

## Chapter 727

# Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

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## Introduction

This module calculates the assurance of *non-inferiority* tests for the difference of two hazard rates which assume an exponential model. These results can also be used for a logrank test and a proportional hazards test. The assurance calculation in this procedure is based on a user-specified prior distribution of the applicable parameters. This procedure may also be used to determine the needed sample size to obtain a specified assurance.

The procedure is based on the *unconditional* method of Chow, Shao, and Wang (2008) which, in turn, is based on the *conditional* methods of Lachin and Foulkes (1986). The conditional procedure does not extend to non-inferiority, non-zero null, or equivalence tests as easily as the unconditional method does (see Chow, Shao, and Wang (2008) page 173). The power calculations are based on the **PASS** procedure *Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model*. Refer to that procedure for more details on the test.

The assurance calculation used in this procedure is based on O'Hagan, Stevens, and Campbell (2005).

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## Assurance

The assurance of a design is the expected value of the power with respect to one or more prior distributions of the design parameters. Assurance is also referred to as *Bayesian assurance*, *expected power*, *average power*, *statistical assurance*, *hybrid classical-Bayesian procedure*, or *probability of success*.

The power of a design is the probability of rejecting the null hypothesis, conditional on a given set of design attributes, such as the test statistic, the significance level, the sample size, and the effect size to be detected. As many of the parameters are typically unknown quantities, the stated power may be very different from the true power if the specified parameter values are inaccurate.

While power is conditional on individual design parameter values, and is highly sensitive to those values, assurance is the average power across a presumed prior distribution of the parameters. Thus, assurance adds a Bayesian element to the frequentist framework, resulting in a hybrid approach to the probability of trial or study success. It should be noted that when it comes time to perform the statistical test on the resulting data, these methods for calculating assurance assume that the traditional (frequentist) tests will be used.

The next section describes some of the ways in which the prior distributions for effect size parameters may be determined.

## Elicitation

In order to calculate assurance, a suitable prior distribution for the applicable parameters must be determined. This process is called the *elicitation* of the prior distribution.

The elicitation may be as simple as choosing a distribution that seems plausible for the parameter(s) of interest, or as complex as combining the informed advice of several experts based on experience in the field, available pilot data, or previous studies. The accuracy of the assurance value depends on the accuracy of the elicited prior distribution. The assumption (or hope) is that an informed prior distribution will produce a more accurate estimate of the probability of trial success than a single value estimate. Because clinical trials and other studies are often costly, many institutions now routinely require an elicitation step.

Two reference texts that focus on elicitation are O'Hagan, Buck, Daneshkhah, Eiser, Garthwaite, Jenkinson, Oakley, and Rakow (2006) and Dias, Morton, and Quigley (2018).

## Technical Details

We assume that a study is to be made to compare the hazard rates of a control group and an experimental group using a non-inferiority test. The control group (group 1) consists of patients that will receive the existing treatment. In cases where no existing treatment exists, group 1 consists of patients that will receive a placebo. Group 2 will receive the new treatment.

We assume that the critical event of interest is death and that two treatments have survival distributions with instantaneous death (hazard) rates,  $\lambda_1$  and  $\lambda_2$ . These hazard rates are a subject's probability of death in a short period of time. The survival times are assumed to be exponential. This section presents the *unconditional* method of Chow, Shao, and Wang (2008).

## Basic Model

Suppose a clinical trial consists of two independent groups labeled "1" and "2" (where group 1 is the control group and group 2 is the treatment group). The total sample size is  $N$  and the sizes of the two groups are  $N_1$  and  $N_2$ . Usually, you would plan to have  $N_1 = N_2$ .

## Non-Inferiority Hypothesis

Assuming that lower hazard rates are better, non-inferiority is established by concluding that the treatment hazard rate is at most, only higher than the control hazard rate by a small margin  $\Delta$ . The statistical hypotheses that yields this conclusion when the null hypothesis is rejected is

$$H_0: (\lambda_2 - \lambda_1) \geq \Delta \quad \text{versus} \quad H_1: (\lambda_2 - \lambda_1) < \Delta$$

or

$$H_0: \lambda_2 \geq \lambda_1 + \Delta \quad \text{versus} \quad H_1: \lambda_2 < \lambda_1 + \Delta$$

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If, however, higher hazard rates are better, non-inferiority is established by concluding that the treatment hazard rate is at most, only slightly lower than the control hazard rate. The statistical hypotheses that yields this conclusion when the null hypothesis is rejected is

$$H_0: (\lambda_2 - \lambda_1) \leq -\Delta \quad \text{versus} \quad H_1: (\lambda_2 - \lambda_1) > -\Delta$$

or

$$H_0: \lambda_2 \leq \lambda_1 - \Delta \quad \text{versus} \quad H_1: \lambda_2 > \lambda_1 - \Delta$$

## Test Statistic

The power and sample size formulas presented below are for the difference of two exponential hazard rates. Simulation studies have shown that they also approximate the power of the logrank test. It is anticipated that the actual test statistic is the regression coefficient from a Cox regression.

## Test Comparing Hazard Rates

The original test statistic is the difference of the hazard rates estimated by maximum likelihood divided by their standard error. The maximum likelihood estimate of an exponential hazard rate for a particular group is

$$\hat{\lambda} = \frac{\text{number of events}}{\text{sum of study time of all subjects}}$$

Chow, Shao, and Wang (2008) indicate that the test statistic

$$Z = \frac{(\hat{\lambda}_2 - \hat{\lambda}_1) - \Delta}{\sqrt{\frac{\sigma^2(\hat{\lambda}_1)}{N_1} + \frac{\sigma^2(\hat{\lambda}_2)}{N_2}}}$$

where

$$\sigma^2(\lambda) = \frac{\lambda^2}{1 + \frac{e^{-\lambda T}(1 - e^{\lambda R})}{\lambda R}}$$

follows the standard normal standard normal distribution at least approximately.

## Power Calculations

Assuming an exponential model with hazard rates  $\lambda_1$  and  $\lambda_2$  for the two groups, Chow et al. (2008) give the following equation relating  $N$  and power of a two-tailed test.

$$\frac{(\lambda_2 - \lambda_1) + \Delta}{\sqrt{\frac{\sigma^2(\lambda_1, \omega_1, G)}{N_1} + \frac{\sigma^2(\lambda_2, \omega_2, G)}{N_2}}} - Z_{1-\alpha/2} = Z_{1-\beta}$$

where

$$\sigma^2(\lambda_i, \omega_i, G) = \frac{\lambda_i^2}{E(d_i|\lambda_i, \omega_i, G)}$$

$$E(d_i|\lambda_i, \omega_i, G) = \left( \frac{\lambda_i}{\lambda_i + \omega_i} \right) \left( 1 + \frac{G \exp\{-(\lambda_i + \omega_i)T\} [1 - \exp\{(\lambda_i + \omega_i - G)R\}]}{(\lambda_i + \omega_i - G)[1 - \exp\{-GR\}]} \right)$$

$$E(d_i|\lambda_i, \omega_i, 0) = \left( \frac{\lambda_i}{\lambda_i + \omega_i} \right) \left( 1 + \frac{\exp\{-(\lambda_i + \omega_i)T\} [1 - \exp\{(\lambda_i + \omega_i)R\}]}{(\lambda_i + \omega_i)R} \right)$$

These parameters are interpreted as follows:

<b><u>Parameter</u></b>	<b><u>Interpretation</u></b>
$\sigma^2(\lambda, \omega, G)$	Variance of $\hat{\lambda}$
$E(d_i \lambda_i, \omega_i, G)$	Expected proportion of events (deaths) in group $i$
$d_i$	Indicates a person does ( $d_i = 1$ ) or does not ( $d_i = 0$ ) die in group $i$
$\lambda_i$	Hazard rate of group $i$ (see below)
$\omega_i$	Loss to follow-up hazard rate of group $i$ (see below)
$A$	Patient entry parameter (see below)
$R$	Accrual time
$T$	Total time
$T - R$	Follow-up time

## Exponential Distribution

The hazard rate from the exponential distribution,  $\lambda$ , is usually estimated using maximum likelihood techniques. In the planning stages, you have to obtain an estimate of this parameter. To see how to accomplish this, let's briefly review the exponential distribution. The density function of the exponential is defined as

$$f(t) = \lambda \exp\{-\lambda t\}, \quad t \geq 0, \lambda > 0.$$

The cumulative survival distribution function is

$$S(t) = \exp\{-\lambda t\}, \quad t \geq 0.$$

Solving this for  $\lambda$  yields

$$\lambda = -\frac{\log\{S(t)\}}{t}$$

Note that  $S(t)$  gives the probability of surviving  $t$  years. To obtain a planning estimate of  $\lambda$ , you need only know the proportion surviving during a particular time period. You can then use the above equation to calculate  $\lambda$ .

## Patient Entry

Patients are enrolled during the accrual period. **PASS** lets you specify the pattern in which subjects are enrolled. Suppose patient entry times are distributed as  $g(t)$  where  $t_i$  is the entry time of the  $i^{th}$  individual and  $0 \leq t_i \leq R$ . Let  $g(t)$  follow the truncated exponential distribution with parameter  $G$ , which has the density

$$g(t) = \begin{cases} \frac{G \exp\{-Gt\}}{1 - G \exp\{-GR\}} & \text{if } 0 \leq t \leq R, \quad G \neq 0 \\ 1 & \text{otherwise} \end{cases}$$

where

$R$  is accrual time.

$G$  is interpreted as follows:

$G > 0$  results in a convex (faster than expected) entry distribution.

$G < 0$  results in a concave (slower than expected) entry distribution.

$G = 0$  results in the uniform entry distribution in which  $g(t) = 1/R$ .

Rather than specify  $G$  directly, **PASS** has you enter the percentage of the accrual time (called  $A$ ) that will be needed to enroll 50% of the subjects. Using an iterative search, the value of  $G$  corresponding to this percentage is calculated and used in the calculations.

## Losses to Follow-Up

The staggered patient entry over the accrual period results in censoring times ranging from  $T - R$  to  $T$  years during the follow-up period. This is often referred to as administrative censoring, since it is caused by the conclusion of the study rather than by some random factor working on an individual. To model the losses to follow-up in each group which come from other causes, we use the exponential distribution again, this time with hazard rates  $\omega_1$  and  $\omega_2$ . You can obtain appropriate loss-to-follow-up hazard rates using the following formula or by using the Survival Parameter Conversion Tool available from the Tools menu or by pressing the small button to the rate of the loss-to-follow-up hazard rate box.

$$\omega = -\frac{\log\{1 - P_{loss}(R)\}}{R}$$

## Assurance Calculation

This assurance computation described here is based on O'Hagan, Stevens, and Campbell (2005).

Let  $P_1(H|\lambda_1, \lambda_2, \omega_1, \omega_2, A)$  be the power function described above where  $H$  is the event that the null hypothesis is rejected conditional on a specific set of parameter values. The specification of  $\lambda_1, \lambda_2, \omega_1, \omega_2$ , and  $A$  is critical to the power calculation, but the actual values are seldom known. Assurance is defined as the expected power where the expectation is with respect to a joint prior distribution for  $\lambda_1, \lambda_2, \omega_1, \omega_2$ , and  $A$ . Hence, assurance is

$$\begin{aligned} \text{Assurance} &= E_{\lambda_1, \lambda_2, \omega_1, \omega_2, A}(P_1(H|\lambda_1, \lambda_2, \omega_1, \omega_2, A)) \\ &= \int \int \int \int \int P_1(H|\lambda_1, \lambda_2, \omega_1, \omega_2, A) f(\lambda_1, \lambda_2, \omega_1, \omega_2, A) d\lambda_1 d\lambda_2 d\omega_1 d\omega_2 dA \end{aligned}$$

where  $f(\lambda_1, \lambda_2, \omega_1, \omega_2, A)$  is the joint prior density of  $\lambda_1, \lambda_2, \omega_1, \omega_2$ , and  $A$ .

In **PASS**, the joint prior distribution can be specified as either a discrete approximation to the joint prior distribution, or as individual prior distributions, one for each parameter.

## Specifying a Joint Prior Distribution

If the joint prior distribution is to be specified directly, the distribution is specified in **PASS** using a discrete approximation to the function  $f(\lambda_1, \lambda_2, \omega_1, \omega_2, A)$ . This provides flexibility in specifying the joint prior distribution. In the five-parameter case, six columns are entered on the spreadsheet: five for the parameters and a sixth for the probability. Each row gives a value for each parameter and the corresponding parameter-combination probability. The accuracy of the distribution approximation is controlled by the number of points (spreadsheet rows) that are used.

An example of entering a joint prior distribution is included at the end of the chapter.

## Specifying Individual Prior Distributions

Ciarleglio, Arendt, and Peduzzi (2016) suggest that more flexibility is available if the joint prior distribution is separated into five independent univariate distributions as follows:

$$f(\lambda_1, \lambda_2, \omega_1, \omega_2, A) = f_1(\lambda_1)f_2(\lambda_2)f_3(\omega_1)f_4(\omega_2)f_5(A)$$

where  $f_1(\lambda_1)$  is the prior distribution of  $\lambda_1$ ,  $f_2(\lambda_2)$  is the prior distribution of  $\lambda_2$ , and so on. This is the definition that is used in **PASS**. The definition of assurance becomes

$$\begin{aligned} \text{Assurance} &= E_{\lambda_1, \lambda_2, \omega_1, \omega_2, A}(P_1(H|\lambda_1, \lambda_2, \omega_1, \omega_2, A)) \\ &= \int \int \int \int \int P_1(H|\lambda_1, \lambda_2, \omega_1, \omega_2, A)f_1(\lambda_1)f_2(\lambda_2)f_3(\omega_1)f_4(\omega_2)f_5(A)d\lambda_1d\lambda_2d\omega_1d\omega_2dA \end{aligned}$$

Using this definition, the assurance can be calculated using numerical integration. There are a variety of pre-programmed, univariate prior distributions available in **PASS**.

## Fixed Values (No Prior) and Custom Values

For any given parameter, **PASS** also provides the option of entering a single fixed value for the prior distribution, or a series of values and corresponding probabilities (using the spreadsheet), rather than one of the pre-programmed distributions.

## Numerical Integration in PASS (and Notes on Computation Speed)

When the prior distribution is specified as independent univariate distributions, **PASS** uses a numerical integration algorithm to compute the assurance value as follows:

The domain of each prior distribution is divided into  $M$  intervals. Since many of the available prior distributions are unbounded on one (e.g., Gamma) or both (e.g., Normal) ends, an approximation is made to make the domain finite. This is accomplished by truncating the distribution to a domain between the two quantiles:  $q_{0.001}$  and  $q_{0.999}$ .

The value of  $M$  controls the accuracy and speed of the algorithm. If only one parameter is to be given a prior distribution, then a value of  $M$  between 50 and 100 usually gives an accurate result in a timely manner. However, if two parameters are given priors, the number of iterations needed increases from  $M$  to  $M^2$ . For example, if  $M$  is 100, 10000 iterations are needed. Reducing  $M$  from 100 to 50 reduces the number of iterations from 10000 to 2500.

The algorithm runtime increases when searching for sample size rather than solving for assurance, as a search algorithm is employed in this case. When solving for sample size, we recommend reducing  $M$  to 20 or less while exploring various scenarios, and then increasing  $M$  to 50 or more for a final, more accurate, result.

## List of Available Univariate Prior Distributions

This section lists the univariate prior distributions that may be used for any of the applicable parameters when the Prior Entry Method is set to Individual.

### No Prior

If 'No Prior' is chosen for a parameter, the parameter is assumed to take on a single, fixed value with probability one.

### Beta (Shape 1, Shape 2, a, c)

A random variable  $X$  that follows the beta distribution is defined on a finite interval  $[a, c]$ . Two shape parameters ( $\alpha$  and  $\beta$ ) control the shape of this distribution. Two location parameters  $a$  and  $c$  give the minimum and maximum of  $X$ .

The probability density function of the beta distribution is

$$f(x|\alpha, \beta, a, c) = \frac{\left(\frac{x-a}{c-a}\right)^{\alpha-1} \left(\frac{c-x}{c-a}\right)^{\beta-1}}{(c-a)B(\alpha, \beta)}$$

where  $B(\alpha, \beta) = \Gamma(\alpha) \Gamma(\beta) / \Gamma(\alpha + \beta)$  and  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_X = \frac{\alpha c + \beta a}{\alpha + \beta}$$

Various distribution shapes are controlled by the values of  $\alpha$  and  $\beta$ . These include

#### Symmetric and Unimodal

$$\alpha = \beta > 1$$

#### U Shaped

$$\alpha = \beta < 1$$

#### Bimodal

$$\alpha, \beta < 1$$

#### Uniform

$$\alpha = \beta = 1$$



**Parabolic**

$$\alpha = \beta = 2$$

**Bell-Shaped**

$$\alpha = \beta > 2$$

**Gamma (Shape, Scale)**

A random variable  $X$  that follows the gamma distribution is defined on the interval  $(0, \infty)$ . A shape parameter,  $\kappa$ , and a scale parameter,  $\theta$ , control the distribution.

The probability density function of the gamma distribution is

$$f(x|\kappa, \theta) = \frac{x^{\kappa-1} e^{-\frac{x}{\theta}}}{\theta^{\kappa} \Gamma(\kappa)}$$

where  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_X = \frac{\kappa}{\theta}$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

**Inverse-Gamma (Shape, Scale)**

A random variable  $X$  that follows the inverse-gamma distribution is defined on the interval  $(0, \infty)$ . If  $Y \sim \text{gamma}$ , then  $X = 1 / Y \sim \text{inverse-gamma}$ . A shape parameter,  $\alpha$ , and a scale parameter,  $\beta$ , control the distribution.

The probability density function of the inverse-gamma distribution is

$$f(x|\alpha, \beta) = \frac{\beta^{\alpha} x^{\alpha-1} e^{-\frac{\beta}{x}}}{\Gamma(\alpha)}$$

where  $\Gamma(z)$  is the gamma function.

The mean of  $X$  is

$$\mu_X = \frac{\beta}{\alpha - 1} \text{ for } \alpha > 1$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Logistic (Location, Scale)

A random variable  $X$  that follows the logistic distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , and a scale parameter,  $s$ , control the distribution.

The probability density function of the logistic distribution is

$$f(x|\mu, s) = \frac{e^{-\frac{x-\mu}{s}}}{s \left(1 + e^{-\frac{x-\mu}{s}}\right)^2}$$

The mean of  $X$  is

$$\mu_X = \mu$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Lognormal (Mean, SD)

A random variable  $X$  that follows the lognormal distribution is defined on the interval  $(0, \infty)$ . A location parameter,  $\mu_{\log(X)}$ , and a scale parameter,  $\sigma_{\log(X)}$ , control the distribution. If  $Z \sim$  standard normal, then  $X = e^{\mu + \sigma Z} \sim$  lognormal. Note that  $\mu_{\log(X)} = E(\log(X))$  and  $\sigma_{\log(X)} = \text{Standard Deviation}(\log(X))$ .

The probability density function of the lognormal distribution is

$$f(x|\mu, \sigma) = \frac{e^{-\frac{1}{2}\left(\frac{\log x - \mu}{\sigma}\right)^2}}{x\sigma\sqrt{2\pi}}$$

The mean of  $X$  is

$$\mu_X = e^{\mu + \frac{\sigma^2}{2}}$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## LogT (Mean, SD)

A random variable  $X$  that follows the logT distribution is defined on the interval  $(0, \infty)$ . A location parameter,  $\mu_{\log(X)}$ , a scale parameter,  $\sigma_{\log(X)}$ , and a shape parameter,  $\nu$ , control the distribution. Note that  $\nu$  is referred to as the *degrees of freedom*.

If  $t \sim \text{Student's } t$ , then  $X = e^{\mu + \sigma t} \sim \text{logT}$ .

The probability density function of the logT distribution is

$$f(x|\mu, \sigma, \nu) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{x\Gamma\left(\frac{\nu}{2}\right)\sigma\sqrt{\nu\pi}} \left(1 + \frac{1}{\nu} \left(\frac{\log x - \mu}{\sigma}\right)^2\right)^{\left(\frac{-\nu-1}{2}\right)}$$

The mean of  $X$  is not defined.

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(\text{Min} \leq X \leq \text{Max})$  where  $\text{Min}$  and  $\text{Max}$  are two truncation bounds.

## Normal (Mean, SD)

A random variable  $X$  that follows the normal distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , and a scale parameter,  $\sigma$ , control the distribution.

The probability density function of the normal distribution is

$$f(x|\mu, \sigma) = \frac{e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}}{\sigma\sqrt{2\pi}}$$

The mean of  $X$  is

$$\mu_X = \mu$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(\text{Min} \leq X \leq \text{Max})$  where  $\text{Min}$  and  $\text{Max}$  are two truncation bounds.

## T (Mean, SD, DF)

A random variable  $X$  that follows Student's  $t$  distribution is defined on the interval  $(-\infty, \infty)$ . A location parameter,  $\mu$ , a scale parameter,  $\sigma$ , and a shape parameter,  $\nu$ , control the distribution. Note that  $\nu$  is referred to as the *degrees of freedom* or *DF*.

The probability density function of the Student's  $t$  distribution is

$$f(x|\mu, \sigma, \nu) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)\sigma\sqrt{\nu\pi}} \left(1 + \frac{1}{\nu} \left(\frac{x-\mu}{\sigma}\right)^2\right)^{\left(\frac{-\nu-1}{2}\right)}$$

The mean of  $X$  is  $\mu$  if  $\nu > 1$ .

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(\text{Min} \leq X \leq \text{Max})$  where  $\text{Min}$  and  $\text{Max}$  are two truncation bounds.

### Triangle (Mode, Min, Max)

Let  $a$  = minimum,  $b$  = maximum, and  $c$  = mode. A random variable  $X$  that follows a triangle distribution is defined on the interval  $(a, b)$ .

The probability density function of the triangle distribution is

$$f(x|a, b, c) = \begin{cases} \frac{2(x-a)}{(b-a)(c-a)} & \text{for } a \leq x < c \\ \frac{2}{b-a} & \text{for } x = c \\ \frac{2(b-x)}{(b-a)(b-c)} & \text{for } c < x \leq b \end{cases}$$

The mean of  $X$  is

$$\frac{a + b + c}{3}$$

### Uniform (Min, Max)

Let  $a$  = minimum and  $b$  = maximum. A random variable  $X$  that follows a uniform distribution is defined on the interval  $[a, b]$ .

The probability density function of the uniform distribution is

$$f(x|a, b) = \begin{cases} \frac{1}{b-a} & \text{for } a \leq x \leq b \end{cases}$$

The mean of  $X$  is

$$\frac{a + b}{2}$$

### Weibull (Shape, Scale)

A random variable  $X$  that follows the Weibull distribution is defined on the interval  $(0, \infty)$ . A shape parameter,  $\kappa$ , and a scale parameter,  $\lambda$ , control the distribution.

The probability density function of the Weibull distribution is

$$f(x|\kappa, \lambda) = \frac{\kappa}{\lambda} \left(\frac{x}{\lambda}\right)^{\kappa-1} e^{-\left(\frac{x}{\lambda}\right)^{\kappa}}$$

The mean of  $X$  is

$$\mu_X = \kappa \Gamma\left(1 + \frac{1}{\kappa}\right)$$

A truncated version of the distribution is constructed by dividing the density by  $1 - \text{Prob}(Min \leq X \leq Max)$  where  $Min$  and  $Max$  are two truncation bounds.

## Custom (Values and Probabilities in Spreadsheet)

This custom prior distribution is represented by a set of user-specified points and associated probabilities, entered in two columns of the spreadsheet. The points make up the entire set of values that are used for this parameter in the calculation of assurance. The associated probabilities should sum to one. Note that custom values and probabilities can be used to approximate any continuous distribution.

For example, a prior distribution of X might be

$X_i$	$P_i$
10	0.2
20	0.2
30	0.3
40	0.2
50	0.1

In this example, the mean of X is

$$\mu_X = \sum_{i=1}^5 X_i P_i$$

## Example 1 – Assurance Over a Range of Sample Sizes

A researcher is planning a clinical trial using a parallel, two-group, equal sample allocation design to compare the survivability of a new treatment with that of the current treatment using a one-sided, non-inferiority test with a significance level of 0.025. The hazard rates of the control and treatment groups are expected to be 0.7 and 0.45, respectively.

The trial will include a recruitment period of one-year after which participants will be followed for an additional two-years. It is assumed that patients will enter the study uniformly over the accrual period. The researcher estimates a loss-to-follow hazard rate of 0.1 in both the control and the experimental groups. The non-inferiority margin is set to 0.2.

To complete their sample size study, the researchers want to run an assurance analysis for a range of group sample sizes from 20 to 80. An elicitation exercise determined the prior distributions of the parameters as:  $\lambda_1 \sim N(0.7, 0.05)$ ,  $\lambda_2 \sim N(0.45, 0.05)$ ,  $\omega_1 \sim N(0.1, 0.01)$ ,  $\omega_2 \sim N(0.1, 0.01)$ , and  $A \sim N(50, 3)$ .

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For .....	<b>Assurance</b>
Prior Entry Method .....	<b>Individual (Enter a prior distribution for each applicable parameter)</b>
Higher Hazard Rates Are .....	<b>Worse (H1: <math>\lambda_2 - \lambda_1 &lt; \Delta</math>)</b>
Alpha .....	<b>0.025</b>
Group Allocation .....	<b>Equal (N1 = N2)</b>
Sample Size Per Group .....	<b>20 40 60 80</b>
Prior Distribution of $\omega_1$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.1</b>
SD .....	<b>0.01</b>
Truncation Boundaries .....	<b>None</b>
Prior Distribution of $\omega_2$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.1</b>
SD .....	<b>0.01</b>
Truncation Boundaries .....	<b>None</b>
Accrual Time (R) .....	<b>1</b>
Follow-Up Time (T - R) .....	<b>2</b>
Prior Distribution of A .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>50</b>
SD .....	<b>3</b>
Truncation Boundaries .....	<b>None</b>
Prior Distribution of $\lambda_1$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.7</b>
SD .....	<b>0.05</b>
Truncation Boundaries .....	<b>None</b>

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Prior Distribution of  $\lambda_2$ .....Normal (Mean, SD)

Mean.....0.45

SD.....0.05

Truncation Boundaries.....None

 $\Delta$  (Non-Inferiority Margin).....0.2

## Options Tab

Number of Computation Points for each.....20

Prior Distribution

Maximum N1 in Sample Size Search .....5000

## Output

Click the Calculate button to perform the calculations and generate the following output.

## Numeric Reports

## Numeric Results

Solve For: Assurance  
 Groups: 1 = Control, 2 = Treatment  
 Non-Inferiority Margin ( $\Delta$ ): 0.2  
 Hypotheses:  $H_0: \lambda_2 - \lambda_1 \geq \Delta$  vs.  $H_1: \lambda_2 - \lambda_1 < \Delta$   
 Test Statistic: Hazard Rate Difference,  $\lambda_2 - \lambda_1$   
 Elapsed Times: Accrual = 1, Follow-Up = 2, Total = 3  
 Prior Type: Independent Univariate Distributions

## Prior Distributions

$\lambda_1$ : Normal (Mean = 0.7, SD = 0.05).  
 $\lambda_2$ : Normal (Mean = 0.45, SD = 0.05).  
 $\omega_1$ : Normal (Mean = 0.1, SD = 0.01).  
 $\omega_2$ : Normal (Mean = 0.1, SD = 0.01).  
 A: Normal (Mean = 50, SD = 3).

Assurance*	Power‡	Required Number of Events			Sample Size			Expected Hazard Rate Group 1 $E(\lambda_1)$	Expected Hazard Rate Group 2 $E(\lambda_2)$	Expected Loss Hazard Rate Group 1 $E(\omega_1)$	Expected Loss Hazard Rate Group 2 $E(\omega_2)$	Expected Percent Accrual Time until 50% Accrued $E(A)$	Alpha
		E1	E2	E	N1	N2	N						
0.52645	0.52773	15	12	27	20	20	40	0.7	0.45	0.1	0.1	50	0.025
0.79703	0.81865	30	24	54	40	40	80	0.7	0.45	0.1	0.1	50	0.025
0.91378	0.94006	45	37	82	60	60	120	0.7	0.45	0.1	0.1	50	0.025
0.96204	0.98210	60	49	109	80	80	160	0.7	0.45	0.1	0.1	50	0.025

\* The number of points used for computation of the prior(s) was 20.

‡ Power was calculated using  $\lambda_1 = E(\lambda_1) = 0.7$ ,  $\lambda_2 = E(\lambda_2) = 0.45$ ,  $\omega_1 = E(\omega_1) = 0.1$ ,  $\omega_2 = E(\omega_2) = 0.1$ , and  $A = E(A) = 50$ .

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

Assurance	The expected power where the expectation is with respect to the prior distribution(s).
Power	The power calculated using the means of the prior distributions as the values of the corresponding parameters.
E1	The required number of events in group 1.
E2	The required number of events in group 2.
E	The total number of required events . $E = E1 + E2$ .
N1	The number of subjects in group 1.
N2	The number of subjects in group 2.
N	The total sample size. $N = N1 + N2$ .
$E(\lambda_1)$	The expected hazard rate in group 1.
$E(\lambda_2)$	The expected hazard rate in group 2.
$E(\omega_1)$	The expected hazard rate at which subjects in group 1 are lost to follow-up.
$E(\omega_2)$	The expected hazard rate at which subjects in group 2 are lost to follow-up.
$E(A)$	The expected percent of the accrual time until 50% of the subjects are accrued.
Alpha	The probability of rejecting a true null hypothesis.

### Summary Statements

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A parallel two-group design will be used to test whether the Group 2 (treatment) hazard rate is non-inferior to the Group 1 (control) hazard rate, with a non-inferiority margin of 0.2 ( $H_0: \lambda_2 - \lambda_1 \geq 0.2$  versus  $H_1: \lambda_2 - \lambda_1 < 0.2$ ). The comparison will be made using a one-sided, two-sample Z-test of the hazard rate difference, with a Type I error rate ( $\alpha$ ) of 0.025. It is intended that subjects will enter the study during an accrual period of 1 time unit(s). The follow-up period is 2 time unit(s). It is assumed that the survival times are exponentially distributed. The prior distribution used for the hazard rate in Group 1 is Normal (Mean = 0.7, SD = 0.05). The prior distribution used for the hazard rate in Group 2 is Normal (Mean = 0.45, SD = 0.05). The prior distribution used for the loss hazard rate in Group 1 is Normal (Mean = 0.1, SD = 0.01). The prior distribution used for the loss hazard rate in Group 2 is Normal (Mean = 0.1, SD = 0.01). The prior distribution used for the percent of accrual time until 50% are accrued is Normal (Mean = 50, SD = 3). With sample sizes of 20 for Group 1 (control) and 20 for Group 2 (treatment), the assurance (average power) is 0.52645.

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### References

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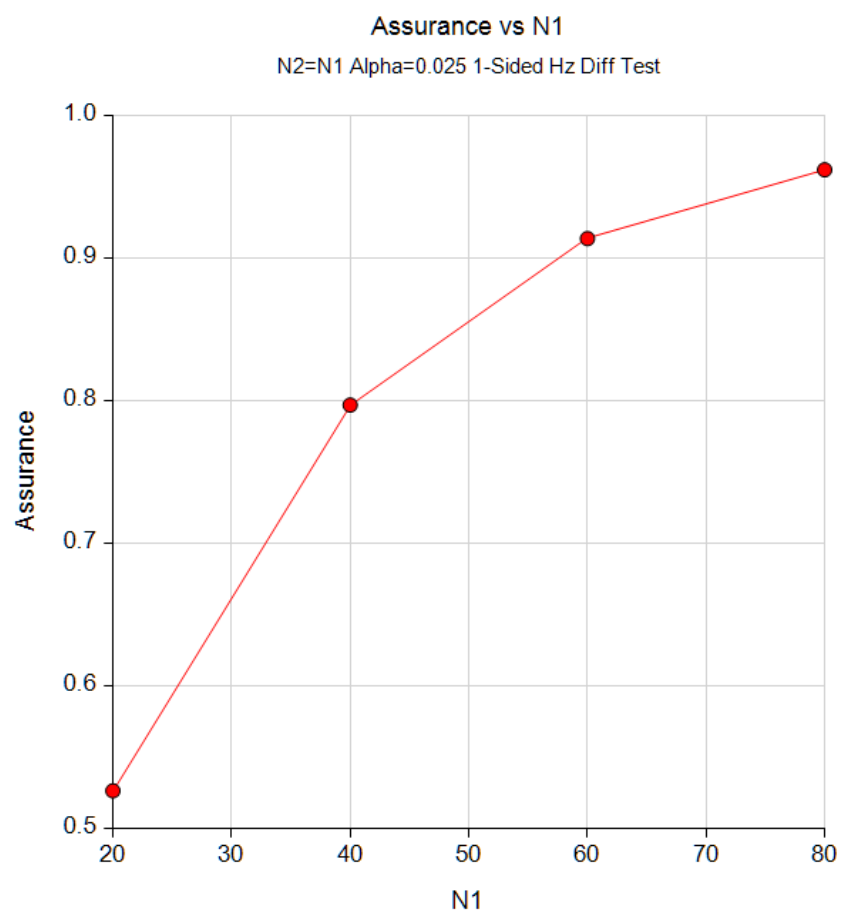
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These reports show the assurance values obtained by each sample size.



## Plots Section

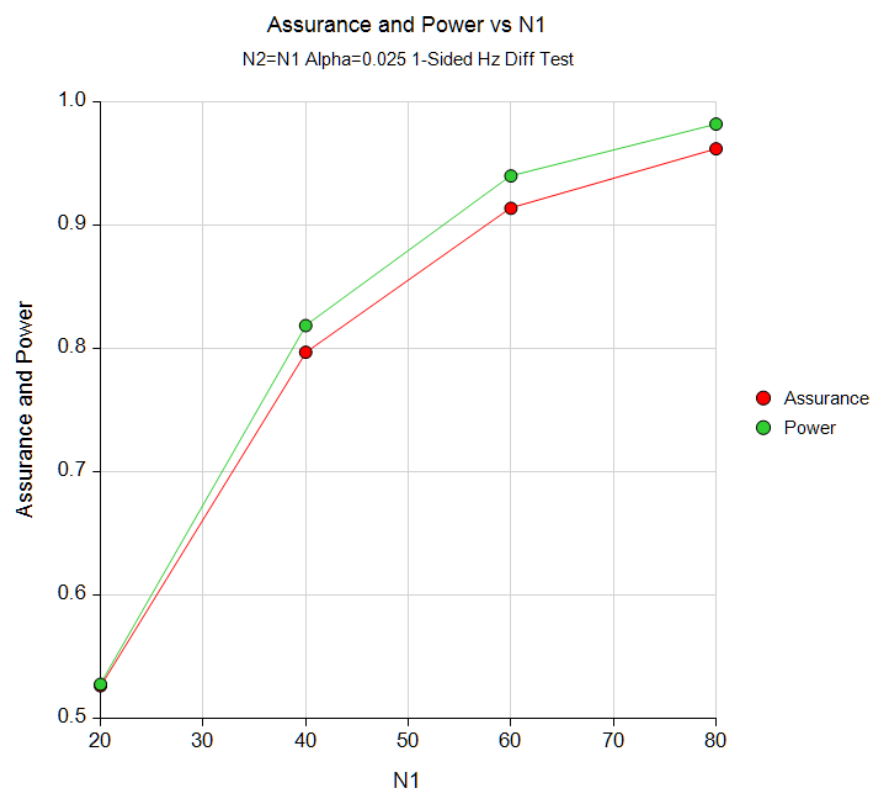
### Plots



This plot shows the relationship between the assurance and sample size.

## Comparison Plots Section

### Comparison Plots



This plot compares the assurance and power across values of sample size.

## Example 2 – Validation using Hand Computation

We could not find a validation example in the literature for this procedure, so we have developed a validation example of our own.

Suppose a non-inferiority test of the difference between hazard rates will be used in which the non-inferiority margin is 0.04,  $N_1 = N_2 = 100$ , and the significance level is 0.025.

The prior distribution of  $\lambda_1$  is approximated by the following table. These are loaded into C1 and C2.

<u><math>\lambda_1</math></u>	<u>Prob</u>
0.6	0.4
0.7	0.6

The prior distribution of the  $\lambda_2$  is approximated by the following table. These are loaded into C3 and C4.

<u><math>\lambda_2</math></u>	<u>Prob</u>
0.3	0.4
0.4	0.6

The prior distribution of the  $\omega_1$  is approximated by the following table. These are loaded into C5 and C6.

<u><math>\omega_1</math></u>	<u>Prob</u>
0.10	0.5
0.16	0.5

The prior distribution of the  $\omega_2$  is approximated by the following table. These are loaded into C7 and C8.

<u><math>\omega_2</math></u>	<u>Prob</u>
0.10	0.5
0.16	0.5

The prior distribution of the  $A$  is approximated by the following table. These are loaded into C9 and C10.

<u><math>A</math></u>	<u>Prob</u>
30	0.5
50	0.5

To run this example, the spreadsheet will need to be loaded with the following ten columns corresponding to the values listed above.

<u>C1</u>	<u>C2</u>	<u>C3</u>	<u>C4</u>	<u>C5</u>	<u>C6</u>	<u>C7</u>	<u>C8</u>	<u>C9</u>	<u>C10</u>
0.6	0.4	0.3	0.4	0.1	0.5	0.1	0.5	30	0.5
0.7	0.6	0.4	0.6	0.16	0.5	0.16	0.5	50	0.5

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

The *Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model* procedure is used to compute the power for each of the 32 parameter combinations. Note that the first report is for  $A = 30$  and the next report is for  $A = 50$  (Uniform).

## Numeric Results

Solve For: **Power**

Groups: 1 = Control, 2 = Treatment

Hypotheses:  $H_0: h_2 \geq h_1 + \Delta$  vs.  $H_a: h_2 < h_1 + \Delta$ 

Accrual: 30% of Accrual Time Results in 50% of Total Enrollment

Power	Sample Size			Hazard Rate		Hazard Rate Difference D	Non-Inferiority		Loss Hazard Rate		Time		Alpha	Report Row
	N	N1	N2	h1	h2		Margin $\Delta$	Boundary B	$\omega_1$	$\omega_2$	Accrual R	Follow-Up T - R		
0.98420	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.10	0.10	1	2	0.025	1
0.77887	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.10	0.10	1	2	0.025	2
0.99807	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.10	0.10	1	2	0.025	3
0.94508	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.10	0.10	1	2	0.025	4
0.98269	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.10	0.16	1	2	0.025	5
0.76984	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.10	0.16	1	2	0.025	6
0.99783	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.10	0.16	1	2	0.025	7
0.94128	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.10	0.16	1	2	0.025	8
0.98062	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.16	0.10	1	2	0.025	9
0.76415	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.16	0.10	1	2	0.025	10
0.99737	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.16	0.10	1	2	0.025	11
0.93747	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.16	0.10	1	2	0.025	12
0.97894	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.16	0.16	1	2	0.025	13
0.75528	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.16	0.16	1	2	0.025	14
0.99708	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.16	0.16	1	2	0.025	15
0.93349	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.16	0.16	1	2	0.025	16

Power	Number of Events			Percent Group 1 %N1	Hazard Ratio HR	Variance		Report Row
	E	E1	E2			$\sigma^2(h_1)$	$\sigma^2(h_2)$	
0.98420	120.8	72.0	48.8	50	0.50000	0.50009	0.18448	1
0.77887	130.4	72.0	58.5	50	0.66667	0.50009	0.27369	2
0.99807	125.5	76.7	48.8	50	0.42857	0.63902	0.18448	3
0.94508	135.1	76.7	58.5	50	0.57143	0.63902	0.27369	4
0.98269	117.7	72.0	45.7	50	0.50000	0.50009	0.19682	5
0.76984	127.0	72.0	55.0	50	0.66667	0.50009	0.29103	6
0.99783	122.4	76.7	45.7	50	0.42857	0.63902	0.19682	7
0.94128	131.7	76.7	55.0	50	0.57143	0.63902	0.29103	8
0.98062	116.9	68.1	48.8	50	0.50000	0.52846	0.18448	9
0.76415	126.6	68.1	58.5	50	0.66667	0.52846	0.27369	10
0.99737	121.6	72.8	48.8	50	0.42857	0.67332	0.18448	11
0.93747	131.2	72.8	58.5	50	0.57143	0.67332	0.27369	12
0.97894	113.8	68.1	45.7	50	0.50000	0.52846	0.19682	13
0.75528	123.1	68.1	55.0	50	0.66667	0.52846	0.29103	14
0.99708	118.5	72.8	45.7	50	0.42857	0.67332	0.19682	15
0.93349	127.8	72.8	55.0	50	0.57143	0.67332	0.29103	16

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

## Numeric Results

Solve For: **Power**  
 Groups: 1 = Control, 2 = Treatment  
 Hypotheses:  $H_0: h_2 \geq h_1 + \Delta$  vs.  $H_a: h_2 < h_1 + \Delta$   
 Accrual: Uniform

Power	Sample Size			Hazard Rate		Hazard Rate Difference D	Non-Inferiority Margin $\Delta$	Non-Inferiority Boundary B	Loss Hazard Rate		Time		Alpha	Report Row
	N	N1	N2	h1	h2				$\omega_1$	$\omega_2$	Accrual R	Follow-Up T - R		
0.98216	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.10	0.10	1	2	0.025	1
0.76937	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.10	0.10	1	2	0.025	2
0.99773	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.10	0.10	1	2	0.025	3
0.94087	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.10	0.10	1	2	0.025	4
0.98057	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.10	0.16	1	2	0.025	5
0.76056	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.10	0.16	1	2	0.025	6
0.99747	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.10	0.16	1	2	0.025	7
0.93702	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.10	0.16	1	2	0.025	8
0.97841	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.16	0.10	1	2	0.025	9
0.75503	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.16	0.10	1	2	0.025	10
0.99697	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.16	0.10	1	2	0.025	11
0.93319	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.16	0.10	1	2	0.025	12
0.97665	200	100	100	0.6	0.3	-0.3	0.04	0.64	0.16	0.16	1	2	0.025	13
0.74640	200	100	100	0.6	0.4	-0.2	0.04	0.64	0.16	0.16	1	2	0.025	14
0.99666	200	100	100	0.7	0.3	-0.4	0.04	0.74	0.16	0.16	1	2	0.025	15
0.92918	200	100	100	0.7	0.4	-0.3	0.04	0.74	0.16	0.16	1	2	0.025	16

Power	Number of Events			Percent Group 1 %N1	Hazard Ratio HR	Variance		Report Row
	E	E1	E2			$\sigma^2(h_1)$	$\sigma^2(h_2)$	
0.98216	117.7	70.5	47.2	50	0.50000	0.51054	0.19058	1
0.76937	127.4	70.5	56.8	50	0.66667	0.51054	0.28149	2
0.99773	122.6	75.3	47.2	50	0.42857	0.65039	0.19058	3
0.94087	132.2	75.3	56.8	50	0.57143	0.65039	0.28149	4
0.98057	114.9	70.5	44.4	50	0.50000	0.51054	0.20277	5
0.76056	124.1	70.5	53.6	50	0.66667	0.51054	0.29860	6
0.99747	119.7	75.3	44.4	50	0.42857	0.65039	0.20277	7
0.93702	128.9	75.3	53.6	50	0.57143	0.65039	0.29860	8
0.97841	114.1	66.9	47.2	50	0.50000	0.53849	0.19058	9
0.75503	123.7	66.9	56.8	50	0.66667	0.53849	0.28149	10
0.99697	118.8	71.6	47.2	50	0.42857	0.68417	0.19058	11
0.93319	128.5	71.6	56.8	50	0.57143	0.68417	0.28149	12
0.97665	111.2	66.9	44.4	50	0.50000	0.53849	0.20277	13
0.74640	120.4	66.9	53.6	50	0.66667	0.53849	0.29860	14
0.99666	116.0	71.6	44.4	50	0.42857	0.68417	0.20277	15
0.92918	125.2	71.6	53.6	50	0.57143	0.68417	0.29860	16

The assurance calculation is made by summing the quantities

$$[(power_{i,j,k,l,m})p(\lambda_{1i})p(\lambda_{2j})p(\omega_{1k})p(\omega_{2l})p(A_m)]$$

as follows

$$\begin{aligned}
 Assurance &= (0.98420 \times 0.4 \times 0.4 \times 0.5 \times 0.5 \times 0.5) + (0.77887 \times 0.4 \times 0.6 \times 0.5 \times 0.5 \times 0.5) + \dots \\
 &\quad + (0.92918 \times 0.6 \times 0.6 \times 0.5 \times 0.5 \times 0.5) \\
 &= 0.91664.
 \end{aligned}$$

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

### Design Tab

Solve For ..... **Assurance**  
 Prior Entry Method ..... **Individual (Enter a prior distribution for each applicable parameter)**  
 Higher Hazard Rates Are ..... **Worse ( $H_1: \lambda_2 - \lambda_1 < \Delta$ )**  
 Alpha ..... **0.025**  
 Group Allocation ..... **Equal ( $N_1 = N_2$ )**  
 Sample Size Per Group ..... **100**  
 Prior Distribution of  $\omega_1$  ..... **Custom (Values and Probabilities in Spreadsheet)**  
 Column of Values ..... **C5**  
 Column of Pr(Values) ..... **C6**  
 Prior Distribution of  $\omega_2$  ..... **Custom (Values and Probabilities in Spreadsheet)**  
 Column of Values ..... **C7**  
 Column of Pr(Values) ..... **C8**  
 Accrual Time (R) ..... **1**  
 Follow-Up Time (T - R) ..... **2**  
 Prior Distribution of A ..... **Custom (Values and Probabilities in Spreadsheet)**  
 Column of Values ..... **C9**  
 Column of Pr(Values) ..... **C10**  
 Prior Distribution of  $\lambda_1$  ..... **Custom (Values and Probabilities in Spreadsheet)**  
 Column of Values ..... **C1**  
 Column of Pr(Values) ..... **C2**  
 Prior Distribution of  $\lambda_2$  ..... **Custom (Values and Probabilities in Spreadsheet)**  
 Column of Values ..... **C3**  
 Column of Pr(Values) ..... **C4**  
 $\Delta$  (Non-Inferiority Margin) ..... **0.04**

### Options Tab

Number of Computation Points for each ..... **20**  
 Prior Distribution  
 Maximum N1 in Sample Size Search ..... **5000**

### Input Spreadsheet Data

Row	C1	C2	C3	C4	C5	C6	C7	C8	C9	C10
1	0.6	0.4	0.3	0.4	0.10	0.5	0.10	0.5	50	0.5
2	0.7	0.6	0.4	0.6	0.16	0.5	0.16	0.5	30	0.5

## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Results

Solve For: Assurance  
 Groups: 1 = Control, 2 = Treatment  
 Non-Inferiority Margin ( $\Delta$ ): 0.04  
 Hypotheses:  $H_0: \lambda_2 - \lambda_1 \geq \Delta$  vs.  $H_1: \lambda_2 - \lambda_1 < \Delta$   
 Test Statistic: Hazard Rate Difference,  $\lambda_2 - \lambda_1$   
 Elapsed Times: Accrual = 1, Follow-Up = 2, Total = 3  
 Prior Type: Independent Univariate Distributions

### Prior Distributions

$\lambda_1$ : Point List (Values = C1, Probs = C2).  
 C1: 0.6 0.7  
 C2: 0.4 0.6  
 $\lambda_2$ : Point List (Values = C3, Probs = C4).  
 C3: 0.3 0.4  
 C4: 0.4 0.6  
 $\omega_1$ : Point List (Values = C5, Probs = C6).  
 C5: 0.1 0.16  
 C6: 0.5 0.5  
 $\omega_2$ : Point List (Values = C7, Probs = C8).  
 C7: 0.1 0.16  
 C8: 0.5 0.5  
 A: Point List (Values = C9, Probs = C10).  
 C9: 50 30  
 C10: 0.5 0.5

Assurance	Power <sup>‡</sup>	Required Number of Events			Sample Size			Expected Hazard Rate Group 1 $E(\lambda_1)$	Expected Hazard Rate Group 2 $E(\lambda_2)$	Expected Loss Hazard Rate Group 1 $E(\omega_1)$	Expected Loss Hazard Rate Group 2 $E(\omega_2)$	Expected Percent Accrual Time until 50% Accrued $E(A)$	Alpha
		E1	E2	E	N1	N2	N						
0.91664	0.95797	72	52	125	100	100	200	0.66	0.36	0.13	0.13	40	0.025

<sup>‡</sup> Power was calculated using  $\lambda_1 = E(\lambda_1) = 0.66$ ,  $\lambda_2 = E(\lambda_2) = 0.36$ ,  $\omega_1 = E(\omega_1) = 0.13$ ,  $\omega_2 = E(\omega_2) = 0.13$ , and  $A = E(A) = 40$ .

PASS has also calculated the assurance as 0.91664 which validates the procedure.

## Example 3 – Finding the Sample Size Needed to Achieve a Specified Assurance

Continuing with Example 1, the researchers want to investigate the sample sizes necessary to achieve assurances of 0.4, 0.6, and 0.8.

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 3** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For .....	<b>Sample Size</b>
Prior Entry Method .....	<b>Individual (Enter a prior distribution for each applicable parameter)</b>
Higher Hazard Rates Are .....	<b>Worse (<math>H_1: \lambda_2 - \lambda_1 &lt; \Delta</math>)</b>
Assurance .....	<b>0.4 0.6 0.8</b>
Alpha .....	<b>0.025</b>
Group Allocation .....	<b>Equal (<math>N_1 = N_2</math>)</b>
Prior Distribution of $\omega_1$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.1</b>
SD .....	<b>0.01</b>
Truncation Boundaries .....	<b>None</b>
Prior Distribution of $\omega_2$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.1</b>
SD .....	<b>0.01</b>
Truncation Boundaries .....	<b>None</b>
Accrual Time (R) .....	<b>1</b>
Follow-Up Time (T - R) .....	<b>2</b>
Prior Distribution of A .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>50</b>
SD .....	<b>3</b>
Truncation Boundaries .....	<b>None</b>
Prior Distribution of $\lambda_1$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.7</b>
SD .....	<b>0.05</b>
Truncation Boundaries .....	<b>None</b>
Prior Distribution of $\lambda_2$ .....	<b>Normal (Mean, SD)</b>
Mean .....	<b>0.45</b>
SD .....	<b>0.05</b>
Truncation Boundaries .....	<b>None</b>
$\Delta$ (Non-Inferiority Margin) .....	<b>0.2</b>



## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

## Options Tab

Number of Computation Points for each ..... **10**

Prior Distribution

Maximum N1 in Sample Size Search ..... **5000**

## Output

Click the Calculate button to perform the calculations and generate the following output.

## Numeric Reports

## Numeric Results

Solve For: [Sample Size](#)  
 Groups: 1 = Control, 2 = Treatment  
 Non-Inferiority Margin ( $\Delta$ ): 0.2  
 Hypotheses:  $H_0: \lambda_2 - \lambda_1 \geq \Delta$  vs.  $H_1: \lambda_2 - \lambda_1 < \Delta$   
 Test Statistic: Hazard Rate Difference,  $\lambda_2 - \lambda_1$   
 Elapsed Times: Accrual = 1, Follow-Up = 2, Total = 3  
 Prior Type: Independent Univariate Distributions

## Prior Distributions

$\lambda_1$ : Normal (Mean = 0.7, SD = 0.05).  
 $\lambda_2$ : Normal (Mean = 0.45, SD = 0.05).  
 $\omega_1$ : Normal (Mean = 0.1, SD = 0.01).  
 $\omega_2$ : Normal (Mean = 0.1, SD = 0.01).  
 A: Normal (Mean = 50, SD = 3).

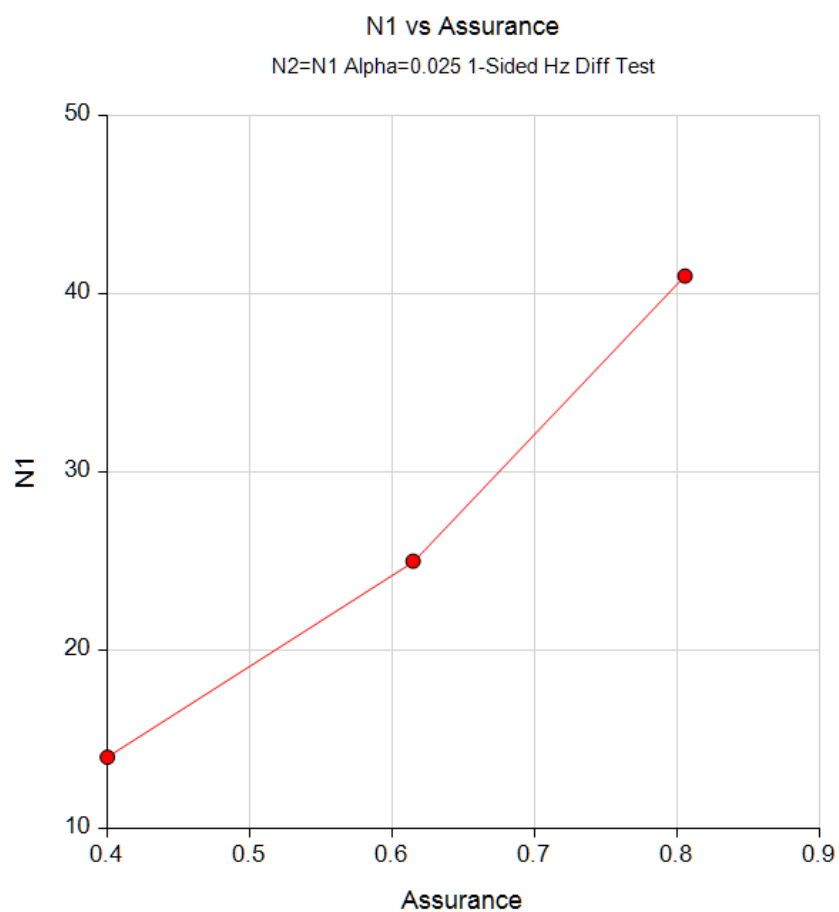
Assurance*	Power‡	Required Number of Events			Sample Size			Expected Hazard Rate Group 1 E( $\lambda_1$ )	Expected Hazard Rate Group 2 E( $\lambda_2$ )	Expected Loss Hazard Rate Group 1 E( $\omega_1$ )	Expected Loss Hazard Rate Group 2 E( $\omega_2$ )	Expected Percent Accrual Time until 50% Accrued E(A)	Alpha
		E1	E2	E	N1	N2	N						
0.40006	0.39668	11	9	19	14	14	28	0.7	0.45	0.1	0.1	50	0.025
0.61462	0.62138	19	15	34	25	25	50	0.7	0.45	0.1	0.1	50	0.025
0.80549	0.82789	31	25	56	41	41	82	0.7	0.45	0.1	0.1	50	0.025

\* The number of points used for computation of the prior(s) was 10.

‡ Power was calculated using  $\lambda_1 = E(\lambda_1) = 0.7$ ,  $\lambda_2 = E(\lambda_2) = 0.45$ ,  $\omega_1 = E(\omega_1) = 0.1$ ,  $\omega_2 = E(\omega_2) = 0.1$ , and  $A = E(A) = 50$ .

## Plots Section

### Plots



This plot shows the relationship between the assurance and sample size.

## Example 4 – Joint Prior Distribution

The following example shows the complexity required to specify a joint distribution for five parameters.

Suppose a non-inferiority test will be used in which  $N_1 = N_2 = (50 \ 100 \ 150 \ 200)$  and the significance level is 0.025. The non-inferiority margin is set at 0.05. Further suppose that the joint prior distribution of the parameters is approximated by the following table. In a real study, the values in this table would be provided by an elicitation study.

Note that the program will rescale the probabilities so they sum to one.

<u><math>\lambda_1</math></u>	<u><math>\lambda_2</math></u>	<u><math>\omega_1</math></u>	<u><math>\omega_2</math></u>	<u>A</u>	<u>Prob</u>
0.6	0.3	0.05	0.05	30	0.07
0.6	0.3	0.1	0.1	30	0.09
0.6	0.3	0.15	0.15	30	0.11
0.65	0.3	0.05	0.05	30	0.07
0.65	0.3	0.1	0.1	30	0.09
0.65	0.3	0.15	0.15	30	0.11
0.7	0.3	0.05	0.05	30	0.07
0.7	0.3	0.1	0.1	30	0.09
0.7	0.3	0.15	0.15	30	0.11
0.6	0.4	0.05	0.05	30	0.27
0.6	0.4	0.1	0.1	30	0.29
0.6	0.4	0.15	0.15	30	0.31
0.65	0.4	0.05	0.05	30	0.27
0.65	0.4	0.1	0.1	30	0.29
0.65	0.4	0.15	0.15	30	0.31
0.7	0.4	0.05	0.05	30	0.27
0.7	0.4	0.1	0.1	30	0.29
0.7	0.4	0.15	0.15	30	0.31
0.6	0.5	0.05	0.05	30	0.17
0.6	0.5	0.1	0.1	30	0.19
0.6	0.5	0.15	0.15	30	0.21
0.65	0.5	0.05	0.05	30	0.17
0.65	0.5	0.1	0.1	30	0.19
0.65	0.5	0.15	0.15	30	0.21
0.7	0.5	0.05	0.05	30	0.17
0.7	0.5	0.1	0.1	30	0.19
0.7	0.5	0.15	0.15	30	0.21
0.6	0.3	0.05	0.05	50	0.12
0.6	0.3	0.1	0.1	50	0.14
0.6	0.3	0.15	0.15	50	0.16
0.65	0.3	0.05	0.05	50	0.12
0.65	0.3	0.1	0.1	50	0.14
0.65	0.3	0.15	0.15	50	0.16
0.7	0.3	0.05	0.05	50	0.12
0.7	0.3	0.1	0.1	50	0.14
0.7	0.3	0.15	0.15	50	0.16

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

0.6	0.4	0.05	0.05	50	0.32
0.6	0.4	0.1	0.1	50	0.34
0.6	0.4	0.15	0.15	50	0.36
0.65	0.4	0.05	0.05	50	0.32
0.65	0.4	0.1	0.1	50	0.34
0.65	0.4	0.15	0.15	50	0.36
0.7	0.4	0.05	0.05	50	0.32
0.7	0.4	0.1	0.1	50	0.34
0.7	0.4	0.15	0.15	50	0.36
0.6	0.5	0.05	0.05	50	0.22
0.6	0.5	0.1	0.1	50	0.24
0.6	0.5	0.15	0.15	50	0.26
0.65	0.5	0.05	0.05	50	0.22
0.65	0.5	0.1	0.1	50	0.24
0.65	0.5	0.15	0.15	50	0.26
0.7	0.5	0.05	0.05	50	0.22
0.7	0.5	0.1	0.1	50	0.24
0.7	0.5	0.15	0.15	50	0.26
0.6	0.3	0.05	0.05	70	0.07
0.6	0.3	0.1	0.1	70	0.09
0.6	0.3	0.15	0.15	70	0.11
0.65	0.3	0.05	0.05	70	0.07
0.65	0.3	0.1	0.1	70	0.09
0.65	0.3	0.15	0.15	70	0.11
0.7	0.3	0.05	0.05	70	0.07
0.7	0.3	0.1	0.1	70	0.09
0.7	0.3	0.15	0.15	70	0.11
0.6	0.4	0.05	0.05	70	0.27
0.6	0.4	0.1	0.1	70	0.29
0.6	0.4	0.15	0.15	70	0.31
0.65	0.4	0.05	0.05	70	0.27
0.65	0.4	0.1	0.1	70	0.29
0.65	0.4	0.15	0.15	70	0.31
0.7	0.4	0.05	0.05	70	0.27
0.7	0.4	0.1	0.1	70	0.29
0.7	0.4	0.15	0.15	70	0.31
0.6	0.5	0.05	0.05	70	0.17
0.6	0.5	0.1	0.1	70	0.19
0.6	0.5	0.15	0.15	70	0.21
0.65	0.5	0.05	0.05	70	0.17
0.65	0.5	0.1	0.1	70	0.19
0.65	0.5	0.15	0.15	70	0.21
0.7	0.5	0.05	0.05	70	0.17
0.7	0.5	0.1	0.1	70	0.19
0.7	0.5	0.15	0.15	70	0.21

To run this example, the above data will need to be loaded into columns C1 to C6.

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 4** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

### Design Tab

Solve For ..... **Assurance**  
 Prior Entry Method ..... **Combined (Enter parameter values and probabilities on spreadsheet)**  
 Higher Hazard Rates Are ..... **Worse (H1:  $\lambda_2 - \lambda_1 < \Delta$ )**  
 Alpha ..... **0.025**  
 Group Allocation ..... **Equal (N1 = N2)**  
 Sample Size Per Group ..... **50 100 150 200**  
 Column of  $\omega_1$  Values ..... **C3**  
 Column of  $\omega_2$  Values ..... **C4**  
 Accrual Time (R) ..... **1**  
 Follow-Up Time (T - R) ..... **2**  
 Column of A Values ..... **C5**  
 Column of  $\lambda_1$  Values ..... **C1**  
 Column of  $\lambda_2$  Values ..... **C2**  
 Column of Pr(Values) ..... **C6**  
 $\Delta$  (Non-Inferiority Margin) ..... **0.05**

### Options Tab

Number of Computation Points for each Prior Distribution    **10**  
 Maximum N1 in Sample Size Search ..... **5000**

### Input Spreadsheet Data

Row	C1	C2	C3	C4	C5	C6
1	0.60	0.3	0.05	0.05	30	0.07
2	0.60	0.3	0.10	0.10	30	0.09
3	0.60	0.3	0.15	0.15	30	0.11
4	0.65	0.3	0.05	0.05	30	0.07
5	0.65	0.3	0.10	0.10	30	0.09
6	0.65	0.3	0.15	0.15	30	0.11
7	0.70	0.3	0.05	0.05	30	0.07
8	0.70	0.3	0.10	0.10	30	0.09
9	0.70	0.3	0.15	0.15	30	0.11
10	0.60	0.4	0.05	0.05	30	0.27
11	0.60	0.4	0.10	0.10	30	0.29
12	0.60	0.4	0.15	0.15	30	0.31
13	0.65	0.4	0.05	0.05	30	0.27
14	0.65	0.4	0.10	0.10	30	0.29
15	0.65	0.4	0.15	0.15	30	0.31
16	0.70	0.4	0.05	0.05	30	0.27
17	0.70	0.4	0.10	0.10	30	0.29
18	0.70	0.4	0.15	0.15	30	0.31
19	0.60	0.5	0.05	0.05	30	0.17
20	0.60	0.5	0.10	0.10	30	0.19
21	0.60	0.5	0.15	0.15	30	0.21
22	0.65	0.5	0.05	0.05	30	0.17

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

23	0.65	0.5	0.10	0.10	30	0.19
24	0.65	0.5	0.15	0.15	30	0.21
25	0.70	0.5	0.05	0.05	30	0.17
26	0.70	0.5	0.10	0.10	30	0.19
27	0.70	0.5	0.15	0.15	30	0.21
28	0.60	0.3	0.05	0.05	50	0.12
29	0.60	0.3	0.10	0.10	50	0.14
30	0.60	0.3	0.15	0.15	50	0.16
31	0.65	0.3	0.05	0.05	50	0.12
32	0.65	0.3	0.10	0.10	50	0.14
33	0.65	0.3	0.15	0.15	50	0.16
34	0.70	0.3	0.05	0.05	50	0.12
35	0.70	0.3	0.10	0.10	50	0.14
36	0.70	0.3	0.15	0.15	50	0.16
37	0.60	0.4	0.05	0.05	50	0.32
38	0.60	0.4	0.10	0.10	50	0.34
39	0.60	0.4	0.15	0.15	50	0.36
40	0.65	0.4	0.05	0.05	50	0.32
41	0.65	0.4	0.10	0.10	50	0.34
42	0.65	0.4	0.15	0.15	50	0.36
43	0.70	0.4	0.05	0.05	50	0.32
44	0.70	0.4	0.10	0.10	50	0.34
45	0.70	0.4	0.15	0.15	50	0.36
46	0.60	0.5	0.05	0.05	50	0.22
47	0.60	0.5	0.10	0.10	50	0.24
48	0.60	0.5	0.15	0.15	50	0.26
49	0.65	0.5	0.05	0.05	50	0.22
50	0.65	0.5	0.10	0.10	50	0.24
51	0.65	0.5	0.15	0.15	50	0.26
52	0.70	0.5	0.05	0.05	50	0.22
53	0.70	0.5	0.10	0.10	50	0.24
54	0.70	0.5	0.15	0.15	50	0.26
55	0.60	0.3	0.05	0.05	70	0.07
56	0.60	0.3	0.10	0.10	70	0.09
57	0.60	0.3	0.15	0.15	70	0.11
58	0.65	0.3	0.05	0.05	70	0.07
59	0.65	0.3	0.10	0.10	70	0.09
60	0.65	0.3	0.15	0.15	70	0.11
61	0.70	0.3	0.05	0.05	70	0.07
62	0.70	0.3	0.10	0.10	70	0.09
63	0.70	0.3	0.15	0.15	70	0.11
64	0.60	0.4	0.05	0.05	70	0.27
65	0.60	0.4	0.10	0.10	70	0.29
66	0.60	0.4	0.15	0.15	70	0.31
67	0.65	0.4	0.05	0.05	70	0.27
68	0.65	0.4	0.10	0.10	70	0.29
69	0.65	0.4	0.15	0.15	70	0.31
70	0.70	0.4	0.05	0.05	70	0.27
71	0.70	0.4	0.10	0.10	70	0.29
72	0.70	0.4	0.15	0.15	70	0.31
73	0.60	0.5	0.05	0.05	70	0.17
74	0.60	0.5	0.10	0.10	70	0.19
75	0.60	0.5	0.15	0.15	70	0.21
76	0.65	0.5	0.05	0.05	70	0.17
77	0.65	0.5	0.10	0.10	70	0.19
78	0.65	0.5	0.15	0.15	70	0.21
79	0.70	0.5	0.05	0.05	70	0.17
80	0.70	0.5	0.10	0.10	70	0.19
81	0.70	0.5	0.15	0.15	70	0.21

## Numeric Results

### Prior Distribution

## Point Lists

[illegible]

Assurance	Power†	Required Number of Events			Sample Size			Expected Hazard Rate Group 1 E( $\lambda_1$ )	Expected Hazard Rate Group 2 E( $\lambda_2$ )	Expected Loss Hazard Rate Group 1 E( $\omega_1$ )	Expected Loss Hazard Rate Group 2 E( $\omega_2$ )	Expected Percent of Accrual Time until 50% are Accrued E(A)	Alpha
		E1	E2	E	N1	N2	N						
0.56131	0.57200	36	29	65	50	50	100	0.65	0.41613	0.10323	0.10323	50	0.025
0.78216	0.85735	73	58	131	100	100	200	0.65	0.41613	0.10323	0.10323	50	0.025
0.87701	0.95986	109	87	196	150	150	300	0.65	0.41613	0.10323	0.10323	50	0.025
0.92436	0.98991	146	116	262	200	200	400	0.65	0.41613	0.10323	0.10323	50	0.025

This report shows the assurance values obtained by each sample size.

## Example 5 – Joint Prior Validation

The problem given in Example 2 will be used to validate the joint prior distribution method. This will be done by running the independent-prior scenario used in that example through the joint-prior method and checking that the assurance values match.

In Example 2, the prior distributions of the parameters are

$\lambda_1$	$P(\lambda_1)$	$\lambda_2$	$P(\lambda_2)$	$\omega_1$	$P(\omega_1)$	$\omega_2$	$P(\omega_2)$	A	$P(A)$
0.6	0.4	0.3	0.4	0.1	0.5	0.1	0.5	30	0.5
0.7	0.6	0.4	0.6	0.16	0.5	0.16	0.5	50	0.5

The joint prior distribution can be found by multiplying the five independent probabilities in each row. This results in the following discrete joint probability distribution.

$\lambda_1$	$\lambda_2$	$\omega_1$	$\omega_2$	A	Prob
0.6	0.3	0.1	0.1	30	0.02
0.6	0.4	0.1	0.1	30	0.03
0.7	0.3	0.1	0.1	30	0.03
0.7	0.4	0.1	0.1	30	0.045
0.6	0.3	0.1	0.16	30	0.02
0.6	0.4	0.1	0.16	30	0.03
0.7	0.3	0.1	0.16	30	0.03
0.7	0.4	0.1	0.16	30	0.045
0.6	0.3	0.16	0.1	30	0.02
0.6	0.4	0.16	0.1	30	0.03
0.7	0.3	0.16	0.1	30	0.03
0.7	0.4	0.16	0.1	30	0.045
0.6	0.3	0.16	0.16	30	0.02
0.6	0.4	0.16	0.16	30	0.03
0.7	0.3	0.16	0.16	30	0.03
0.7	0.4	0.16	0.16	30	0.045
0.6	0.3	0.1	0.1	50	0.02
0.6	0.4	0.1	0.1	50	0.03
0.7	0.3	0.1	0.1	50	0.03
0.7	0.4	0.1	0.1	50	0.045
0.6	0.3	0.1	0.16	50	0.02
0.6	0.4	0.1	0.16	50	0.03
0.7	0.3	0.1	0.16	50	0.03
0.7	0.4	0.1	0.16	50	0.045
0.6	0.3	0.16	0.1	50	0.02
0.6	0.4	0.16	0.1	50	0.03
0.7	0.3	0.16	0.1	50	0.03
0.7	0.4	0.16	0.1	50	0.045
0.6	0.3	0.16	0.16	50	0.02
0.6	0.4	0.16	0.16	50	0.03
0.7	0.3	0.16	0.16	50	0.03
0.7	0.4	0.16	0.16	50	0.045



## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

To run this example, the spreadsheet is loaded with the following six columns.

<b><u>C1</u></b>	<b><u>C2</u></b>	<b><u>C3</u></b>	<b><u>C4</u></b>	<b><u>C5</u></b>	<b><u>C6</u></b>
0.6	0.3	0.1	0.1	30	0.02
0.6	0.4	0.1	0.1	30	0.03
0.7	0.3	0.1	0.1	30	0.03
0.7	0.4	0.1	0.1	30	0.045
0.6	0.3	0.1	0.16	30	0.02
0.6	0.4	0.1	0.16	30	0.03
0.7	0.3	0.1	0.16	30	0.03
0.7	0.4	0.1	0.16	30	0.045
0.6	0.3	0.16	0.1	30	0.02
0.6	0.4	0.16	0.1	30	0.03
0.7	0.3	0.16	0.1	30	0.03
0.7	0.4	0.16	0.1	30	0.045
0.6	0.3	0.16	0.16	30	0.02
0.6	0.4	0.16	0.16	30	0.03
0.7	0.3	0.16	0.16	30	0.03
0.7	0.4	0.16	0.16	30	0.045
0.6	0.3	0.1	0.1	50	0.02
0.6	0.4	0.1	0.1	50	0.03
0.7	0.3	0.1	0.1	50	0.03
0.7	0.4	0.1	0.1	50	0.045
0.6	0.3	0.1	0.16	50	0.02
0.6	0.4	0.1	0.16	50	0.03
0.7	0.3	0.1	0.16	50	0.03
0.7	0.4	0.1	0.16	50	0.045
0.6	0.3	0.16	0.1	50	0.02
0.6	0.4	0.16	0.1	50	0.03
0.7	0.3	0.16	0.1	50	0.03
0.7	0.4	0.16	0.1	50	0.045
0.6	0.3	0.16	0.16	50	0.02
0.6	0.4	0.16	0.16	50	0.03
0.7	0.3	0.16	0.16	50	0.03
0.7	0.4	0.16	0.16	50	0.045

## Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 5** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

### Design Tab

Solve For ..... **Assurance**  
 Prior Entry Method ..... **Combined (Enter parameter values and probabilities on spreadsheet)**  
 Higher Hazard Rates Are ..... **Worse (H1:  $\lambda_2 - \lambda_1 < \Delta$ )**  
 Alpha ..... **0.025**  
 Group Allocation ..... **Equal (N1 = N2)**  
 Sample Size Per Group ..... **100**  
 Column of  $\omega_1$  Values ..... **C3**  
 Column of  $\omega_2$  Values ..... **C4**  
 Accrual Time (R) ..... **1**  
 Follow-Up Time (T - R) ..... **2**  
 Column of A Values ..... **C5**  
 Column of  $\lambda_1$  Values ..... **C1**  
 Column of  $\lambda_2$  Values ..... **C2**  
 Column of Pr(Values) ..... **C6**  
 $\Delta$  (Non-Inferiority Margin) ..... **0.04**

### Options Tab

Number of Computation Points for each ..... **20**  
 Prior Distribution  
 Maximum N1 in Sample Size Search ..... **5000**

### Input Spreadsheet Data

Row	C1	C2	C3	C4	C5	C6
1	0.6	0.3	0.10	0.10	30	0.020
2	0.6	0.4	0.10	0.10	30	0.030
3	0.7	0.3	0.10	0.10	30	0.030
4	0.7	0.4	0.10	0.10	30	0.045
5	0.6	0.3	0.10	0.16	30	0.020
6	0.6	0.4	0.10	0.16	30	0.030
7	0.7	0.3	0.10	0.16	30	0.030
8	0.7	0.4	0.10	0.16	30	0.045
9	0.6	0.3	0.16	0.10	30	0.020
10	0.6	0.4	0.16	0.10	30	0.030
11	0.7	0.3	0.16	0.10	30	0.030
12	0.7	0.4	0.16	0.10	30	0.045
13	0.6	0.3	0.16	0.16	30	0.020
14	0.6	0.4	0.16	0.16	30	0.030
15	0.7	0.3	0.16	0.16	30	0.030
16	0.7	0.4	0.16	0.16	30	0.045
17	0.6	0.3	0.10	0.10	50	0.020
18	0.6	0.4	0.10	0.10	50	0.030
19	0.7	0.3	0.10	0.10	50	0.030
20	0.7	0.4	0.10	0.10	50	0.045
21	0.6	0.3	0.10	0.16	50	0.020

## Assurance for Non-Inferiority Tests for the Difference of Two Hazard Rates Assuming an Exponential Model

22	0.6	0.4	0.10	0.16	50	0.030
23	0.7	0.3	0.10	0.16	50	0.030
24	0.7	0.4	0.10	0.16	50	0.045
25	0.6	0.3	0.16	0.10	50	0.020
26	0.6	0.4	0.16	0.10	50	0.030
27	0.7	0.3	0.16	0.10	50	0.030
28	0.7	0.4	0.16	0.10	50	0.045
29	0.6	0.3	0.16	0.16	50	0.020
30	0.6	0.4	0.16	0.16	50	0.030
31	0.7	0.3	0.16	0.16	50	0.030
32	0.7	0.4	0.16	0.16	50	0.045

## Output

Click the Calculate button to perform the calculations and generate the following output.

### Numeric Results

Solve For: [Assurance](#)  
 Groups: 1 = Control, 2 = Treatment  
 Non-Inferiority Margin ( $\Delta$ ): 0.04  
 Hypotheses:  $H_0: \lambda_2 - \lambda_1 \geq \Delta$  vs.  $H_1: \lambda_2 - \lambda_1 < \Delta$   
 Test Statistic: Hazard Rate Difference,  $\lambda_2 - \lambda_1$   
 Elapsed Times: Accrual = 1, Follow-Up = 2, Total = 3  
 Prior Type: Joint Multivariate Distribution

### Prior Distribution

#### Point Lists

$\lambda_1$ : C1: 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7 0.6 0.6 0.7 0.7  
 $\lambda_2$ : C2: 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4 0.3 0.4  
 $\omega_1$ : C3: 0.1 0.1 0.1 0.1 0.1 0.1 0.1 0.1 0.16 0.16 0.16 0.16 0.16 0.16 0.16 0.16 0.1 0.1 0.1 0.1 0.1 0.1 0.16 0.16 0.16 0.16 0.16 0.16  
 $\omega_2$ : C4: 0.1 0.1 0.1 0.1 0.16 0.16 0.16 0.16 0.1 0.1 0.1 0.1 0.16 0.16 0.16 0.16 0.1 0.1 0.1 0.1 0.16 0.16 0.16 0.16 0.1 0.1 0.1 0.16 0.16  
 A: C5: 30 30 30 30 30 30 30 30 30 30 30 30 30 30 50 50 50 50 50 50 50 50 50 50 50 50 50 50  
 Prob: C6: 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045  
 0.02 0.03 0.03 0.045 0.02 0.03 0.03 0.045

Assurance	Power†	Required Number of Events			Sample Size			Expected Hazard Rate Group 1 $E(\lambda_1)$	Expected Hazard Rate Group 2 $E(\lambda_2)$	Expected Loss Hazard Rate Group 1 $E(\omega_1)$	Expected Loss Hazard Rate Group 2 $E(\omega_2)$	Expected Percent of Accrual Time until 50% are Accrued $E(A)$	Alpha
		E1	E2	E	N1	N2	N						
0.91664	0.95797	72	52	125	100	100	200	0.66	0.36	0.13	0.13	40	0.025

† Power was calculated using  $\lambda_1 = E(\lambda_1) = 0.66$ ,  $\lambda_2 = E(\lambda_2) = 0.36$ ,  $\omega_1 = E(\omega_1) = 0.13$ ,  $\omega_2 = E(\omega_2) = 0.13$ , and  $A = E(A) = 40$ .

PASS has also calculated the assurance as 0.91664 which matches Example 2 and thus validates the procedure.