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Chapter 141

Non-Unity Null Tests for the Ratio of Within-Subject Variances in a 2×2M Replicated Cross-Over Design

Introduction

This procedure calculates power and sample size of inequality tests of within-subject variabilities from a 2×2M replicated cross-over design for the case when the ratio assumed by the null hypothesis is not necessarily equal to one. This routine deals with the case in which the statistical hypotheses are expressed in terms of the ratio of the within-subject variances.

This design is used to compare two treatments which are administered to subjects in different orders. It has two treatment sequences. Here, M is the number of times a particular treatment is received by a subject. For example, if M = 2, the design is a 2×4 cross-over. The two sequences would often be

sequence 1: RTRT sequence 2: TRTR

It is assumed that either there is no carry-over from one measurement to the next, or there is an ample washout period between measurements.

Technical Details

This procedure uses the formulation given in Chow, Shao, Wang, and Lokhnygina (2018).

Suppose x_{ijkl} is the response in the *i*th sequence (i = 1, 2), *j*th subject (j = 1, ..., Ni), *k*th treatment (k = T, C), and *l*th replicate (l = 1, ..., M). The mixed effect model analyzed in this procedure is

$$x_{ijkl} = \mu_k + \gamma_{ikl} + S_{ijk} + e_{ijkl}$$

where μ_k is the kth treatment effect, γ_{ikl} is the fixed effect of the kth replicate on treatment k in the kth sequence, S_{ij1} and S_{ij2} are random effects of the kth subject, and kth is the within-subject error term which is normally distributed with mean 0 and variance $V_k = \sigma_{Wk}^2$.

Unbiased estimators of these variances are found after applying an orthogonal transformation matrix *P* to the x's as follows

$$z_{ijk} = P'x_{ijk}$$

where P is an $m \times m$ matrix such that P'P is diagonal and $var(z_{ijkl}) = \sigma_{Wk}^2$.

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For example, in a 2×4 cross-over design the z's become

$$z_{ijk1} = \frac{x_{ijk1} + x_{ijk2}}{2} = \bar{x}_{ijk}.$$

and

$$z_{ijk2} = \frac{x_{ijk1} + x_{ijk2}}{\sqrt{2}} = \bar{x}_{ijk}.$$

In this case, the within-subject variances are estimated as

$$\hat{V}_T = \frac{1}{(N_1 + N_2 - 2)(M - 1)} \sum_{i=1}^{2} \sum_{j=1}^{N_i} \sum_{l=1}^{M} (z_{ijTl} - \bar{z}_{i.Tl})^2$$

and

$$\hat{V}_C = \frac{1}{(N_1 + N_2 - 2)(M - 1)} \sum_{i=1}^{2} \sum_{j=1}^{N_i} \sum_{l=1}^{M} (z_{ijCl} - \bar{z}_{i.Cl})^2$$

Testing Variance Inequality with a Non-Unity Null

The following three sets of statistical hypotheses are used to test for variance inequality with a non-unity null

$$H_0: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} \ge R0$$
 versus $H_1: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} < R0$,

$$H_0: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} \le R0$$
 versus $H_1: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} > R0$,

$$H_0: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} = R0$$
 versus $H_1: \frac{\sigma_{WT}^2}{\sigma_{WC}^2} \neq R0$,

where R0 is the variance ratio assumed by the null hypothesis.

The corresponding test statistics are $T=(\hat{V}_T/\hat{V}_C)/R0$. Upon making the usual normality assumptions, T is distributed as an $F_{d,d}$ random variable where

$$d = (N_1 + N_2 - 2)(M - 1).$$

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Power

The corresponding powers of these three tests are given by

Power =
$$P\left(F < \frac{R0}{R1} F_{\alpha,d,d}\right)$$

Power = $1 - P\left(F < \frac{R0}{R1} F_{1-\alpha,d,d}\right)$
Power = $P\left(F < \frac{R0}{R1} F_{\alpha/2,d,d}\right) + 1 - P\left(F < \frac{R0}{R1} F_{1-\alpha/2,d,d}\right)$

where F is the common F distribution with the indicated degrees of freedom, α is the significance level, and R1 is the value of the variance ratio stated by the alternative hypothesis. Lower quantiles of F are used in the equation.

A simple binary search algorithm can be applied to this power function to obtain an estimate of the necessary sample size.

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Example 1 - Finding Sample Size

A company has developed a generic drug for treating rheumatism and wants to compare it to the standard drug in terms of the within-subject variability. A 2 x 4 cross-over design will be used to test the inequality using a two-sided test.

Company researchers set the variance ratio under the null hypothesis to 0.75, the significance level to 0.05, the power to 0.90, M to 2, and the actual variance ratio values between 0.5 and 1.2. They want to investigate the range of required sample size values assuming that the two group sample sizes are equal.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 1** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Sample Size
Alternative Hypothesis	Two-Sided (H1: σ²wτ/σ²wc ≠ R0)
Power	0.90
Alpha	0.05
Sequence Allocation	Equal (N1 = N2)
M (Number of Replicates)	2
R0 (H0 Variance Ratio)	0.75
R1 (Actual Variance Ratio)	0.5 0.6 0.9 1 1.1 1.2

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Output

Click the Calculate button to perform the calculations and generate the following output.

Numeric Reports

Numeric Results

Solve For: Sample Size Variance Ratio: σ^2 WT / σ^2 WC

Hypotheses: H0: $\sigma^2 w \tau / \sigma^2 w c = R0$ vs. H1: $\sigma^2 w \tau / \sigma^2 w c \neq R0$

Dou			Sequenc		Number of	Variance	Ratio	
Pow Target	Actual	N1	N2	N	Replicates M	H0 (Null) R0	Actual R1	Alpha
0.9	0.9015	130	130	260	2	0.75	0.5	0.05
0.9	0.9003	424	424	848	2	0.75	0.6	0.05
0.9	0.9001	634	634	1268	2	0.75	0.9	0.05
0.9	0.9006	256	256	512	2	0.75	1.0	0.05
0.9	0.9005	145	145	290	2	0.75	1.1	0.05
0.9	0.9011	97	97	194	2	0.75	1.2	0.05

	hypothesis.
Actual Power	The actual power achieved. Because N1 and N2 are discrete, this value is usually slightly larger than the
	target power.
N1	The number of subjects in sequence 1.
N2	The number of subjects in sequence 2.
N	The total number of subjects. $N = N1 + N2$.

Μ

The number of replicates. That is, it is the number of times a treatment measurement is repeated on a R0 The within-subject variance ratio used to define the null hypothesis, H0.

Target Power The desired power value entered in the procedure. Power is the probability of rejecting a false null

The value of the within-subject variance ratio ($\sigma^2 w \tau / \sigma^2 w c$) at which the power is calculated. R1 Alpha The probability of rejecting a true null hypothesis.

Summary Statements

A 2x2M replicated cross-over design will be used to test whether the within-subject variance ratio (σ^2 wt / σ^2 wc = σ²Within,Treatment / σ²Within,Control) is different from 0.75 (H0: σ²wτ / σ²wc = 0.75 versus H1: σ²wτ / σ²wc ≠ 0.75). Each subject will alternate treatments (T and C), with an assumed wash-out period between measurements to avoid carry-over. With 2 replicate pairs, each subject will be measured 4 times. For those in the Sequence 1 group, the first treatment will be C, and the sequence is [C T C T]. For those in the Sequence 2 group, the first treatment will be T, and the sequence is [T C T C]. The comparison will be made using a two-sided, variance-ratio F-test (with the treatment within-subject variance in the numerator), with a Type I error rate (α) of 0.05. To detect a within-subject variance ratio (σ^2 wr / σ^2 wc) of 0.5 with 90% power, the number of subjects needed will be 130 in Group/Sequence 1, and 130 in Group/Sequence 2.

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Dropout-Inflated Sample Size

	s	Sample S	ize	ı	pout-Inf Enrollme Sample S	ent	1	Expected Number of Dropouts	of
Dropout Rate	N1	N2	N	N1'	N2'	N'	D1	D2	D
20%	130	130	260	163	163	326	33	33	66
20%	424	424	848	530	530	1060	106	106	212
20%	634	634	1268	793	793	1586	159	159	318
20%	256	256	512	320	320	640	64	64	128
20%	145	145	290	182	182	364	37	37	74
20%	97	97	194	122	122	244	25	25	50
Dropout Rate	The percentag		٠,	•		e lost at randor e treated as "r	_		
N1, N2, and N	The evaluable	sample si	zes at which p	ower is com	puted. If N	I1 and N2 sub	jects are eva	aluated ou	t of the
		•			, ,	gn will achieve			
N1', N2', and N'	inflating N1 a	sed on the and N2 usi ded up. (S	assumed dro ng the formula	pout rate. Af as N1' = N1 / A. (2010) pa	ter solving ′ (1 - DR) a	n order to obta for N1 and Ni and N2' = N2 / or Chow, S.C	2, N1' and N (1 - DR), wi	2' are calc th N1' and	ulated b
D1, D2, and D	The expected	` ,			20 NO	NO and D F	M . DO		

Dropout Summary Statements

Anticipating a 20% dropout rate, 163 subjects should be enrolled in Group 1, and 163 in Group 2, to obtain final group sample sizes of 130 and 130, respectively.

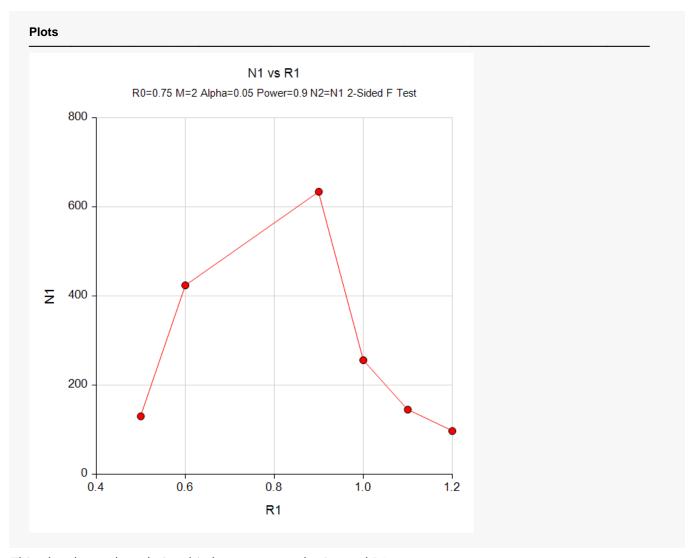
References

Chow, S.C., Shao, J., Wang, H., and Lokhnygina, Y. 2018. Sample Size Calculations in Clinical Research, Third Edition. Taylor & Francis/CRC. Boca Raton, Florida.

Chow, S.C., and Liu, J.P. 2014. Design and Analysis of Clinical Trials, Third Edition. John Wiley & Sons. Hoboken, New Jersey.

This report gives the sample sizes for the indicated scenarios.

Plots Section



This plot shows the relationship between sample size and R1.

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Example 2 - Validation using Chow and Liu (2014)

We will use an example from Chow and Liu (2014) page 509 to validate this procedure.

In this example, the significance level to 0.05, M is 2, the power is 0.80, and the actual variance ratio is 0.3/0.45 or about 0.667. The resulting sample size is found to be 98 per sequence.

Setup

If the procedure window is not already open, use the PASS Home window to open it. The parameters for this example are listed below and are stored in the **Example 2** settings file. To load these settings to the procedure window, click **Open Example Settings File** in the Help Center or File menu.

Solve For	Sample Size
Alternative Hypothesis	Two-Sided (H1: σ²wτ/σ²wc ≠ R0)
Power	0.80
Alpha	0.05
Sequence Allocation	Equal (N1 = N2)
M (Number of Replicates)	2
R0 (H0 Variance Ratio)	1
R1 (Actual Variance Ratio)	0.667

Output

Click the Calculate button to perform the calculations and generate the following output.

Solve Fo Variance Hypothes	Ratio: σ	Sample Size σ^2 wr / σ^2 wc H0: σ^2 wr/ σ^2 wc = R0 vs. H1:		: σ²wτ/σ²wc ≠ R0				
Power			Sequence Sample Size		Number of	Variance Ratio		
Target	Actual	N1	N2	N	Replicates M	H0 (Null) R0	Actual R1	Alpha
0.8	0.8032	98	98	196	2	1	0.667	0.05

The sample sizes match Chow and Liu (2014).