

## Chapter 313

# Non-Unity Null Tests for Two Total Variances in a 2×2 Cross-Over Design

## Introduction

This procedure calculates power and sample size of tests of total variabilities (between + within) from a 2×2 cross-over design for the case when the ratio assumed by the null hypothesis is not necessarily equal to one. This routine deals with the case in which the statistical hypotheses are expressed in terms of the ratio of the total variances.

This design is used to compare two treatments which are administered to subjects in different orders. The design has two treatment sequences. The two sequences are

sequence 1: C T

sequence 2: T C

It is assumed that either there is no carry-over from one measurement to the next, or there is an ample washout period between measurements.

## Technical Details

This procedure uses the formulation given in Chow, Shao, Wang, and Lokhnygina (2018), pages 224 - 227.

Suppose  $x_{ijk}$  is the response in the  $i$ th sequence ( $i = 1, 2$ ),  $j$ th subject ( $j = 1, \dots, N_i$ ), and  $k$ th treatment ( $k = T, C$ ). The mixed effect model analyzed in this procedure is

$$x_{ijk} = \mu_k + \gamma_{ik} + S_{ijk} + e_{ijk}$$

where  $\mu_k$  is the  $k$ th treatment effect,  $\gamma_{ik}$  is the interaction between sequence  $i$  and treatment  $k$ ,  $S_{ijT}$  and  $S_{ijC}$  are random effects of the  $j$ th subject, and  $e_{ijk}$  is the within-subject error term which is normally distributed with mean 0 and variance  $V_k = \sigma_{Wk}^2$ .

Let  $N_s = N_1 + N_2 - 2$ . The total variances ( $\sigma_{Tk}^2 = \sigma_{Bk}^2 + \sigma_{Wk}^2$ ,  $k = T, C$ ) are estimated by

$$\hat{\sigma}_{Tk}^2 = s_{Tk}^2 = \frac{1}{N_s} \sum_{i=1}^2 \sum_{j=1}^{N_i} (\bar{x}_{ijk} - \bar{x}_{i.k})^2$$

where

$$\bar{x}_{i.k} = \frac{1}{N_i} \sum_{j=1}^{N_i} \bar{x}_{ijk}$$

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The sample between-subject covariance is calculated using

$$s_{BTC}^2 = \frac{1}{N_s} \sum_{i=1}^2 \sum_{j=1}^{N_i} (\bar{x}_{ijT} - \bar{x}_{i.T})(\bar{x}_{ijC} - \bar{x}_{i.C})$$

Using this value, the sample between-subject correlation is easily calculated.

## Testing Variance Inequality with a Non-Unity Null

The following three sets of statistical hypotheses are used to test for total variance inequality with a non-unity null

$$H_0: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} \geq R0 \quad \text{versus} \quad H_1: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} < R0,$$

$$H_0: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} \leq R0 \quad \text{versus} \quad H_1: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} > R0,$$

$$H_0: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} = R0 \quad \text{versus} \quad H_1: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} \neq R0,$$

where  $R0$  is the variance ratio assumed by the null hypothesis.

Let  $\eta = \sigma_{TT}^2 - R0\sigma_{TC}^2$  be the parameter of interest. The test statistic is  $\hat{\eta} = \hat{\sigma}_{TT}^2 - R0\hat{\sigma}_{TC}^2$ .

### Two-Sided Test

For the two-sided test, compute two limits,  $\hat{\eta}_L$  and  $\hat{\eta}_U$ , using

$$\hat{\eta}_L = \hat{\eta} - \sqrt{\Delta_L}$$

$$\hat{\eta}_U = \hat{\eta} + \sqrt{\Delta_U}$$

Reject the null hypothesis if  $\hat{\eta}_L > 0$  is or  $\hat{\eta}_U < 0$ .

The  $\Delta$ s are given by

$$\Delta_L = h\left(1 - \frac{\alpha}{2}, N_s\right) \lambda_1^2 + h\left(\frac{\alpha}{2}, N_s\right) \lambda_2^2$$

$$\Delta_U = h\left(\frac{\alpha}{2}, N_s\right) \lambda_1^2 + h\left(1 - \frac{\alpha}{2}, N_s\right) \lambda_2^2$$

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where

$$h(A, B) = \left(1 - \frac{B}{\chi_{A,B}^2}\right)^2$$

$$\lambda_i^2 = \left( \frac{s_{TT}^2 - (R_0)s_{TC}^2 \pm \sqrt{(s_{TT}^2 + (R_0)s_{TC}^2)^2 - 4(R_0)s_{BTC}^4}}{2} \right) \text{ for } i = 1, 2$$

and  $\chi_{A,B}^2$  is the upper quantile of the chi-square distribution with  $B$  degrees of freedom.

### One-Sided Test

For the lower, one-sided test, compute the limit,  $\hat{\eta}_U$ , using

$$\hat{\eta}_U = \hat{\eta} + \sqrt{\Delta_U}$$

Reject the null hypothesis if  $\hat{\eta}_U < 0$ .

The  $\Delta_U$  is given by

$$\Delta_U = h(\alpha, N_s)\lambda_1^2 + h(1 - \alpha, N_s)\lambda_2^2$$

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## Power

### Two-Sided Test

The power of the two-sided test is given by

$$\text{Power} = 1 - \Phi\left(z_{1-\frac{\alpha}{2}} - \frac{(R_1 - R_0)\sigma_{TC}^2}{\sqrt{\sigma^{*2}/N_s}}\right) + \Phi\left(z_{\alpha/2} - \frac{(R_1 - R_0)\sigma_{TC}^2}{\sqrt{\sigma^{*2}/N_s}}\right)$$

where

$$R_1 = \frac{\sigma_{TT}^2}{\sigma_{TC}^2}$$

$$\sigma_{TT}^2 = R_1\sigma_{TC}^2$$

$$\sigma^{*2} = 2[\sigma_{TT}^4 + R_0^2\sigma_{TC}^4 - 2R_0\sigma_{BT}^2\sigma_{BC}^2\rho^2]$$

where  $R_1$  is the value of the variance ratio stated by the alternative hypothesis and  $\Phi(x)$  is the standard normal CDF.

A simple binary search algorithm can be applied to the power function to obtain an estimate of the necessary sample size.

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**One-Sided Test**

The power of the lower, one-sided test,  $H_0: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} \geq R_0$  versus  $H_1: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} < R_0$ , is given by

$$\text{Power} = \Phi \left( z_\alpha - \frac{(R_1 - R_0)\sigma_{TC}^2}{\sqrt{\sigma^{*2}/N_s}} \right)$$

The power of the upper, one-sided test,  $H_0: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} \leq R_0$  versus  $H_1: \frac{\sigma_{TT}^2}{\sigma_{TC}^2} > R_0$ , is given by

$$\text{Power} = 1 - \Phi \left( z_{1-\alpha} - \frac{(R_1 - R_0)\sigma_{TC}^2}{\sqrt{\sigma^{*2}/N_s}} \right)$$

## Example 1 – Finding Sample Size

A company has developed a generic drug for treating rheumatism and wants to compare it to the standard drug in terms of the total variability. A 2 x 2 cross-over design will be used to test the inequality using a two-sided test.

Company researchers set the variance ratio under the null hypothesis to 0.8, the significance level to 0.05, the power to 0.90, and the actual variance ratio values between 0.5 and 1.3. They also set  $\sigma^2_{TC} = 0.8$ ,  $\sigma^2_{WT} = 0.2$ ,  $\sigma^2_{WC} = 0.3$ , and  $\rho = 0.7$ . They want to investigate the range of required sample size values assuming that the two sequence sample sizes are equal.

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For .....	<b>Sample Size</b>
Alternative Hypothesis .....	<b>Two-Sided (<math>H_1: \sigma^2_{TT}/\sigma^2_{TC} \neq R_0</math>)</b>
Power.....	<b>0.90</b>
Alpha.....	<b>0.05</b>
Sequence Allocation .....	<b>Equal (<math>N_1 = N_2</math>)</b>
$R_0$ ( $H_0$ Variance Ratio).....	<b>0.8</b>
$R_1$ (Actual Variance Ratio) .....	<b>0.5 0.7 0.9 1 1.1 1.3</b>
$\sigma^2_{TC}$ (Control Variance).....	<b>0.8</b>
$\sigma^2_{WT}$ (Treatment Variance) .....	<b>0.2</b>
$\sigma^2_{WC}$ (Control Variance).....	<b>0.3</b>
$\rho$ (Treatment, Control Correlation) .....	<b>0.7</b>

## Non-Unity Null Tests for Two Total Variances in a 2x2 Cross-Over Design

## Output

Click the Calculate button to perform the calculations and generate the following output.

## Numeric Reports

## Numeric Results

Solve For: [Sample Size](#)

Hypotheses:  $H_0: \sigma^2_{TT}/\sigma^2_{TC} = R_0$  vs.  $H_1: \sigma^2_{TT}/\sigma^2_{TC} \neq R_0$

Power		Sequence Sample Size			Total Variance			Within-Subject Variance		Between- Subject (Treatment, Control) Correlation ρ		Alpha
					Ratio							
					H0 (Null) R0	Actual R1	Control σ²rc	Treatment σ²wr	Control σ²wc			
Target	Actual	N1	N2	N								
0.9	0.9012	91	91	182	0.8	0.5	0.8	0.2	0.3	0.7	0.05	
0.9	0.9001	957	957	1914	0.8	0.7	0.8	0.2	0.3	0.7	0.05	
0.9	0.9000	1190	1190	2380	0.8	0.9	0.8	0.2	0.3	0.7	0.05	
0.9	0.9006	336	336	672	0.8	1.0	0.8	0.2	0.3	0.7	0.05	
0.9	0.9011	169	169	338	0.8	1.1	0.8	0.2	0.3	0.7	0.05	
0.9	0.9026	78	78	156	0.8	1.3	0.8	0.2	0.3	0.7	0.05	

Target Power	The desired power value entered in the procedure. Power is the probability of rejecting a false null hypothesis.
Actual Power	The actual power achieved. Because N1 and N2 are discrete, this value is usually slightly larger than the target power.
N1	The number of subjects in sequence 1.
N2	The number of subjects in sequence 2.
N	The total number of subjects. $N = N1 + N2$ .
R0	The total variance ratio used to define the null hypothesis, $H_0$ .
R1	The value of the total variance ratio at which the power is calculated. $R1 = \sigma^2_{TT} / \sigma^2_{TC}$ .
$\sigma^2_{TC}$	The total variance of measurements in the control group. Note that $\sigma^2_{TC} = \sigma^2_{BC} + \sigma^2_{WC}$ .
$\sigma^2_{WT}$	The within-subject variance of measurements in the treatment group.
$\sigma^2_{WC}$	The within-subject variance of measurements in the control group.
$\rho$	The between-subject correlation of the treatment versus control measurements.
Alpha	The probability of rejecting a true null hypothesis.

## Summary Statements

A 2x2 cross-over design will be used to test whether the total variance ratio ( $\sigma^2_{TT} / \sigma^2_{TC} = \sigma^2_{Total, Treatment} / \sigma^2_{Total, Control}$ ) is different from 0.8 ( $H_0: \sigma^2_{TT} / \sigma^2_{TC} = 0.8$  versus  $H_1: \sigma^2_{TT} / \sigma^2_{TC} \neq 0.8$ ). Each subject will alternate treatments (T and C), with an assumed wash-out period between measurements to avoid carry-over. For those in the Sequence 1 group, the first treatment will be C, and the sequence is [C T]. For those in the Sequence 2 group, the first treatment will be T, and the sequence is [T C]. The comparison will be made using a two-sided, variance-difference test (treatment minus control) as described in Chow, Shao, Wang, and Lokhnygina (2018), with a Type I error rate ( $\alpha$ ) of 0.05. For the control group, the total variance ( $\sigma^2_{TC}$ ) is assumed to be 0.8, and the within-subject variance is assumed to be 0.3. The within-subject variance of the treatment group is assumed to be 0.2. The between-subject correlation between the treatment and control measurements per subject is assumed to be 0.7. To detect a total variance ratio ( $\sigma^2_{TT} / \sigma^2_{TC}$ ) of 0.5 with 90% power, the number of subjects needed will be 91 in Group/Sequence 1, and 91 in Group/Sequence 2.

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## Dropout-Inflated Sample Size

Dropout Rate	Sample Size			Dropout-Inflated Enrollment Sample Size			Expected Number of Dropouts		
	N1	N2	N	N1'	N2'	N'	D1	D2	D
20%	91	91	182	114	114	228	23	23	46
20%	957	957	1914	1197	1197	2394	240	240	480
20%	1190	1190	2380	1488	1488	2976	298	298	596
20%	336	336	672	420	420	840	84	84	168
20%	169	169	338	212	212	424	43	43	86
20%	78	78	156	98	98	196	20	20	40

Dropout Rate The percentage of subjects (or items) that are expected to be lost at random during the course of the study and for whom no response data will be collected (i.e., will be treated as "missing"). Abbreviated as DR.

N1, N2, and N The evaluable sample sizes at which power is computed. If N1 and N2 subjects are evaluated out of the N1' and N2' subjects that are enrolled in the study, the design will achieve the stated power.

N1', N2', and N' The number of subjects that should be enrolled in the study in order to obtain N1, N2, and N evaluable subjects, based on the assumed dropout rate. After solving for N1 and N2, N1' and N2' are calculated by inflating N1 and N2 using the formulas  $N1' = N1 / (1 - DR)$  and  $N2' = N2 / (1 - DR)$ , with N1' and N2' always rounded up. (See Julious, S.A. (2010) pages 52-53, or Chow, S.C., Shao, J., Wang, H., and Lokhnygina, Y. (2018) pages 32-33.)

D1, D2, and D The expected number of dropouts.  $D1 = N1' - N1$ ,  $D2 = N2' - N2$ , and  $D = D1 + D2$ .

## Dropout Summary Statements

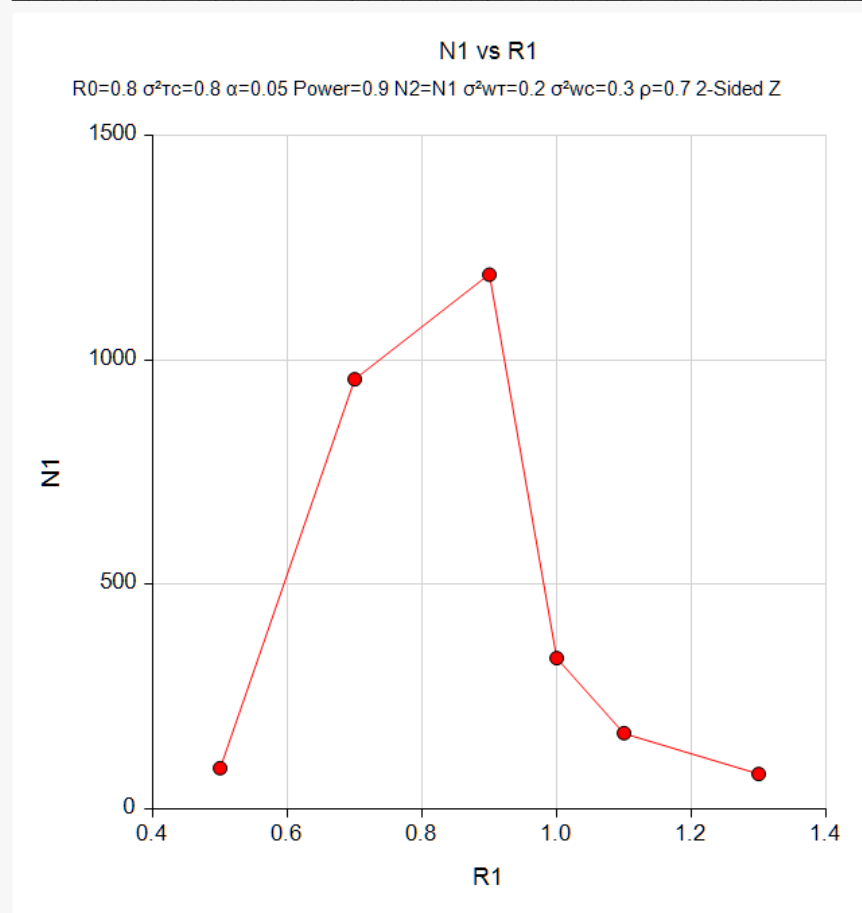
Anticipating a 20% dropout rate, 114 subjects should be enrolled in Group 1, and 114 in Group 2, to obtain final group sample sizes of 91 and 91, respectively.

## References

- Chow, S.C., Shao, J., Wang, H., and Lokhnygina, Y. 2018. Sample Size Calculations in Clinical Research, Third Edition. Taylor & Francis/CRC. Boca Raton, Florida.
- Chow, S.C., and Liu, J.P. 2014. Design and Analysis of Clinical Trials, Third Edition. John Wiley & Sons. Hoboken, New Jersey.

## Plots Section

### Plots



This plot shows the relationship between sample size and R1.



## Example 2 – Validation using Chow et al. (2018)

Chow *et al.* (2018) page 227 contains an example that we will use to validate this procedure.

Set power = 0.8,  $R_0 = 1.21$ , significance level = 0.05, and  $R_1 = 0.52$ . Also,  $\sigma^2_{TC} = 0.25$ ,  $\sigma^2_{WT} = 0.04$ ,  $\sigma^2_{WC} = 0.09$ , and  $\rho = 1.0$ . The sample size per sequence is computed to be 17 for a lower, one-sided test.

Note that there is a typo in the book. The variance estimate of 0.147 is changed to 0.153 in the formula for  $n_s$ . This error is not enough to change the final result.

### Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

#### Design Tab

Solve For ..... **Sample Size**  
 Alternative Hypothesis ..... **One-Sided ( $H_1: \sigma^2_{TT}/\sigma^2_{TC} < R_0$ )**  
 Power..... **0.80**  
 Alpha..... **0.05**  
 Sequence Allocation ..... **Equal ( $N_1 = N_2$ )**  
 $R_0$  ( $H_0$  Variance Ratio)..... **1.21**  
 $R_1$  (Actual Variance Ratio) ..... **0.52**  
 $\sigma^2_{TC}$  (Control Variance)..... **0.25**  
 $\sigma^2_{WT}$  (Treatment Variance) ..... **0.04**  
 $\sigma^2_{WC}$  (Control Variance)..... **0.09**  
 $\rho$  (Treatment, Control Correlation)..... **1**

### Output

Click the Calculate button to perform the calculations and generate the following output.

#### Numeric Results

Solve For: **Sample Size**

Hypotheses:  $H_0: \sigma^2_{TT}/\sigma^2_{TC} \geq R_0$  vs.  $H_1: \sigma^2_{TT}/\sigma^2_{TC} < R_0$

Power		Sequence Sample Size			Total Variance			Within-Subject Variance		Between- Subject (Treatment, Control) Correlation ρ	Alpha
					Ratio						
Target	Actual	N1	N2	N	H0 (Null) R0	Actual R1	Control σ²rc	Treatment σ²wt	Control σ²wc		
0.8	0.8157	17	17	34	1.21	0.52	0.25	0.04	0.09	1	0.05

The sample size is computed to be 17 per sequence and matches Chow et al. (2018).