.016
3	0.46	0.44	1.1	0.012
4	0.46	0.50	0.9	0.036
5	0.46	0.50	1.0	0.048
6	0.46	0.50	1.1	0.036
7	0.46	0.56	0.9	0.012
8	0.46	0.56	1.0	0.016
9	0.46	0.56	1.1	0.012
10	0.50	0.44	0.9	0.036
11	0.50	0.44	1.0	0.048
12	0.50	0.44	1.1	0.036
13	0.50	0.50	0.9	0.108
14	0.50	0.50	1.0	0.144
15	0.50	0.50	1.1	0.108
16	0.50	0.56	0.9	0.036
17	0.50	0.56	1.0	0.048
18	0.50	0.56	1.1	0.036
19	0.54	0.44	0.9	0.012
20	0.54	0.44	1.0	0.016
21	0.54	0.44	1.1	0.012
22	0.54	0.50	0.9	0.036
23	0.54	0.50	1.0	0.048

Assurance for Equivalence Tests for Two Survival Curves using Cox's Proportional Hazards Model

24	0.54	0.50	1.1	0.036
25	0.54	0.56	0.9	0.012
26	0.54	0.56	1.0	0.016
27	0.54	0.56	1.1	0.012

## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Results

Solve For: Assurance  
 Groups: 1 = Control, 2 = Treatment  
 Hazard Ratio (HR): Hazard Rate 2 / Hazard Rate 1  
 Hypotheses: H0:  $HR \leq 1 / EQHR$  or  $HR \geq EQHR$  vs. H1:  $1 / EQHR < HR < EQHR$   
 Test Statistic: Treatment Regression Coefficient from Cox Regression  
 Prior Type: Joint Multivariate Distribution

### Prior Distribution

Point Lists

Pev1: C1: 0.46 0.46 0.46 0.46 0.46 0.46 0.46 0.46 0.46 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.5 0.54 0.54 0.54 0.54  
 0.54 0.54 0.54 0.54 0.54  
 Pev2: C2: 0.44 0.44 0.44 0.5 0.5 0.5 0.56 0.56 0.56 0.44 0.44 0.44 0.5 0.5 0.5 0.56 0.56 0.56 0.44 0.44 0.44  
 0.5 0.5 0.5 0.56 0.56 0.56  
 HR: C3: 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1 0.9 1 1.1  
 Prob: C4: 0.012 0.016 0.012 0.036 0.048 0.036 0.012 0.016 0.012 0.036 0.048 0.036 0.108 0.144 0.108  
 0.036 0.048 0.036 0.012 0.016 0.012 0.036 0.048 0.036 0.012 0.016 0.012

Assurance	Power‡	Required Number of Events			Sample Size			Prob Event Grp 1 E(Pev1)	Expected Prob Event Grp 2 E(Pev2)	Expected Hazard Ratio E(HR)	Equiv Hazard Ratio EQHR	Alpha
		E1	E2	E	N1	N2	N					
0.66788	0.86918	400	400	800	800	800	1600	0.5	0.5	1	1.25	0.05

‡ Power was calculated using Pev1 = E(Pev1) = 0.5, Pev2 = E(Pev2) = 0.5, and HR = E(HR) = 1.

PASS has also calculated the assurance as 0.66788 which matches Example 2 and thus validates the procedure.