Chapter 228

Group-Sequential Non-Inferiority Tests for the Odds Ratio of Two Proportions (Simulation)

This procedure uses simulation for the calculation of the boundaries as well as for calculation of power (and sample size). Futility boundaries are limited. A variety of test statistics are available.

Introduction

This procedure can be used to determine power, sample size and/or boundaries for group sequential non-inferiority tests comparing the odds ratio of proportions from two groups. The tests that can be simulated in this procedure are the common two-sample Z-test with or without pooled standard error and with or without continuity correction, the T-test, and three score tests. Significance and futility boundaries can be produced. The spacing of the looks can be equal or custom specified. Boundaries can be computed based on popular alpha- and beta-spending functions (O'Brien-Fleming, Pocock, Hwang-Shih-DeCani Gamma family, linear) or custom spending functions. Boundaries can also be input directly to verify alpha- and/or beta-spending properties. Futility boundaries can be binding or non-binding. Maximum and average (expected) sample sizes are reported as well as the alpha and/or beta spent and incremental power at each look. Corresponding P-Value boundaries are given for each boundary statistic. Plots of boundaries are also produced.

Technical Details

This section outlines many of the technical details of the techniques used in this procedure including the simulation summary, the test statistic details, and the use of spending functions.

An excellent text for the background and details of many group-sequential methods is Jennison and Turnbull (2000).

Simulation Procedure

In this procedure, a large number of simulations are used to calculate boundaries and power using the following steps

- 1. Based on the specified proportions, random samples of size N1 and N2 are generated under the null distribution and under the alternative distribution. These are simulated samples as though the final look is reached.
- 2. For each sample, test statistics for each look are produced. For example, if N1 and N2 are 100 and there are 5 equally spaced looks, test statistics are generated from the random samples at N1 = N2 = 20, N1 = N2 = 40, N1 = N2 = 60, N1 = N2 = 80, and N1 = N2 = 100 for both null and alternative samples.

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3. To generate the first significance boundary, the null distribution statistics of the first look (e.g., at N1 = N2 = 20) are ordered and the percent of alpha to be spent at the first look is determined (using either the alpha-spending function or the input value). The statistic for which the percent of statistics above (or below, as the case may be) that value is equal to the percent of alpha to be spent at the first look is the boundary statistic. It is seen here how important a large number of simulations is to the precision of the boundary estimates.

- 4. All null distribution samples that are outside the first significance boundary at the first look are removed from consideration for the second look. If binding futility boundaries are also being computed, all null distribution samples with statistics that are outside the first futility boundary are also removed from consideration for the second look. If non-binding futility boundaries are being computed, null distribution samples with statistics outside the first futility boundary are not removed.
- 5. To generate the second significance boundary, the remaining null distribution statistics of the second look (e.g., at N1 = N2 = 40) are ordered and the percent of alpha to be spent at the second look is determined (again, using either the alpha-spending function or the input value). The percent of alpha to be spent at the second look is multiplied by the total number of simulations to determine the number of the statistic that is to be the second boundary statistic. The statistic for which that number of statistics is above it (or below, as the case may be) is the second boundary statistic. For example, suppose there are initially 1000 simulated samples, with 10 removed at the first look (from, say, alpha spent at Look 1 equal to 0.01), leaving 990 samples considered for the second look. Suppose further that the alpha to be spent at the second look is 0.02. This is multiplied by 1000 to give 20. The 990 still-considered statistics are ordered and the 970th (20 in from 990) statistic is the second boundary.
- 6. All null distribution samples that are outside the second significance boundary and the second futility boundary, if binding, at the second look are removed from consideration for the third look (e.g., leaving 970 statistics computed at N1 = N2 = 60 to be considered at the third look). Steps 4 and 5 are repeated until the final look is reached.

Futility boundaries are computed in a similar manner using the desired beta-spending function or custom beta-spending values and the alternative hypothesis simulated statistics at each look. For both binding and non-binding futility boundaries, samples for which alternative hypothesis statistics are outside either the significance or futility boundaries of the previous look are excluded from current and future looks.

Because the final futility and significance boundaries are required to be the same, futility boundaries are computed beginning at a small value of beta (e.g., 0.0001) and incrementing beta by that amount until the futility and significance boundaries meet.

When boundaries are entered directly, this procedure uses the null hypothesis and alternative hypothesis simulations to determine the number of test statistics that are outside the boundaries at each look. The cumulative proportion of alternative hypothesis statistics that are outside the significance boundaries is the overall power of the study.

Small Sample Considerations

When the sample size is small, say 200 or fewer per group, the discrete nature of the number of possible differences in proportions in the sampling distribution comes into play. This has led to a large number of proposed tests for comparing two proportions (or testing the 2 by 2 table of counts). For example, Upton (1982) considers twenty-two alternative tests for comparing two proportions. Sweeping statements about the power of one test over another are impossible to make, because the size of the Type I error depends upon the proportions used. At some proportions, some tests are overly conservative while others are not, while at other proportions the reverse may be true.

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This simulation procedure, however, is based primarily on the ordering of the sample statistics in the simulation. The boundaries are determined by the spending function alphas. Thus, if a test used happens to be conservative in the single-look traditional sense, the boundaries chosen in the simulation results of this procedure will generally remove the conservative nature of the test. This makes comparisons to the one-look case surprising in many cases.

Definitions

Suppose you have two populations from which dichotomous (binary) responses will be recorded. The probability (or risk) of obtaining the event of interest in population 1 (the treatment group) is p_1 and in population 2 (the control group) is p_2 . The corresponding failure proportions are given by $q_1 = 1 - p_1$ and $q_2 = 1 - p_2$.

The assumption is made that the responses from each group follow a binomial distribution. This means that the event probability, p_i , is the same for all subjects within the group and that the response from one subject is independent of that of any other subject.

Random samples of m and n individuals are obtained from these two populations. The data from these samples can be displayed in a 2-by-2 contingency table as follows

Group	Success	Failure	Total
Treatment	а	С	m
Control	b	d	n
Total	S	f	Ν

The following alternative notation is also used.

Group	Success	Failure	Total
Treatment	x_{11}	x_{12}	n_1
Control	x_{21}	x_{22}	n_2
Total	m_1	m_2	N

The binomial proportions p_1 and p_2 are estimated from these data using the formulae

$$\hat{p}_1 = \frac{a}{m} = \frac{x_{11}}{n_1}$$
 and $\hat{p}_2 = \frac{b}{n} = \frac{x_{21}}{n_2}$

Comparing Two Proportions

Let $p_{1.0}$ represent the group 1 proportion tested by the null hypothesis, H_0 . The power of a test is computed at a specific value of the proportion which we will call $p_{1.1}$. Let δ represent the smallest difference (margin of non-inferiority) between the two proportions that still results in the conclusion that the new treatment is not inferior to the current treatment. For a non-inferiority test, $\delta < 0$. The set of statistical hypotheses that are tested is

$$H_0: p_{1,0} - p_2 \le \delta$$
 versus $H_1: p_{1,0} - p_2 > \delta$

which can be rearranged to give

$$H_0: p_{1,0} \le p_2 + \delta$$
 versus $H_1: p_{1,0} > p_2 + \delta$

There are multiple methods of specifying the margin of non-inferiority. The most direct is to simply give values for p_2 and $p_{1.0}$. However, it is often more meaningful to give p_2 and then specify $p_{1.0}$ implicitly by specifying the difference, ratio, or odds ratio. Mathematically, assuming higher proportions are better, the definitions of these parameterizations are

<u>Parameter</u>	<u>Computation</u>	<u>Hypotheses</u>
Difference	$\delta = p_{1.0} - p_2$	$H_0: p_{1.0} - p_2 \le \delta_0 \text{ vs. } H_1: p_{1.0} - p_2 > \delta_0, \delta_0 < 0$
Ratio	$\phi = p_{1.0} / p_2$	$H_0: p_1 / p_2 \le \phi_0 \text{ vs. } H_1: p_1 / p_2 > \phi_0, \ \phi_0 < 1$
Odds Ratio	$\psi = 0dds_{1.0} / 0dds_2$	$H_0: o_{1.0} / o_2 \le \psi_0 \text{ vs. } H_1: o_{1.0} / o_2 > \psi_0, \psi_0 < 1$

Odds Ratio

The odds ratio, $\psi = (p_{1.0} / (1 - p_{1.0})) / (p_2 / (1 - p_2))$, gives the relative change in the odds of the response. Testing non-inferiority uses the formulation

$$H_0: \psi \leq \psi_0$$
 versus $H_1: \psi > \psi_0$

when higher proportions are better.

Test Statistics

This section describes the test statistics that are available in this procedure.

Miettinen and Nurminen's Likelihood Score Test

Miettinen and Nurminen (1985) proposed a test statistic for testing whether the odds ratio is equal to a specified value, ψ_0 . Because the approach they used with the difference and ratio does not easily extend to the odds ratio, they used a score statistic approach for the odds ratio. The regular MLE's are \hat{p}_1 and \hat{p}_2 . The constrained MLE's are \tilde{p}_1 and \tilde{p}_2 . These estimates are constrained so that $\tilde{\psi}=\psi_0$. A correction factor of N/(N-1) is applied to make the variance estimate less biased. The significance level of the test statistic is based on the asymptotic normality of the score statistic.

The formula for computing the test statistic is

$$z_{MNO} = \frac{\frac{(\hat{p}_1 - \tilde{p}_1)}{\tilde{p}_1 \tilde{q}_1} - \frac{(\hat{p}_2 - \tilde{p}_2)}{\tilde{p}_2 \tilde{q}_2}}{\sqrt{\left(\frac{1}{n_1 \tilde{p}_1 \tilde{q}_1} + \frac{1}{n_2 \tilde{p}_2 \tilde{q}_2}\right) \left(\frac{N}{N-1}\right)}}$$

where

$$\tilde{p}_1 = \frac{\tilde{p}_2 \psi_0}{1 + \tilde{p}_2 (\psi_0 - 1)}$$

$$\tilde{p}_2 = \frac{-B + \sqrt{B^2 - 4AC}}{2A}$$

$$A = n_2(\psi_0 - 1)$$

$$B = n_1 \psi_0 + n_2 - m_1 (\psi_0 - 1)$$

$$C = -m_1$$

Farrington and Manning's Likelihood Score Test

Farrington and Manning (1990) indicate that the Miettinen and Nurminen statistic may be modified by removing the factor N/(N-1).

The formula for computing this test statistic is

$$z_{FMO} = \frac{\frac{(\hat{p}_1 - \tilde{p}_1)}{\tilde{p}_1 \tilde{q}_1} - \frac{(\hat{p}_2 - \tilde{p}_2)}{\tilde{p}_2 \tilde{q}_2}}{\sqrt{\left(\frac{1}{n_1 \tilde{p}_1 \tilde{q}_1} + \frac{1}{n_2 \tilde{p}_2 \tilde{q}_2}\right)}}$$

where the estimates \tilde{p}_1 and \tilde{p}_2 are computed as in the corresponding test of Miettinen and Nurminen (1985) given above.

Spending Functions

Spending functions can be used in this procedure to specify the proportion of alpha or beta that is spent at each look without having to specify the proportion directly.

Spending functions have the characteristics that they are increasing and that

$$\alpha(0) = 0$$

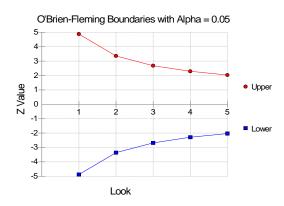
$$\alpha(1) = \alpha$$

The last characteristic guarantees a fixed α level when the trial is complete. This methodology is very flexible since neither the times nor the number of analyses must be specified in advance. Only the functional form of $\alpha(\tau)$ must be specified.

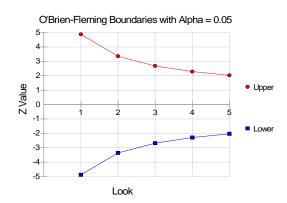
PASS provides several popular spending functions plus the ability to enter and analyze your own percents of alpha or beta spent. These are calculated as follows (beta may be substituted for alpha for beta-spending functions):

1. Hwang-Shih-DeCani (gamma family)

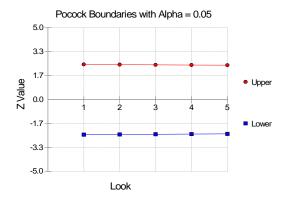
$$\alpha \left[\frac{1 - e^{-\gamma t}}{1 - e^{-\gamma}} \right], \gamma \neq 0; \quad \alpha t, \gamma = 0$$



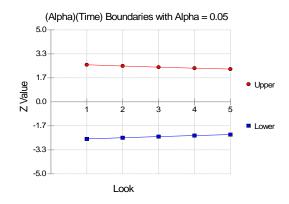
2. O'Brien-Fleming Analog $2-2\Phi\left(\frac{Z_{\alpha/2}}{\sqrt{t}}\right)$



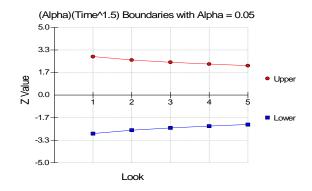
3. Pocock Analog $\alpha \cdot \ln(1 + (e-1)t)$



4. Alpha * time $\alpha \cdot t$

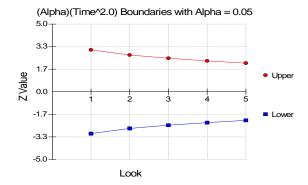


5. Alpha * time^1.5 $\alpha \cdot t^{3/2}$

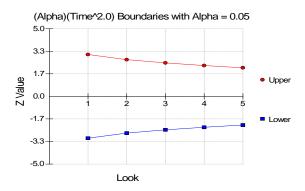


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6. Alpha * time^2 $\alpha \cdot t^2$



7. Alpha * time^C $\alpha \cdot t^C$



8. User Supplied Percents

A custom set of percents of alpha to be spent at each look may be input directly.

The O'Brien-Fleming Analog spends very little alpha or beta at the beginning and much more at the final looks. The Pocock Analog and (Alpha or Beta)(Time) spending functions spend alpha or beta more evenly across the looks. The Hwang-Shih-DeCani (C) (gamma family) spending functions and (Alpha or Beta)(Time^C) spending functions are flexible spending functions that can be used to spend more alpha or beta early or late or evenly, depending on the choice of C.

Example 1 – Power and Output

A clinical trial is to be conducted over a two-year period to compare the proportion response of a new treatment to that of the current treatment. The current response proportion is 0.58. The researchers would like to show that the odds ratio of new treatment to the old is at least 0.8. Although the researchers do not know the true proportion of patients that will survive with the new treatment, they would like to examine the power that is achieved if the proportion under the new treatment is also 0.58 (odds ratio of 1.0). The sample size at the final look is to be 1000 per group. Testing will be done at the 0.05 significance level. A total of five tests are going to be performed on the data as they are obtained. The O'Brien-Fleming (Analog) boundaries will be used.

Find the power and test boundaries assuming equal sample sizes per arm.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Power
Test Type	Likelihood Score (Farr. & Mann.)
Higher Proportions Are	Better
Simulations	100000
Random Seed	5143819 (for Reproducibility)
Alpha	0.05
Group Allocation	Equal (N1 = N2)
Sample Size Per Group	1000
OR0 (Non-Inferiority Odds Ratio)	0.8
OR1 (Actual Odds Ratio)	1.0
P2 (Group 2 Proportion)	0.58
Looks & Boundaries Tab	
Specification of Looks and Boundaries	Simple
Number of Equally Spaced Looks	5
Alpha Spending Function	O'Brien-Fleming Analog
Type of Futility Boundary	None

Output

Click the Calculate button to perform the calculations and generate the following output.

Scenario 1 Numeric Results for Group Sequential Testing Odds Ratio = NIOR.

Solve For: Power

Hypotheses: H0: Odds 1 / Odds 2 = NIOR; H1: Odds 1 / Odds 2 > NIOR

Test Statistic: Likelihood Score (Farrington and Manning) Test

Zero Adjustment Method: None

Alpha-Spending Function: O'Brien-Fleming Analog

Beta-Spending Function: None
Futility Boundary Type: None
Number of Looks: 5
Simulations: 100000

Random Seed: 5143819 (User-Entered)

Numeric Summary for Scenario 1

	Power				Alpha		
Value	95% LCL	95% UCL	Target	Actual	95% LCL	95% UCL	Beta
0.775	0.772	0.777	0.05	0.05	0.049	0.051	0.225

		A	verage S	erage Sample Size						
	_	Given	Given H0		en H1					
N1 N2 Grp	Grp1	Grp2	Grp1	Grp2	OR0	OR1	P1.0	P1.1	P2	
1000	1000	992	992	794	794	0.8	1	0.52	0.58	0.58
Power			propo		rejecting a fa					
Power 9	95% LCL ar	d UCL			er confiden		the power	estimate. T	he width o	f the interval is
Target A	Alpha		The us	er-specified	probability	of rejecting	a true null	hypothesis	. It is the to	tal alpha spent.
_	r Actual Alp	ha	The alp	oha İevel tha		ılly achieve	d by the ex	periment. It	is the total	proportion of
Alpha 9	5% LCL an	d UCL	The lov	wer and upp	er confiden	ce limits for	the actual			
Beta					accepting a ations that o					of alternative
N1 and	N2				of each grou					
Average	Sample S	ze Given H0	propo							are based on the tility boundaries
Average	e Sample S	ze Given H1	propo		rnative hyp					are based on the e or futility
OR0				n-inferiority thesis, H0.	odds ratio is	s the odds	ratio betwe	en groups a	assuming th	ne null
OR1			The od	lds ratio bet	ween group	s assuming	the alterna	ative hypoth	nesis, H1.	
			The pro	oportion use	ed in the sim	ulations for	Group 1 u	nder H0.		
P1.0										
P1.0 P1.1 P2					ed in the simed in the sime	nulations for	r Group 1 u	nder H1.		

Summary Statements

A group sequential trial with sample sizes of 1000 and 1000 at the final look achieve 77% power to detect an odds ratio of 1 of the treatment group proportion to a control group proportion of 0.58 with a non-inferiority odds ratio of 0.8 with an overall Type I error rate (α) of 0.05 using a one-sided Likelihood Score (Farrington and Manning) Test.

Accumulated Information Details for Scenario 1

	Accumulated Information	Accumulated Sample Size			
Look	Percent	Group 1	Group 2	Total	
1	20	200	200	400	
2	40	400	400	800	
3	60	600	600	1200	
4	80	800	800	1600	
5	100	1000	1000	2000	

Look

Accumulated Information Percent Accumulated Sample Size Group 1 Accumulated Sample Size Group 2 Accumulated Sample Size Total The number of the look.

The percent of the sample size accumulated up to the corresponding look.

The total number of individuals in group 1 at the corresponding look.

The total number of individuals in group 2 at the corresponding look.

The total number of individuals in the study (group 1 + group 2) at the corresponding look.

Boundaries for Scenario 1

Look 1 2	Significanc	e Boundary
Look	Z-Value Scale	P-Value Scale
1	4.32394	0.00001
2	2.85340	0.00216
3	2.32637	0.01000
4	1.97242	0.02428
5	1.75382	0.03973

Look

Significance Boundary Z-Value Scale

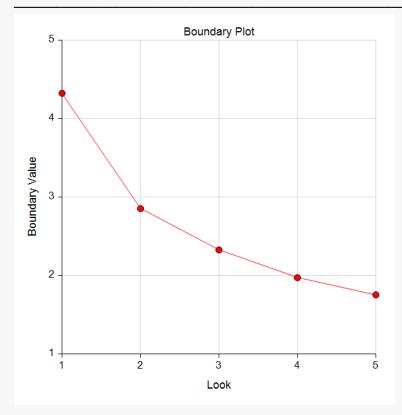
The number of the look.

The value such that statistics outside this boundary at the corresponding look indicate termination of the study and rejection of the null hypothesis. They are sometimes called efficacy boundaries.

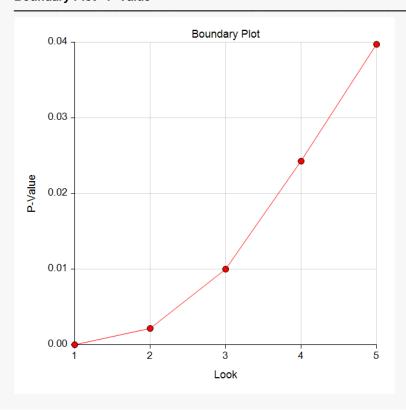
Significance Boundary P-Value Scale

The value such that P-Values outside this boundary at the corresponding look indicate termination of the study and rejection of the null hypothesis. This P-Value corresponds to the Z-Value Boundary and is sometimes called the nominal alpha.

Boundary Plot



Boundary Plot - P-Value



Significance Boundaries with 95% Simulation Confidence Intervals for Scenario 1

		-Value Bound	ary	P	P-Value Boundary				
Look	Value	95% LCL	95% UCL	Value	95% LCL	95% UCL			
1	4.32394			0.00001					
2	2.85340	2.85145	2.92356	0.00216	0.00173	0.00218			
3	2.32637	2.27759	2.33140	0.01000	0.00987	0.01138			
4	1.97242	1.96750	2.00920	0.02428	0.02226	0.02456			
5	1.75382	1.72549	1.76202	0.03973	0.03903	0.04222			

of the study and rejection of the null hypothesis. They are sometimes called efficacy boundaries.

P-Value Boundary Value

The value such that P-Values outside this boundary at the corresponding look indicate termination of the study and rejection of the null hypothesis. This P-Value corresponds to the Z-Value Boundary and is sometimes called the nominal alpha.

95% LCL and UCL

The lower and upper confidence limits for the boundary at the given look. The width of the interval is based on the number of simulations.

Alpha-Spending and Null Hypothesis Simulation Details for Scenario 1

			Tai	rget	Ac	tual	D	Cum
•		·		Cum. Spending		Cum.	Proportion H1 Sims Outside	Cum. H1 Sims Outside
		Spending Spending Function Function Alpha Alpha		Alpha Alpha Spent Spent		Signif. Boundary	Signif. Boundary	
1	4.32394	0.00001	0.000	0.000	0.000	0.000	0.001	0.001
2	2.85340	0.00216	0.002	0.002	0.002	0.002	0.091	0.092
3	2.32637	0.01000	0.009	0.011	0.009	0.011	0.246	0.338
4	1.97242	0.02428	0.017	0.028	0.017	0.028	0.262	0.600
5	1.75382	0.03973	0.022	0.050	0.022	0.050	0.174	0.775

Look

Significance Boundary Z-Value Scale

Significance Boundary P-Value Scale

Spending Function Alpha

Cumulative Spending Function Alpha

Alpha Spent

Cumulative Alpha Spent

Proportion H1 Sims Outside Significance Boundary

Cumulative H1 Sims Outside Significance Boundary

The number of the look.

The value such that statistics outside this boundary at the corresponding look indicate termination of the study and rejection of the null hypothesis. They are sometimes called efficacy boundaries.

The value such that P-Values outside this boundary at the corresponding look indicate termination of the study and rejection of the null hypothesis. This P-Value corresponds to the Significance Z-Value Boundary and is sometimes called the nominal alpha.

The intended portion of alpha allocated to the particular look based on the alpha-spending function.

The intended accumulated alpha allocated to the particular look. It is the sum of the Spending Function Alpha up to the corresponding look.

The proportion of the null hypothesis simulations resulting in statistics outside the Significance Boundary at this look.

The proportion of the null hypothesis simulations resulting in Significance Boundary termination up to and including this look. It is the sum of the Alpha Spent up to the corresponding look.

The proportion of the alternative hypothesis simulations resulting in statistics outside the Significance Boundary at this look. It may be thought of as the incremental power.

The proportion of the alternative hypothesis simulations resulting in Significance Boundary termination up to and including this look. It is the sum of the Proportion H1 Sims Outside Significance Boundary up to the corresponding look.

Group-Sequential Non-Inferiority Tests for the Odds Ratio of Two Proportions (Simulation)

Numeric Results for Group Sequential Testing Proportion Odds Ratio = NIOR.

Solve For: Power

Hypotheses: H0: Odds 1 / Odds 2 = NIOR; H1: Odds 1 / Odds 2 > NIOR

Test Statistic: Likelihood Score (Farrington and Manning) Test

Zero Adjustment Method: None

Alpha-Spending Function: O'Brien-Fleming Analog

Beta-Spending Function: None
Futility Boundary Type: None
Number of Looks: 5
Simulations: 100000

Random Seed: 5143819 (User-Entered)

Numeric Summary of Scenarios

Scenario	Power	N1	N2	Alpha	OR0	OR1	P1.0	P1.1	P2
1	0.775	1000	1000	0.05	0.8	1	0.52	0.58	0.58
Power	•		•	ull hypothesi			It is the tota	l proportion	n of alternative
Alpha	•		,	chieved by the gnificance be	•	nt. It is the	total propor	rtion of the	null hypothesis
N1 and N2	The sample	sizes of eac	ch group if the	he study read	ches the fin	al look.			
OR0	The non-infe	riority odds	ratio is the	odds ratio be	etween grou	ıps assumi	ng the null l	hypothesis	, H0.
OR1	The odds rat	io between	groups ass	uming the alt	ernative hy	pothesis, F	1 1.		
P1.0	The proportion	on used in t	he simulation	ons for Group	1 under H	0.			
P1.1	The proportion	on used in t	he simulation	ons for Group	1 under H	1.			
P2	The proporti	on used in t	he simulation	ons for Group	2 under H	0 and H1.			

Power and Alpha Summary

		Power						
Scenario	Value	95% LCL	95% UCL	Target	Actual	95% LCL	95% UCL	Beta
1	0.775	0.772	0.772 0.777 0.05 0.05 0.049 0.051					
Power			oility of rejecting a				It is the total propoundaries.	portion of
Power 95% L	_CL and UCL		and upper confid mber of simulation		r the power e	stimate. The wid	Ith of the interval	is based
Target Alpha		The user-sp	pecified probabili	ty of rejecting	a true null h	ypothesis. It is th	ne total alpha spe	ent.
Alpha or Actu	ual Alpha	•	evel that was ac	,	, ,		total proportion	of the null
Alpha 95% L	CL and UCL		and upper confid the number of s		r the actual a	lpha estimate. T	he width of the ir	nterval is
Beta			oility of accepting is simulations the				ortion of alternation	ve

Group-Sequential Non-Inferiority Tests for the Odds Ratio of Two Proportions (Simulation)

Sample Size Summary

						Average S	Sample Siz	:e	
					Give	en H0	Give	en H1	
Scenario	Power	Alpha	N1	N2	Grp1	Grp2	Grp1	Grp2	
1	0.775	0.05	1000	1000	992	992	794	794	
Power		pre		, .	•	hypothesis simulations			
Alpha						eved by the are outside			ital proportion of aries.
N1 and N2		The	sample siz	es of each	group if the	study reach	nes the final	look.	
Average San	nple Size Give	pre							e are based on the futility boundaries
Average San	nple Size Give	pre	oportion of	•	hypothesis		•		e are based on the ince or futility

Run Time: 4.35 minutes.

References

Jennison, C., Turnbull, B.W. 2000. Group Sequential Methods with Applications to Clinical Trials. Chapman & Hall. Boca Raton. FL.

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The values obtained from any given run of this example will vary slightly due to the variation in simulations.

Example 2 - Power with Futility Boundaries

Continuing with Example 1, suppose that the researchers would also like to terminate the study early if there is indication that the treatment is not non-inferior.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Power
Test Type	Likelihood Score (Farr. & Mann.)
Higher Proportions Are	Better
Simulations	100000
Random Seed	<i>Blank</i> or Random
Alpha	0.05
Group Allocation	Equal (N1 = N2)
Sample Size Per Group	1000
OR0 (Non-Inferiority Odds Ratio)	0.8
OR1 (Actual Odds Ratio)	1.0
P2 (Group 2 Proportion)	0.58
Looks & Boundaries Tab	
Specification of Looks and Boundaries	Simple
Number of Equally Spaced Looks	5
Alpha Spending Function	O'Brien-Fleming Analog
Type of Futility Boundary	Non-Binding
Number of Skipped Futility Looks	0
Beta Spending Function	O'Brien-Fleming Analog

Output

Click the Calculate button to perform the calculations and generate the following output.

Scenario 1 Numeric Results for Group Sequential Testing Odds Ratio = NIOR.

Solve For: Power

Hypotheses: H0: Odds 1 / Odds 2 = NIOR; H1: Odds 1 / Odds 2 > NIOR

Test Statistic: Likelihood Score (Farrington and Manning) Test

Zero Adjustment Method: None

Alpha-Spending Function: O'Brien-Fleming Analog Beta-Spending Function: O'Brien-Fleming Analog

Futility Boundary Type: Non-Binding

Number of Looks:

100000

Simulations:

Random Seed: 5772343 (Computer-Generated)

Numeric Summary for Scenario 1

	Power		Alpha				
Value	95% LCL	95% UCL	Target	Actual	95% LCL	95% UCL	Beta
0.708	0.705	0.71	0.05	0.041	0.04	0.042	0.292

		,	Average S	ample Siz	е					
		Give	en H0	Give	en H1					
N1	N2	Grp1	Grp2	Grp1	Grp2	OR0	OR1	P1.0	P1.1	P2
1000	1000	496	496	690	690	0.8	1	0.52	0.58	0.58

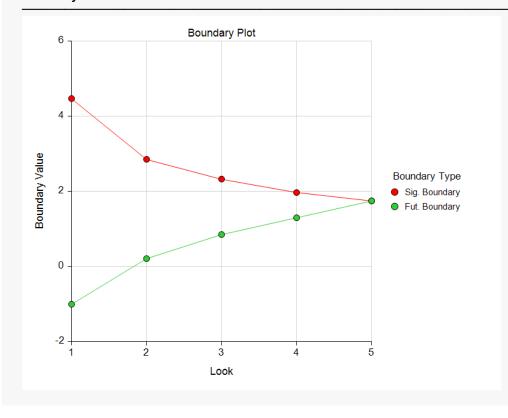
Accumulated Information Details for Scenario 1

	Accumulated Information	Accumulated Sample Size				
Look	Percent	Group 1	Group 2	Total		
1	20	200	200	400		
2	40	400	400	800		
3	60	600	600	1200		
4	80	800	800	1600		
5	100	1000	1000	2000		

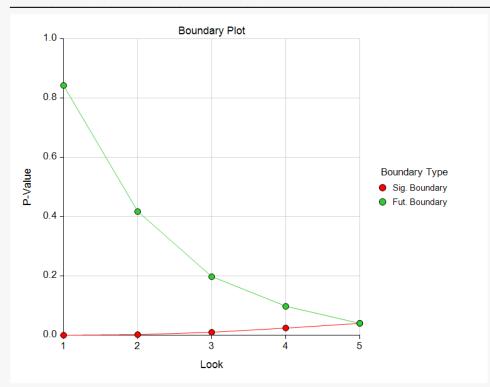
Boundaries for Scenario 1

	Significand	e Boundary	Futility B	Boundary
Look	Z-Value Scale	P-Value Scale	Z-Value Scale	P-Value Scale
1	4.47018	0.00000	-1.00492	0.84253
2	2.85302	0.00217	0.20907	0.41720
3	2.32700	0.00998	0.85077	0.19745
4	1.96817	0.02452	1.29679	0.09735
5	1.74684	0.04033	1.74684	0.04033

Boundary Plot



Boundary Plot - P-Value



Significance Boundaries with 95% Simulation Confidence Intervals for Scenario 1

	Z	Z-Value Boundary P-Value Boundary				
Look	Value	95% LCL	95% UCL	Value	95% LCL	95% UCL
1	4.47018			0.00000		
2	2.85302	2.85094	2.92318	0.00217	0.00173	0.00218
3	2.32700	2.27692	2.33187	0.00998	0.00985	0.01140
4	1.96817	1.96156	1.97296	0.02452	0.02425	0.02491
5	1.74684	1.72173	1.75774	0.04033	0.03940	0.04256

Futility Boundaries with 95% Simulation Confidence Intervals for Scenario 1

	Z-Value Boundary			P	P-Value Boundary				
Look	Value	95% LCL	95% UCL	Value	95% LCL	95% UCL			
1	-1.00492	-1.01677	-0.98021	0.84253					
2	0.20907	0.20085	0.22049	0.41720	0.41274	0.42041			
3	0.85077	0.84371	0.85570	0.19745	0.19608	0.19942			
4	1.29679	1.29191	1.30241	0.09735	0.09639	0.09819			
5	1.74684	1.74003	1.76065	0.04033	0.03915	0.04093			

Alpha-Spending and Null Hypothesis Simulation Details for Scenario 1

			Tar	get	Ac	tual		
	Signif. B	Boundary	Spending	Cum. Spending		Cum.	Proportion H0 Sims Outside	Cum. H0 Sims Outside
Look	Z-Value Scale	P-Value Scale	Function Alpha	Function Alpha	Alpha Spent	Alpha Spent	Futility Boundary	Futility Boundary
1	4.47018	0.00000	0.000	0.000	0.000	0.000	0.157	0.157
2	2.85302	0.00217	0.002	0.002	0.002	0.002	0.428	0.585
3	2.32700	0.00998	0.009	0.011	0.009	0.011	0.234	0.818
4	1.96817	0.02452	0.017	0.028	0.016	0.027	0.098	0.917
5	1.74684	0.04033	0.022	0.050	0.013	0.041	0.042	0.959

Beta-Spending and Alternative Hypothesis Simulation Details for Scenario 1

			Taı	get	Ac	tual	D	•
	Futility B	Boundary	Spending	Cum. Spending		Cum.	Proportion H1 Sims Outside	Cum. H1 Sims Outside
Look	Z-Value Scale	P-Value Scale	Function Beta	Function Beta	Beta Spent	Beta Spent	Signif. Boundary	Signif. Boundary
1	-1.00492	0.84253	0.019	0.019	0.019	0.019	0.000	0.000
2	0.20907	0.41720	0.078	0.096	0.078	0.096	0.093	0.093
3	0.85077	0.19745	0.078	0.175	0.078	0.174	0.245	0.338
4	1.29679	0.09735	0.065	0.240	0.065	0.239	0.251	0.590
5	1.74684	0.04033	0.053	0.293	0.053	0.292	0.118	0.708

The values obtained from any given run of this example will vary slightly due to the variation in simulations.

Example 3 - Enter Boundaries

With a set-up similar to Example 2, suppose we wish to investigate the properties of a set of significance (3, 3, 3, 2, 1) and futility (-2, -1, 0, 0, 1) boundaries.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 3** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Alpha and Power (Enter Boundaries)
Test Type	Likelihood Score (Farr. & Mann.)
Higher Proportions Are	Better
Simulations	100000
Random Seed	
Group Allocation	Equal (N1 = N2)
Sample Size Per Group	1000
OR0 (Non-Inferiority Odds Ratio)	0.8
OR1 (Actual Odds Ratio)	1.0
P2 (Group 2 Proportion)	0.58
Looks & Boundaries Tab	
Number of Looks	5
Equally Spaced	Checked
Types of Boundaries	Significance and Futility Boundaries
Significance Boundary	3 3 3 2 1 (for looks 1 through 5)
Futility Boundary	2 -1 0 0 1 (for looks 1 through 5)

Output

Click the Calculate button to perform the calculations and generate the following output.

Scenario 1 Numeric Results for Group Sequential Testing Odds Ratio = NIOR.

Solve For: Alpha and Power (Enter Boundaries)

Hypotheses: H0: Odds 1 / Odds 2 = NIOR; H1: Odds 1 / Odds 2 > NIOR

Test Statistic: Likelihood Score (Farrington and Manning) Test

Zero Adjustment Method: None

Type of Boundaries: Significance and Futility Boundaries

Number of Looks: 5

Simulations: 100000

Random Seed: 5804029 (Computer-Generated)

Numeric Summary for Scenario 1

	Power			Alpha			
Value	95% LCL	95% UCL	Value	95% LCL	95% LCL	Beta	
0.918	0.917	0.92	0.151	0.149	0.153	0.082	

		,	Average S							
		Give	n H0	Give	n H1					
N1	N2	Grp1	Grp2	Grp1	Grp2	OR0	OR1	P1.0	P1.1	P2
1000	1000	739	739	813	813	0.8	1	0.52	0.58	0.58

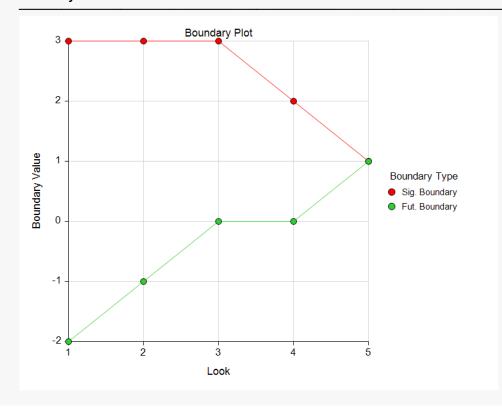
Accumulated Information Details for Scenario 1

Look	Accumulated Information	Accumulated Sample Size				
	Percent	Group 1	Group 2	Total		
1	20	200	200	400		
2	40	400	400	800		
3	60	600	600	1200		
4	80	800	800	1600		
5	100	1000	1000	2000		

Boundaries for Scenario 1

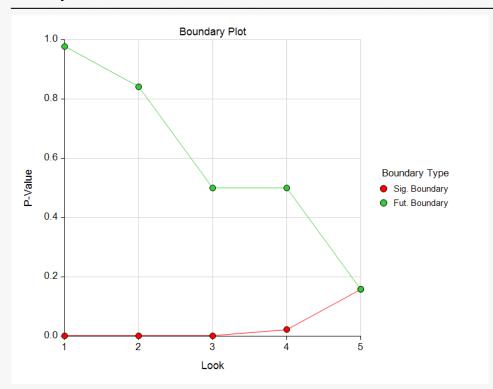
	Significand	e Boundary	Futility Boundary		
Look	Z-Value Scale	P-Value Scale	Z-Value Scale	P-Value Scale	
1	3	0.00135	-2	0.97725	
2	3	0.00135	-1	0.84134	
3	3	0.00135	0	0.50000	
4	2	0.02275	0	0.50000	
5	1	0.15866	1	0.15866	

Boundary Plot



Group-Sequential Non-Inferiority Tests for the Odds Ratio of Two Proportions (Simulation)





Alpha-Spending and Null Hypothesis Simulation Details for Scenario 1

	Signif. B	Boundary			Proportion H0 Sims	Cum. H0 Sims
Look	Z-Value Scale	P-Value Scale	Alpha Spent	Cum. Alpha Spent	Outside Futility Boundary	Outside Futility Boundary
1	3	0.00135	0.002	0.002	0.023	0.023
2	3	0.00135	0.001	0.002	0.139	0.162
3	3	0.00135	0.001	0.003	0.341	0.503
4	2	0.02275	0.021	0.024	0.081	0.585
5	1	0.15866	0.127	0.151	0.264	0.849

Beta-Spending and Alternative Hypothesis Simulation Details for Scenario 1

	Futility E	Boundary			Proportion H1 Sims	Cum. H1 Sims
Look	Z-Value Scale	P-Value Scale	Beta Spent	Cum. Beta Spent	Outside Signif. Boundary	Outside Signif. Boundary
1	-2	0.97725	0.001	0.001	0.031	0.031
2	-1	0.84134	0.005	0.006	0.054	0.084
3	0	0.50000	0.024	0.030	0.081	0.165
4	0	0.50000	0.004	0.033	0.418	0.583
5	1	0.15866	0.048	0.082	0.335	0.918

The values obtained from any given run of this example will vary slightly due to the variation in simulations.

Example 4 - Validation using Simulation

With a set-up similar to Example 2, we examine the power and alpha generated by the set of significance (3.83553, 2.92248, 2.32439, 1.96539, 1.74400) and futility (-0.99709, 0.20512, 0.85245, 1.30452, 1.74400) boundaries.

For reproducibility, we'll use a random seed of 8485162.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 4** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Alpha and Power (Enter Boundaries)
Test Type	Likelihood Score (Farr. & Mann.)
Higher Proportions Are	Better
Simulations	100000
Random Seed	8485162 (for Reproducibility)
Group Allocation	Equal (N1 = N2)
Sample Size Per Group	1000
OR0 (Non-Inferiority Odds Ratio)	0.8
OR1 (Actual Odds Ratio)	1.0
P2 (Group 2 Proportion)	0.58
Looks & Boundaries Tab	
Number of Looks	5
Equally Spaced	Checked
Types of Boundaries	Significance and Futility Boundaries
Significance Boundary	3.83553, 2.92248, 2.32439, 1.96539, 1.74400
Futility Boundary	0.99709, 0.20512, 0.85245, 1.30452, 1.74400

Output

Click the Calculate button to perform the calculations and generate the following output.

Scenario 1 Numeric Results for Group Sequential Testing Odds Ratio = NIOR.

Solve For: Alpha and Power (Enter Boundaries)

Hypotheses: H0: Odds 1 / Odds 2 = NIOR; H1: Odds 1 / Odds 2 > NIOR

Test Statistic: Likelihood Score (Farrington and Manning) Test

Zero Adjustment Method: None

Type of Boundaries: Significance and Futility Boundaries

Number of Looks:

Simulations: 100000

Random Seed: 8485162 (User-Entered)

Numeric Summary for Scenario 1

	Power			Alpha					
Value	95% LCL	95% UCL	Value	95% LCL	95% LCL	Beta			
0.705	0.702	0.708	0.041	0.039	0.042	0.295			

N1			Average S	ample Siz	e					
		Given H0		Given H1						
	N2	Grp1	Grp2	Grp1	Grp2	OR0	OR1	P1.0	P1.1	P2
1000	1000	496	496	691	691	0.8	1	0.52	0.58	0.58

The values obtained from any given run of this example will vary slightly due to the variation in simulations. The power and alpha generated with these boundaries are very close to the values of Example 2.